

1197 TM: Hard ware, Software for "Ultimate TM":

1/2 ~~+~~

01: 96 TM 68.40 This is an old idea that I had on a long Airplane trip from (I think) Conf. (or so):
I did write it up but can't find it. It was an outgrowth of some work on how to use ~~the~~ ~~many~~ buppy CPUs to do good work — One result was that I probably wouldn't ~~the~~ ~~many~~ have much advantage in using "untested", say, CPUs.

The I units use untested CPUs this way: Say a normal production line for CPUs has a 75% yield — w 100% testing. Then I accept every other CPU (untested), as long as I know the other CPUs ~~are~~ have 75% pass rate. If they untested ones are much cheaper, this would perhaps be fine! I doubt I'd only need $n \times \frac{1}{.75} = n \times \frac{4}{3}$ many CPUs.

10 (Anyway, as I remember, I first approximation had many CPUs working in parallel. They had certain modules for commonly used operations & would queue-up to use them. If the queue got too long, dup's of those modules would be made.

If a module wasn't used enough, it would be removed & its parts used to make ^{new} other parallel problem solvers, or other ~~new~~ commonly used modules.

20 Second Approx ~~Another approximation~~ To realize this one could have something like the "Internet" in which messages would be passed to various CPUs & modules, & the outputs of modules sent to other CPUs & modules. So the entire device was just this message passing system.

In 20, the allocation of resources for modules is "CPU's" ^{≡ makes of trials} is somehow interrelated w. the ~~is~~ constancy of $\frac{PC}{CC}$ resources. & each pt. in time. The whole system of detailed systems of how this has to be done, has to be worked out.

It could be like the Early Connection Machines: that each computer unit has an assoc unit for passing messages only. — In fact, it may be well to use a "Connection Machine" Architecture.

Perhaps the idea of 20 is ~~useful~~ useful only if each CPU infrequently waits for results from submodules but "pipelines" many trials at the ~~same~~ same time. So each CPU is working all of the time. If all trials ^{for a given CPU} are waiting for info from modules, then the CPU starts a new trial.

So, except for the assoc. w. messages being transported, the system can be very efficient (i.e. no time spent waiting): In fact, it would seem that there would be no point in making duplicate modules, since long queues would not be sufficient.

1.40 Corresponding to a "long que" in a non-pipelined TM, would be a very large no. of trials being done in 11 by a "pipelined" ~~arch~~ architecture.

Also note: If a module has an empty que much of the time, then it is being Very inefficient! So probably it is best to have modules with que's containing 3 or 4 "waiters".

Also note: There is no point in having a module unless the module is somehow more efficient than having each CPU doing the task. On the other hand, it may be best to have each "cpu" that runs trials, have no computing capability other than clearing — i.e. keeping track of I/O strings for each trial. — seeing that the I/O needs of each trial are satisfied.

10-13 is very imp.! I don't immediately see how to do this. I'm thinking of a FPU (e.g.) How would this be "simulated" in a TM of t. kind I'm contemplating?

Re: Many "crumby" cpus in 11 trying to run code a corpus: There was t. problem that if one CPU did find a "short code", it would not be possi. to know how to continue that code, because of r. uncerainty in just what parts "fail" or "ambiguous":

A (perhaps) soln: We have all these cpus in 11 making trials: If one does succeed, it automatically tries random continuings of its code to find probable continuings. of F. corpus.

A big dirty is in implementing an exhaustive L search. We can't do it unless w. any certainty w. "Crumby" machines. But a crumby machine can be programmed in each program has a reasonable distribution for its output.

A possibly disturbing thing: There is some area, short codes but was the machine ~~and~~ it was assigned to, didn't do it rite.

Is this a nacy dirty — i.e. is there not a good chance that a dirty ^{short} code of this, or some other "crumby" machine, would by mistake, do that code?

A crumby machine can be modeled by a ~~crumby~~ non-crumby machine w. 2 inputs: 1 is b. regular p.p.m., 2 to other input is random. The degree of "Crumbyness" is measured by relative rate at which t. machine asks for t. 2 inputs: A ^{simulation} very crumby machine mainly asks for random inputs. A conversly for very non-crumby machines.

Some (maybe new) approaches to using randomness in Leach:

1) At a gn. pt. in t. search we consider all codes of length L or less. We choose one of these at random w. = prob. i we work on it w. $c \propto 2^{-L}$, if it is of length R. After working on it for that time, we mark it, i do another random trial. If we hit one that's been waded on in this round, we skip it & make a new random trial. The ^{fraction of} ~~time~~ time spent in hitting sites again and again may be quite small. — BUT this will have to be checked.

We do maybe 2L trials, (?) (maybe much less): Then we maybe double t. time we spend on each trial. — or, ~~if~~ we random the state of each machine waded on, we can spend time $\propto 2^{-L}$ on each CPU, whenever we hit it: No t. in time spent per hit: This amounts to Levin Search w. time sharing, but random choices.

2) Another (crazy) idea: Make checks probly $\propto 2^{-L}$: also, am. of time spent is $\propto 2^{-L}$.

This \sqrt{p} d.f. is famous for 2-Phy's! ^{Spec.} $\rightarrow 42.01$

01' 6TM 68, 40 spec: An Attractive Idea:

Take a fairly difficult math problem & find one or more conc. nets for it.

Doing it in Euc. Geom might be easy for me. I'm afraid that human "Geometric intuition" would be difficult to express in machine form ... but I think others in A.I. have worked on this & I may be able to use/understand their work.

I was thinking of avoiding the need of human Geom intuitions - but it may well be that I need many parallel concept nets for different "domains" - so TM can make analogies betw. them.

If I try to get TM to use my methods of solving problems, it will have to have lots of geometric intuition!

I might be able to get some other people interested in this problem of really explaining how various hard problems really can be solved - For use in both Human & Machine Learning.

Before, I wanted to write TSO's "from t. bottom up" (starting w. easy, fundamental, "primitive" problems.) - Now, it seems that "top down" may be easier/better.

In Euc. geom, we already have a lot of ordering of t. Thms. by Euclid & others. T. Geometry of Euclid is, of course, very incomplete, but perhaps I want to leave it that way. Humans are able to work w. it - prove Theorems in it.

2000

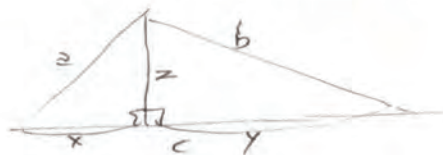
A poss. starting point: Pyth. Thm.

First: How to conjecture t. Thm!

Well, because of 5.3. congruency, if one knows 2 sides & th. included \angle , t. Third side must be determined uniquely.

Well, with 2 rt. \angle . If we had formula for opposite side as func. of 2 legs! Would we be able to solve other problems?

Well, given Δ w. 3 known sides



If $z = f(a, x)$

$z = f(b, y)$

Then $f(a, x) = f(b, y)$

also $xy = c$

So we can solve for x & y & then get z & get area of Δ . So it might be

Worth while!

From similarity considerations,

find z as a func. of x .



By doing drawing and measuring, we find for small x

$z \approx (1 + \frac{1}{2} x^2)$

for large x , $z \approx x$.

A good guess would be

$z = \sqrt{1 + x^2}$ since it has

desired properties: Checking w. carefully drawn diagrams confirms t. conjecture.

By similarity arguments, if we multiply every thing by a ;

$az = \sqrt{a^2 + (ax)^2}$ we can check

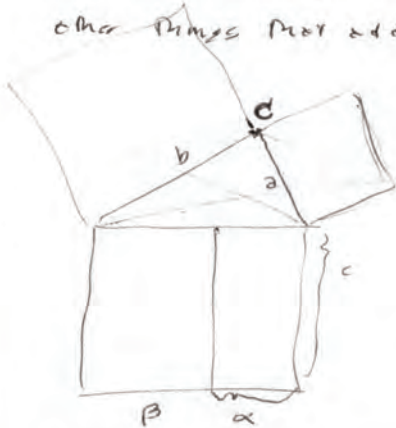
$a^2 z^2 = a^2 + a^2 x^2$ & get this ok.

so in general

hypot $z = \sqrt{a^2 + b^2}$.

or $h^2 = a^2 + b^2$.

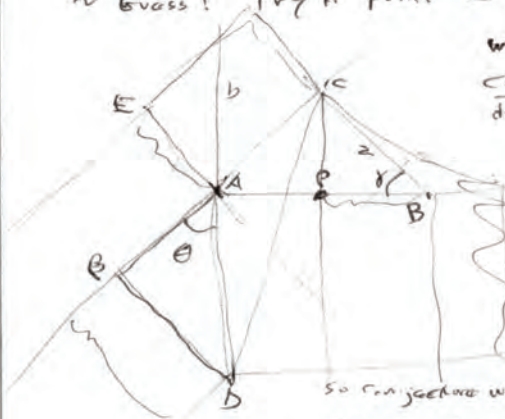
01 2b.40: The proof! One way would be to show ^{2 rows} squares on 2 legs summed to sq. on hypot. Since 1. squares themselves don't fit into 1. large sq. — lets try to show 1. 2 ~~leg~~ leg squares are = to 2 other things that add up to 1. hypot.



Now, there must be an α , \exists 1. square 2 has same area as 1. rect angle (c, α) .

08

14 ~ Guess! Try 1. point C as 2.5 divided:



we can verify this conjecture by drawing pictures!
Also, conjecture is reasonable, since BP gets small when α is small. Also, when $b=c$ $PA=PB$ —

So conjecture works on ends and in middle.

If $\Delta ADC = \frac{1}{2} \text{ area of } b^2$

from dist of D from line AC, must

$\equiv AE = AC$

also if $AD = c$ and $DP = b$

then βA must be α

So $\Delta \beta DA$ is congruent to ABC

O.H. prob's if we compare

Since $\angle \theta = \angle \alpha$ \therefore 1. 2 Δ 's must be \sim \therefore having 1. side size $AB = AD$ in common, they are congruent —
So 1. rest of the proof follows.

Some diff. parts of 1. "proof"

1) The conjecture that $a^2 + b^2 = c^2$ was obtained using partly algebraic reasoning.

2) .08 is a large jump

3) .14 is another large jump.

These 3 points are interesting! They give an index of the difficulty of 1. proof.

There is some q. as to how much detail I want to put into 1. "proofs" at this point.

Another tack: Probably I'll want to pursue the T.S.Q. in Algebra in 11 w. Geometry. This will make it poss. to use analogs from 1. 2 domains to help each other.

I can't immediately think of a NY Point to work backward from. Euclid made a nice Geometry "TREE" — a \sim conc. net. But I have no clear idea of such a net for Algebra.

Maybe "Ask around" — Marvin? $\rightarrow 4b.03$

.20

N.B. All of the details of my "proofs" in conc. nets don't have to be "correct" (i.e. of reasonable c.j.s.). If most of them are O.K., this will still give me much needed practice.

For "Conjectures" we first aim for "Quick Abort" i.e. a quick disproof.

— By taking extreme conditions — which usually simplify 1. analysis.

\rightarrow Also, Go over ~~the~~ these (recent) long lists of new ways to make T.S.Q.S

Refs:

01 3.10b: On "HINTS": A "Hint" is ^(any kind of something) that narrows the search space a/o reduces such (time).
 Examples of "Hints" ① Any Subgoal: like that.
 Conjecture of 3, 142 ② We can show $a^2 + b^2 = c^2$ by finding some # $\alpha, \beta, \gamma \Rightarrow a + \gamma = \gamma$ and $a^2 = \alpha, b^2 = \beta, c^2 = \gamma$. ③

IMPT Q: After TM solves

This problem (or a set of problems), with or w/o "hints"; How much easier is it for TM to solve other problems? Well, for Proust being proved: TM ~~has~~ has to Proust with "set of facts". — (which is a kind of $\{ng\}$). Also, it has \uparrow to SSZ for use of each of the heurs it vsd. Also, it has \uparrow to corpus of examples on which to search for heurs. (i.e. examples of successful problem solns)

There is a large Corpus of problems in the A.I. literature that have been "Solved by heuristic methods". Each of these solns. is a (perhaps useful) example of the kind of work I'll be doing. T. only by differences:
 ① I will eventually put in detail so that the "demo" is doable by Machine w. acceptable cc. ② I will have quantitative methods to get pc's, cc's for \angle such ③ I will have methods to search for heurs of all kinds, amongst successful problem solns

33 ④ A Hint differs from a heur. A hint is usually relatively narrowly applicable to a particular problem. A heur is more generally applicable. A "hint" may be somewhat Ad Hoc.

In general, heur, a person or TM would probably want to remember known "hints" & — they are good candidates to be

heurs ^{or} parts of heurs — or may ^{specialize} _{generalize} ~~heurs~~ to heurs.

03 36.20 Actually, I don't really need a "Euclid" of Algebra! Just start at any heuristic pt. (Term.) & work downward. In fact, this is what I'll do for the Pyth. Thm. ... I don't have any Plane Geometry Texts around.

Some needed ideas in the "proof" of Pyth:

- 1) I. idea of Area: "Area" is assoc. "region" w. any region of a plane bounded by lines or curves. T. "region" has property that one can move between any 2 pts. in it w/o crossing the boundary. T. interior pts. are an "equiv. class"
- 2) Area is Additive: T. sum of areas of all curves inside curve C, is equal to the Area of C.
- 3) 2 concurrent figures have same area.
- 4) properties ② & ③ pretty well make area unique if one specifies area of some specific figure (say a unit square). By making small squares one can give upper & lower bounds on area (numerical) of most any "curved" region.

5) In the present "Pyth Thm": T. sum of ~~sq~~ areas of sps. on legs = area of sq. on Hypot. There is no numerical interpretation of "Area"

— Hvr, in the Analysis of 2b.12: Motivation for conjecturing the Term. and 2b.25ff; on the conjecture of the Thm itself — there is lots of Algebra & numerical interpretation of "Area". I think we assume that if "a" is edge of a square, its "Area" is a^2 .

1497 TM Language, Human

.01 96TM 43.40: (Perhaps most relevant to papers & discuss. therein, of 96TM 43.16-19).

.03 If Grammars of Human lang. are largely engraved onto t. DNA, & this is not true of other Animals, then this might explain Man's ability to acquire many kinds of technology & much better than Animals. Hvr. t. learning ability of animals w. large Brains (Elephants, Whales, other aquatic mammals) has ~~been~~ probably not been adequately examined!

Also, consider the visual learning capabilities of Pigeons (it true!).

This (.03) ^{would} put animals in a "Tower of Babel" situation w.r.t. man's ~~relative~~ relatively easy communication w. one another.

I'm assuming this "Language capability" makes certain (otherwise diff.) kinds of induction much easier.

Another possy. is that Lang. doesn't so much affect learning ability directly as it does ease of intercommunication — so animals w. a good lang. can build up more useful (interesting, complex) social str.

[The social insects have elaborate social str. ... I'm not sure t. situation is at all comparable to what humans have]

Social str. is very important for technology. E.G. note lack of development of technology in China: tho they originated many impt inventions, t. society did not encourage their further development — as was done in Europe.

Language learning: (TSQ)

5
"Computer Jan 97. p 72: Discn. & refs: Some Ppms to read Edgar reports & get specific info from them: T. P. M. does do some "learning".

"Eloise" is Name of Ppm.

Also "Frump"

This Art. has lots of other things on learning in A.I. — Refs to many Ppms of interest.

- 01 4A0b: A Big Q: There are difrent %s of rigor in proofs. Very rigorous proofs are easy to express to Machines. Non-rigorous proofs are very unclear to Machines.

This has to do w. \rightarrow assumptns humans normally make ... Assumptns not necly in t. proofs of f. Thm being proved!

Can we give t. machine those Assumptns?

Can we define any criteria for a non-rigorous proof? — We hope that usually a non-rig. proof can be modified (w.o. too much diffy) to become a rigorous proof!

Perhaps I should create a study some examples of rigorous/non-rigorous proofs — see if I can formalize a non-rig. proof well enuf ~~for a machine~~ to devise a criterion for acceptance/rejection.

Any old \bullet Plane Geom. text should have fine non-rigorous proofs.

~~How~~ How did A.I. workers deal w. this?

Look in Encyc. of A.I.

In general, if I provide criteria for acceptance of proofs, this will be an unambiguous criterion, so it will be a "formal System". T. only Q~~ue~~ then is: is it consistent? This is usually undecidable, but one can have strong intuitive argts for its consistency. Or one can only allow use of t. system at a certain logical depth: in which case we ~~restrictive~~ could, in theory, verify its "consistency" at \leq that depth.

The main problem w. unrigorous ideas is — we have to explicitate them to some extent. This can be a big job. — Anyway \rightarrow 7a.01

List of immediate diffys:

① Problem of rigor (6a.01) \equiv (6a.01)

② Clear Soln. of GA probs by Lsrch.

e.g. 11 input Muxer: try

$$11 = 3 + 2^2 \quad | \quad 2 + 2^2 = 6 \text{ input muxer}$$

$$20 = 4 + 2^4$$

Lsrch has to be at least as good as probably much better than ordinary Gen. Algs.

③ Do try to find "proof" of idea in footnote of Sol89: i.e. all heurs can be expressed as modifn. of t. p.d. if one is allowed an adequate set of input params. for t. induction: I should (at least) be able to show that t. system can deal w.

"Quick Abort":

"Input params" include "traces" of previous solns. (i non-solns) of problems — i suppling else TM may ask for.

01 62.40: I'll just have to look at some non-rigorous proofs. In fact, the stuff I have on 2b.01ff may ~~be~~ give some good examples!

Very probly I wouldn't have to look beyond my own work ^{on this problem} to find many usable examples

Re: General Prob. Solving: Perhaps
Look at New/simon "Human Prob. Solving"
Perhaps I could formalize part of
what the prob. solvers were doing in their
"protocols"

30 690 935

This is a kind of essay/lecture, to be distributed to friends/colleagues on t.net, etc.

Part I : Explain my position: That I like to think of A.I. as ushering in the next stage of the evolution of Man; The Trans-human; Discussed by Moravec, Minsky,

Hvr. Discuss ^{of roles} of S.F.: To examine poss. scenarios - fully of development. While this part of S.F. has been Dystopia-prone

as it must be.

Discuss Asimov's Robot ~~st~~ stories: "3 laws".

Why they would be passed: why Prog would be inadequate.

1) obey Don't harm humans by action in action.

2) ...obey humans.

Then ① The "not-conscious" but very smart machine.

Development by Military or Gov. (NSA Always gets Biggest, fastest Machines first)

Use for developing Military A.W./S.W.

Use for control.

That first Nation to get it can "win" — by preventing any one ...

Domestic or foreign from working on A.I.

Discuss "Science of 100 or 200 yrs hence".

"Supremacy" effect where IPC $\rightarrow \infty$ in future time; what "0" means.

①② What happens when > 1 power simultly have very smart machines?

② Discuss Consciousness in machines:

obtained by ② very smart machines in suitable envt.

③ Technologically Augmented humans

Idea that as dominant species becomes more powerful, the (likelihood of destruction of the world in a war, ↑.

That humans set desired w. goals appropriate to small amt. of power per individual — so ^{normal, rational, well} a person striving to become more powerful in his culture (or world),

is not a threat to the world as a whole, because that person couldn't do that much.

This particular goal could, instead, serve to develop a person in a way, useful to the culture as a whole.

A very strong person like Napoleon, Alexander, Genghis Khan ... could have very powerful effect on the whole world.

There is the ditty in very intelligent machines that are goal-oriented. In general, they will be able to change the goal or trivialize its achievement:

An early version: McCarty's reinforcement machine that soldered its own reinforcement channel to "On" (\equiv "Maximum").

In Humans, Morphine, Heroin... can perform similar function. In general evolution, master brain obtains ^{some} sexual pleasure w.o. attaining evolution's "goal": Tho usually the reproductive goal is attained by social & other pressures that make normal reproduction have additional pleasure.

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So, how do we deal w. this problem in very intelligent machines? As soon as the machine has much understanding of R.W. & its position in it, it will realize that there are alternative, simple, ways to achieve its "goal" — by internal rewiring.

We could fix it so that part of its goal was that it was not allowed to do internal rewiring (or otherwise modify its goal). Is this possible?

250
295A

Also, there is always the possy. Part of goal of machine could become accidentally modified by ^{Bug} H.W. or S.W. malfunction — or Accidentally modified — by earthquake, Airplane crash or other catastrophic accident.

I could study the problem at a micro level. See just what occurs when T.M. "realizes" its in a R.W. & that it had best interact intelligently w. R.W. in order to achieve its own ~~goals~~ goals

Superficially, for any kind of goal (or prob), a suff. smart machine could find gross short-cuts to that goal by internal rewiring & re-prior.

SI=20
CU=55
=80
MT=.242

[4.14.98] Another (rough) approach: That TM is not completely governed by a "Goal": That it has other things controlling its parts.

There may be some real ditty that TM gets by "Evolution" mode — so we really don't know in what direction it will evolve.

1.20.97 TM EXPD

10

I may want to rewrite (revise) the Luen "paper": Update w. more recent markets by Levin for book on \mathbb{Z} & INV. probs. Also discuss what it means by "optimality" of Luen: Also perhaps look at Li-vitanyi's discn. of it.

Mention that my pen and ... was one way of generating Prob. Sets.

12697: Things to do
things

Admin

Less critical

11 +

- 1) SM work:
- a) OXB FT Model/quest
 - b) other econ. G.A.
 - c) Related to TM. (General Problem).

Pack stuff at 26B.

- Get Pent. to work web.
- Perhaps find low cost Internet Provider.
- Fix up 10th floor Put books on shelves
- Fix TV set. work on Basement.

Try indexing ~~the~~ Packed Boxes
 " " Stacked Bookshelves

Fix phone so it sounds better
 or get better phone.

Cleanup desk:
 Other things to facilitate rapid work!
 Write "T. dress"
 Write "T. Story of Harvey".

- 2) TM stuff a) CAP: Get Gunkel's list.
 b) Plot General path for TM work.
 say several poss. paths.
 c)

- 3) General:
- a) finish letter to dtd
 - b) Read Sasha's papers
 - c) Read Jurgen's papers; reply to his letters
 - d) Finish "2 bands of induction"
 - e) Perhaps write Review of ALP for Sci (Am or Science, how ALP is eval. approaches almost all "earnings")
 - f) Finish paper for dtd etc.
 - g) revise "Lsrch" paper: look at Wittne's remarks.
 - h) Write paper on "What to do next"

NB How many of these "Things to do" do I want to, have to, do?
 Perhaps make PERT diagram of various projects, Deadlines (if any)
 Make chart of various projects, how much time spent on them, what fraction completed.

25:09: Look at that recent bunch of lists/discn. of TSP's etc.

on Beach washing tool

My present bias is a) work pbs in GA at various levels of difficulty using Lsrch
 b) Take old AI, "exp. systems" (the AI is get it to develop properly, maybe a useful medical or other a.s.?) that have lots of external
 it to coming in.

try capacitor across Mike, inside handle read white from black

~~the~~ List deadlines for projects if they exist.

This will help assess times.

Estimate deadlines for various projects

2 middle = white, green, coupling
 2 outside = white, red, black
 Books ~~from~~
 handset.
 Perhaps try Black handset on Best Phone or other handset.
 5-530 545-600

Phone Steve Beer about 2 arts in recent science on Beach prop. in vivo
 12697 Cont. Radiation Science D20, 1993, Jan 10!

Is it? Could Grace handle it?
 Write Andrew, Isabel

Maybe write proposal for NSF

Things to Do!

1) Re: Pont & 486! See that >lt. imp. files from 486 are on Zip & In feet.
 Q: Why were comports ~~22~~ accepted on 486? why didn't from Pont?
 Try ~~re-creating~~ w/ ~~Mike~~ why didn't from Pont?

This overall planning of the rest of my life: can be useful in helping decide what to get rid of.

General idea: Whether I want to write papers on certain matters or not, I ~~must~~ should write a clear exposition for my own use, of important ideas & discoveries.

The Clarification of the differences of ^{production} GDP, ^{consumption} MNC, MDH act. This
 is important to me, since they have useful approxs (sometimes) — but I must know just what their limitations are, so I can effectively transcend them.

Make list of items w. near deadlines & do them

SM: write up imp. results, techniques, Plurms

first, the most recent stuff on portfolio Management &
 Then how much to bet when one has pd. v.s. default pd.
 Then the default pd. as given by ^{actual} / ^{assumed} prices
 Then to "trap" Black-Scholes model. (second ^{difference} of option price = p.d.)
 Include refs to wherever, where work was done.
 Also stuff on Kelly betting: so yield = $\exp(\sum (\frac{r_i}{\sigma_i})^2)$ ^{gains!}
 Also recent understanding of "the eye test".



Most urgent:

- 1) write DLD
- 2) " Andreas / Isbell — 1/18/96 (letter to Ann) ← ANDREAS Read Lipton's Stuff, Reply. to Hans on 2 first call
- 3) Read Sogoh's papers (get from Grace)

Get Madzara, GMD Straighten out. find out how best to use it.

A. Big Problem: I will not have enough time in my life to do most of the things I've listed on pp 1,2. I really am spreading myself too thin. By conceivably limiting my activity to a few well-chosen projects, I will get much more done.

Also, I can probably get rid of a lot of books, documents & hardware — **once I decide what I don't want to do!**

- .08 as part of ~~trying~~ trying to understand how humans/scientists make discoveries & work problems & make decisions.
- .09 work problems & make decisions.

At the present time, I think I know about the solution to ~~Prob~~ (.08-.09) as well as the more general problem. Tho I like dabbling in ~~all kinds of~~ many things of great variety, I really don't have time for very much of that, just now!

Some lists:

1) Things to read:

- a) Sasha's papers
- b) Jürgen's "
- c) Lipton's "
- d) Gunkel's "future of AI" list
- e) Winfree's comments.

2) Things to write:

- a) Finish 2 kinds of ~~the~~ ^{introduction}
- b) write up SM results (2,20)
- c) What to "do next" paper

- 3) AR
- 4) Write "T. dress"
- 5) "Harvey"

Sci News
⋮


- d) Summary of impl. ideas, Plans in my TM work, ... Bibliography
- e) Finish paper for DLD
- f) may be modify Burez.
- g) Review of ALP for Sci Amer ^{Nature} _(Science).
- h) NSF proposal

Start work on "New Problem!" I want to decide what to work on next: a/o roughly how much time share to put on various projects: (Mainly w.r.t. TM Goal)

One of the poss. paths to TM! Working backwards! TM has knowl. & good prob-solving ability in 1 or more areas. He also understands English well and so read ordinary text about these areas.

To get there! ^{in general:} TM's knowl. in the field ^(s) of interest is not very deep. He can't solve diff. probs., but he does know simple relationships — has some "commonsense", but not very much. He is able to answer simple Q's about the field. He can read in that field & understand and to be able to answer some Q's in a "open book" exam.

Such a TM could perhaps do interesting induction on ordinary text about his ~~the~~ fields of knowledge.

By asking Q's & noting quickly ^{slowly} ~~slowly~~ ^{wrongness} of answers, I get an idea of what TM needs to know, & ask Q's leading in proper directions.

Could TM ~~possibly~~ make use of a dictionary, encyc., or ~~or~~ books? The older approach is to ~~get~~ ^{use} a TM who is very smart in a particular field, then ~~teach~~ ^{teach} it to converse about things in that field.

SW TM could learn to speak well by being given its own spectral sequences, those of acceptable humans, & ~~some~~ ^{relative} "humanness" scores for its various utterances. (Reinforcement ^{lrng.} ~~lrng.~~).

Another Approach: Take a problem (like Pyth ^{conjecture} ~~proof~~ ^{proof}). work out soln. See what concs are needed. Take a different ~~prob.~~, somewhat a problem, see what concs its soln. needs! etc. The ~~prob.~~ ^{prob.} ~~do~~ ^{do} this for many problems & see what needed concs. They ~~are~~ ^{are} in common. We can possibly give these to TM ab initio. This may save a lot of time. (Or, if concs may have common soln concs so we can give TM a lang. to generate those concs. & (presumably) other useful concs.

Re: ① & ②; "Blocks world" is a case in point! But there wasn't any lrng. involved — author of "Blocks world" itself, or of the lang. during the world. Blocks world is rather simple — as is the lang. ~~it~~ ^{used} to describe it.

A Pyth Proof attracts me to ② (1.22): That if we get TM To be very bright in a particular area by lrng. This is a much stronger result for me than a TM that learned easy things & then an easy lang. assoc. w. it. (Re it could learn easy things w. a rather complex lang. assoc. w. it.)

GO
5050420 2/6/97
A.I. ~~scenarios~~
scenarios
72642.
35740
5
Courageous.com
Address
P.O. No.
to Gopher.

Another area of "easy" task design (question): Solving opten probs by Leuch for ²inv solns.
We can find a field w. a very low path from the bottom to top of hill.

SM is another example of this sort of thing: Start in 1920, say & acquire
new, more difficult techniques as we approach the present.

Or, just stay in ~~the~~ present time & try for better & better strategies.

05 would be (impossibly) interesting only if the info used in its predns. was rather general -
kind obtainable from magazines, newspapers.

Well: re: 06 say T.M. really did do a good job re 05. The only thing it
would be interested in were the various params it used in predn., w. prices & bet on.
The task of learning to extract such info. from English durns of the date or newspaper reports
or stockholder reports would seem to be quite separate from TM's ability to
do good SM predn. We could just as well try to teach TM how to extract financial params
from text reports, w/o first teaching TM what the params "mean".

A possibly more interesting literature for TM to learn to understand, would
be not ~~more~~ reports of data, but discussion of various SM techniques.

↳ This would seem to be rather difficult! Ideally, TM would read this
stuff and try to see if it was useful in prediction, or in augmenting its own schemes.

This would seem related to TM's realization that (±)RW. is related to its
main problem (RW: maybe not: because TM is not getting direct f.B.
from RW).

On the other hand: Learning to understand English, may very often involve
some rather complex (not "linguistic") (imp... e.g. figuring out what the referents
of pronouns are - or in general, what is going on in a story that's being told

↳ ~~more~~ Could SW "Agents" develop any kind of useful
more generally useful understanding of humans?

My really main criticism of Schmidu is that I don't see
a clear "Path" in his work, to a very smart TM. He was working
much on polishing the final TM, but not ~~was~~ doing much towards getting something
that was "polishable".

"Path" as a new
sub-TM stream.

Luke's ph. no.
492 3772.

|||| times

A General related Q about a ~~minimal~~ minimal path is the Q
of what ^{to} some likely paths in the existing A.I., ML community? To
estimate this last would involve reviewing the whole A.I., ML field.

I could, but, try this, using only the parts of ^{current} AI/ML that I'm
familiar w.

Perhaps write paper "Paths to A.I." : Maybe publish in "Analogy" -
Send copy to interested people.

Maybe publish "Scarceness of Decm" for A.I. - also in "Analogy".

2.12.96 TM (PAM):

In Paper, perhaps also discuss 4 or 5 aspects of "Consciousness" & how they are not difficult for A.I. to achieve. Also mention "overly eager subordinate" & "serious admin problem" in purely human culture.

B Paper itself: Outline:

1) Classical Approach to A.I. "expert Systems": ~~classical~~
Need for "commonsense" - This is an artifact of the way machines ~~are built~~ acquire info ("knowledge") in "classical A.I." They get it from a human who is good at communicating w/ people who have a common background knowledge.

In a more reasonable type env't, the machine (in fact) starts out w. certain info. ~~the~~ general info. that makes it easy for it to learn the "commonsense" knowledge of the world it will live in. The env't starts w. an amount it needs to get along in the world of 1 year olds.

Idea that since we can make a.s. in very many env'ts. we should be able to combine them in a grand C.S. that would be ~~as~~ as good as a human, but much faster in many areas. Lenati's

Except, perhaps, for the Cyc project (~~Lenati~~), which the hope that this would occur in near future (50 yrs) has withered away from its early euphoria.

Since about 1980 there has been a great ^{upsurge} of interest in machine learning. Neural nets, statistical learning, and genetic algorithms are ~~among~~ but a few of the great variety of techniques applied in this area. In fact, the diversity of techniques has been so great that some scientists have expressed doubt that they have enough in common for the ~~concept~~ idea of "learning" to be at all a useful scientific concept! ~~Conclusion~~

One of the ideas we will develop in this paper is that just about all learning does have a common ^{mathematical} basis, ~~and from this basis~~ ^{from which} we can evaluate the effectiveness of ~~a learning technique~~ ^{various} ~~various techniques~~.

Various techniques and obtain suggestions for improving them. This basis is Algorithmic probability - or program length complexity. ^(probability)
(FN: Also called Kolmogorov complexity, M.D.L, M.M.L, Stochastic ^{complexity})
Solomonoff, Komparev, Chaitin ^{complexity} (in order of discovery)
Furthermore, ~~it~~ enables us to understand ~~the~~ several important ~~paradoxes~~ resolve many difficulties in our understanding of the inductive process - not for least use of which is to give us a good understanding of randomness.

2.14.97: TM (PAR)

17 7

~~ALP also~~ In addition to giving a detailed understanding of how Learning works, it also gives us a tree structure showing ~~how~~ what concepts can be learned from what more fundamental concepts. — much like the tree structure of species in organic evolution. There is an important difference, ^{while} ~~that~~ there is a unique evolutionary tree, there are many possible tree structures for various learning tasks. ALP enables us to evaluate them and tell which are more likely, which involve less computation cost.

Perhaps discuss basis of A.I. in Introspective Psychology:

How it adds 1 more constraint: ~~the~~ programmability: ~~that~~ it needed 1 more constraint: learnability. Until ~ 1960, there was no clear criterion for learnability. (make clear that I'm not talking about Valiant "learnability", which ^{posse} ~~is a~~ narrow ^{restriction on how} ~~learning~~ learning may occur.)

Discuss Sol 86, 90 in detail: what "all info is in P.D. really means." Just how hard are discovery of a ~~given~~ learning & soln. of an ~~env~~ env. problem can be "simulated".

2.22 2.17.96: In fact, SM may be a v.p. tsg. if I do it via G.A. as a sequence of "FUV" problems! This could be very good! In general, it may be possible to give TM a ~~problem~~ problem in this form, in a "sufficiently complex (interesting) environment", so it will learn interesting things, even ~~tho~~ I don't do much in it. way off "designing" t. T.S.Q. ∴ it's just a very good problem envt!

It might be possible to give TM a small set of diff or probs in various interesting ^{environments} ~~envts~~. (have TM do them sequentially or in ||) & end up being very smart, & have a set of mtd ^{are} ~~able~~ to work other diff ~~inv.~~ probs & even oz probs (done in proper way!),

01: 7.90: ^{Star} This approach to TM is like the "problem pool" approach to TSPQ design.

T. Problem pool seems to be a non-cl. problem in which TM itself would have to face its own administrative problems — e.g. how much time to spend on reorganizing old data ("recompression") v.s. how much ~~time~~ time to spend on new problems.

I think one big new "BREAKTHRU" is the realization that I might be able to get the "problem pool" approach to work using not the general "optimum" soln. to the OZ problem, but by using the much simpler "INV" approach to OZ probs! \swarrow OZI method
(SN) In fact, I haven't really worked out the details of the OZ_{INV} method:

One ~~works~~ works on the Optm. problem w. INV (Lsrch) for a while, then after making some progress, one compresses data ~~is~~ ^{loop to (1)}. T. Q's are: How long does one do INV before compression, & how much time is spent on compression? Maybe working on INV & compression in time — share of \sim ~~is~~ ^{is} dwell time can't be > a factor of 2 wrong.

Def.

My impression of what Def. inv (\equiv OZI) is: One ~~is~~ ^{is} ~~initially~~ ^{initially} hunts for improvements in solns in a hypersphere about pt. α in the search space. As better solns are found, they replace α . In the case of what look like multiple peaks, we might use > 1 " α " pts. & time share betw. them.

When I speak of "hypersphere about α " the radius is in terms of "conditional complexity of trial wrt α ".

I'm not so easy about how to "Lsearch about α ". One α way! Find some short codes for α & add bits (or "info") to them. Ordinarily, one will know "short codes for α " because α was created in Lsrch & it tends to have a "short code".

ULTIMATE More ULTIMATELY we will want TM to do OZI search to find an optimum all-over psm. for itself, that "goes up hill" as rapidly as poss. — So it will use OZI to transcend OZI!

Actually OZI may be "Adaptate" (to non-optimal) in the sense that it ~~can~~ ^{can} solve most any OZ problem — but it can use other search plans if I like, it may seem more appropriate & seem universal enuf.

On "branching" in OZI! Perhaps one should spend time on a branch or its pc & on its ΔG (linearized G). Somehow ΔG is ^{somehow} related to the negative cc penalty. ΔG is linearized "Good" & cc is linearized "bad".

2/17/97 TM (Part)

The constant of proportionality betw. cc & ΔG is unknown, hvr. ... can we do anything w.o. any knowledge of it?? Perhaps time to show a pc. ΔG? Well, this is dependent on "origin" of ΔG. Adding a constraint to all ΔG's will change relative time shares. → 13.01

Perhaps get ideas on how to do this from GA; i from SGA. — In GA, iterative (Also ALife) Review lots of strange techniques used. See how I can formalize them w. ALP & Optimize them. Re-emphasize "ALP is the final solution for all induction problems". ∴ it should (in theory) tell how to solve 82cs problems.

One view: I own a set of pts P₁ vs for: each has known: (G_i, P_i, c_i) from P₁ set, what to work on next, & how hard?

Q: Is P₁ related to t. "what to work on next" problem (which is normally used for INDU probs. only)

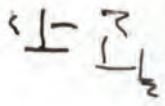
Is $\frac{\Delta G_i}{\Delta c_i}$ or $\frac{\Delta G_i}{\Delta P_i}$ like "velocity or momentum"?

ABCDEFIJKL
ABCDEFIJKL

- 1) List t. most immediate problems in OZI. Then suggest poss. solns.
- 2) Criticisms of those proposals.

ABC
ABCD
A
ABCD

- 2) Draw up a minimal (imple) system that could do OZI. See just how many & which features it needs in order for it to be able to improve itself & transcend, completely, its initial pgm. (Hvr, P₁ list may depend on giving it good problems, a good problem pool, or a good ISG.)



(SN): For Level of All kinds for t. Time-sharing method: Store all info used in each problem in a stack. When we stop work on a particular prob, we switch t. state of t. machine by switching stacks. The stack tells how far we've gone in a sub-prob.

We can have stack length limits in t. trials, as well as time limits.

Hvr, it's not clear how I should limit stack length as a fract of pc. If stack length is unlimited, then t. time limit limits it: I could try that: see if t. time limit was small enough to prevent stack from getting too big.

So, no stack length limit (to start off): but when I store t. stack, I have to say how long it is. Stack info can be stored in Ram, like Hard discs store files — not nearly contiguous, but w. jump positions somehow indicated.

One poss. way to do OZI: Pick the 100 best ΔG pts. P₁ vs for and do Level out from them by somehow "adding to their codes".

At the end of a certain time, pick the 100 best P₁ vs for & continue them } for a certain time. } loop to .37.

An impl. Q: If f being I'm looking for is a finite string, \forall perms for Perm will be "self-delimiting". How do I generate strings that are "close to" strings?
 say I know a short code for S .

Well! look at Sol 78 P 427 col II: section V on Chaitin's cond. complexity.

It uses Σ times w. Σ inputs. One input is f . min code for S .

$U(r, t^*) = S$: ~~the set of acceptable~~ The set of acceptable r , for each value of t^* , form a prefix set. t^* is the shortest code for r : i.e. $U(t^*, \Lambda) = t^*$.

If $U(\cdot, \cdot)$ has a unidirectional unidirectional input tape for first arg, & unidir. output tape, $U(\cdot, \cdot) = S$ means U stops after S is printed, Perm t .

~~legal~~ / legal first inputs must form a prefix set.

We write also constrain U so that $U(\Lambda, a) = U(a, \Lambda)$. Probably not poss. Set .20

Then $H^c(S/S^*) = \min |T| \ni U(r, S^*) = S$ so $r = \Lambda$ if $|\Lambda| = 0$

Hrr, if I want to use probs, maybe not so easy to get $H^c(S/S^*) = 0$. dist from S to $S^* = 1$.

.20 A bit of trouble! If $U(\Lambda, a) = S$, says then no other first input (other than Λ) is legal. $\sum 2^{-d(\Lambda)} = 2^0 = 1$: Λ is a complete prefix set.

I think my deriv. of cond. proby was (perhaps exactly):

.24 $P(S/t) = \sum_i 2^{-|r_i(t)|} \cdot 2^{-|r_j(S, r_i(t))|}$

The j th code for $(S, \text{ given the } i$ th code for t)

.25 $P(S/t) = \sum_i \sum_j 2^{-|r_i(t)| - |r_j(S, r_i(t))|}$

means the j th code for $(S, \text{ given } r_i(t))$

Can I show $P(S/S) = 1$? This could only be true if lengths of t exponents in .25 could be shown to be $(d(S) + 1)$ (lengths of some complete prefix set).

Maybe in .25 I should divide the denom by $P(t/\Lambda)$.

The expressions present ex expressions in .25 may give all t codes to produce (S, t) (in that order).

It would seem that t 2 strings referred to in t exponents in .25 have an order into them to produce S and t individually.


.37 .25 is gives something like f pc of $P(S, t)$.

Thinking about it, .37 is reasonable, but it is t pc of $P(S, t)$ in that order.


we have 2 strings, $r_i(t)$ and $r_j(S, r_i(t))$, and they, in that order, enable one to get (S, t) .

2/19/97 TM (Part)

2/11

So 10.25 gives + pc of + 2 objects s, t, in that order, using a certain kind of UMC. 

as a reference. A natural q: does $P(s, t) = P(t, s)$?

03 Also, does $P(s/s) = 1$? i.e. does $P(s, s) = P(s)$? (which would imply $P(s/s) = 1$)  (12.16)

For my own Applications in OZ I: Consider $U(x, r)$:

r is a (relatively short) code for some object, s. Now I want short strings, x, to modify r, so as to produce new objects, that are in some sense "Informationally close" to $U(r, \Lambda)$.


Seems like a good idea that $U(x, r)$ should have available to it $U(x, \Lambda)$. — on second thought, this seems rather trivial thing to satisfy.

I guess what is needed is a way for x to modify r to produce a new p.p.n. z; no = give output $U(z, \Lambda)$. ($U(z, \Lambda) = U(x, r)$).


Consider r: p.p.n. r, to be not a binary string, but r. (p.p.n.) of a function,

and the function is described by telling how to construct it from other functions.


Certain functions are functionals. Their inputs are various functions, & their outputs are functions. The "composition" operator is one kind of functional. Another classifies recursive functions (i.e. p.p. functs) from other functions.

 We may actually need only a few (maybe three) functionals, to get all p.p. functs. (I wrote about how to do this ~ 1990 in Saarb)

Perhaps what I had in mind for OZ I: we search until we get a set of somewhat good trials (say Ray are functions). We then pass a grammar thru these functs — But we want the grammar, in addition to be universal. The construction of this Grammar constitutes

 "Compression of the Corpus of 'Good trials.'" We then use this Grammar to Generate new trials.

-30

In SGA, we passed a grammar thru the data in a way so that the pc of each trial was either G or GG. — Hrr, it wasn't clear as to what we should do here: what  monotonic A function G we should make pc of to:

From a practical standpoint, we could make pc of trials being monotonic function of the "pc" of G. Grammar (or the G's of G. trials)

One way I thought of getting to Grammar was to use SGA or G (order) for each trial. This could result in a grammar that concatenated very many second where the sample trials were. I then used "(pc)" to control digests of new trials. Large G for concatenation of trials at high pc, small G for filling out the grammar at lower pc areas.

01 → Contrast 11.30 → 90 with the scheme I used for the II input Muxer.

I started w. a stock grammar on functions. When a function by G was found, it was added to the list of "defined" functions. I don't know what pc I assigned to those new functions. Perhaps I just took as inputs the set of ~~variables~~ "Good" trials and looked at the frequencies of the various "basic functions" within them.

A thing to think about: In SGA, I got that by making a large (essentially making all pc's basic or 1) that I could set a grammar to ~~generate~~ only trials near the "Peak" G . However, in .01-.06, I don't think there is trial .01-.06 was a "Bernoulli grammar" Each point had a fixed frequency. On the other hand, in SGA, I had the pc's vary w. each use of each branch pt. in the generation of the objects. I read not "Bernoulli Grammar" by a phrase str. Grammar (w. or w/o. context dependence).

16:11.03 : To actually use this $P(t/s)$ idea: Just list various useful functions & find ways to xpm them in to "nearby functions". Whatever ideas I have about "nearness", I can formalize so that $P(t/s)$ expresses those ideas quantitatively. The behavior of $P(t/s)$ is under my control (w. a few restrictions) — I will not be (very) much constrained by a priori ideas of what $P(t/s)$ should look like!

Q: Is it possible that I could use my own more optimal (So) 86, 90) system for this kind of TSCP? — or a Schmidt's system?
Essentially to $TM_2 = TM_1$, idea: Or just have to machine work on "external" problems, then alternate work on ~~my~~ its own system optimization.

It would be well to write up the basic idea — The idea of just how SM gives v.g. $T \neq Q^2$: In envt detail so that maxes of 1 million G moor 1 yr from now will be able to understand it. The main argt. of the present idea starts at 7.22

Essentially, the idea is, we get TM to climb an infinitely high hill. If he gets stuck in a local max, we have him climb to other local maxima — several of them — then we recombine the ~~the~~ concepts of these several local max's to try to get higher.

SGA has another method of dealing w. local max. — by using small δ , to fill out trials that are dist off from the known peak regions.

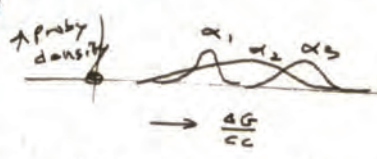
2/20/97 TM (P₂th)

$\sqrt{(8.20)} [\alpha_i]$

of: Spec 9.03 Say we've been working on several "α" pts (timesharing). We have this normal relationship betw p.c. & cc that we follow for each α pt. — How much time share should I give each α_i pts? Well, consider at each α_i the $\frac{\Delta G}{cc}$ obtained. ΔG is ^{that of the} hyst. pt. obtain, working out of that α_i.

Remember the goal is $\max \frac{\Delta G}{cc}$, where ΔG is the hyst. peak for that α_i. Well at each α_i, we have $[\frac{\Delta G}{cc}] \neq 0$. i.e. we're not really certain about the long term prospects of $\frac{\Delta G}{cc}$ at each of the α_i. We might work on the hyst. peak until its ΔG

Def



$\frac{\Delta G}{cc}_i (\equiv \dot{\Delta G}_i) \equiv \frac{\Delta G}{cc}$ at α_i $\dot{\Delta G}_i$ } Note! later we drop the "d" in both ΔG & $\dot{\Delta G}$

begins to get near small: Next, we work on the different α_i that has hyst. $\dot{\Delta G}_i$; The our choice may also depend on how far below the current peak we are. Perhaps choose the α_i that has least expected cc to reach the current peak. →

(4.0)

On the other hand, it maybe best to work on the α_i that has max $\dot{\Delta G}_i$.

For a given cc limit, we can choose α_i on the basis of both its $\dot{\Delta G}_i$ from peak and its $\dot{\Delta G}_i$. We choose the α_i $\rightarrow [\dot{\Delta G}_i, cc_0 + \Delta G_i]$ is max:

This is the expected ΔG we would have w. cc of cc₀ working on that α_i. Could we design an "anytime algorithm" so we wouldn't have to assume a particular cc₀?

If I work on the α_i w. max $\dot{\Delta G}_i$, then eventually I will hit the abs. peak, or I will end up w. identical $\dot{\Delta G}_i$ for ^{most} α_i. This is because usually $\dot{\Delta G}_i \downarrow$ as one works on it — so one switches to ~~some~~ ^{the} book as it gets \leftarrow that of the next highest $\dot{\Delta G}_i$. One then time shares betw. these 2 until their remaining $\dot{\Delta G}$ is \leq that of another α_i one then time shares ~~between~~ between 1, 3 — etc.

or In .19-.25, assume some fixed cc₀ and pick α_i viz. .21 and switch as $\dot{\Delta G}_i \downarrow$. In .21 both $\dot{\Delta G}_i$ and ΔG_i change with time:

$\int \Delta G_i dt = \int \Delta G_i dt$ so look at $\Delta G_i, cc_0 + \int \Delta G_i dt + \Delta G_i(\text{initial})$

Examine the effects of using cc₀ of various sizes: small, medium, large. Larger cc₀ (say cc₀ = ∞) leads to .26 ft (i.e. picking α_i of max $\dot{\Delta G}_i$) Small cc₀ leads to picking α_i of max ΔG_i .

Perhaps chose cc₀ = elapsed time since start, or some fraction ($>$ or $<$ 1) of that time. Another perhaps better way: do cc₀ = cc₀ for while then cc₀ ← cc₀ × 2 repeat "until done" (∞). Maybe a "while" = whatever cc is →

01:13.18 Some imp. ideas! 1) But in many applications, there is much uncertainty of G_i at \tilde{G}_i . See fig. 13.11 : 12. There is Noise in G .

2) In SGA: In fitting curves to $G(x_i)$: we may be mainly interested in the ordering of G by the Grammar: On the other hand, if G is a true Utility, it is defined within a constant factor.

3) The "input muxer" may not be such a "typical" problem. It has 2 features that make it ~~prohibitively~~ easier to solve than most:

a) It has a Noiseless G : (MTM)

b) Each solution ~~at each level~~ function at each level is a 2^n input muxer & the function itself is very useful for the next level: a 2^{n+1} input muxer. So TM doesn't have to search usable functions for common Subfunctions.

13: 13.40: So we: ^{t. problem of} w/ 13.0 ff ^{we normally} have uncertainty in both G_i & \tilde{G}_i .

Even in MTM problems, where we know G_i , our value of \tilde{G}_i is for past work & is not a perfect estimate of future G_i . We can, however, thru our past experience, collect data on G_i, \tilde{G}_i pairs off. past & how G_i (& \tilde{G}_i) continued (as a function of α).

Also look into 13.39 - 40 it seems to be not bad!

122 T. way it would work: we start w. $cc_0 = \delta$, which is small: we use 13.21 to pick the α_i to work on for $cc = \delta$. At first for small cc we work on α_i w. max G_i . Then from double cc so it = 2δ . we can use 13.21 again to decide what to work on. Then loop to α_i .

During the loop, the \tilde{G}_i of the α_i worked on will change (usually \downarrow).

But any way as $cc_0 \uparrow$, we will eventually start working on α_i 's that have high \tilde{G}_i but low G_i .

Well Ok., the process of .22 ff would get us "out of the hole" —

But we end up with largeish cc_0 , & we may be in a position where

we'd want to start all over w. $cc_0 = \delta$ again!

Another thing! I had been thinking in terms of a few α_i 's:

Actually, each terminal pt. will be an α_i : we will have an enormous no. of Procs.

Furthermore, ^{an} extensions of α_i maybe identifiable (of an extension of α_j ($i \neq j$))

This increases the probability of the resultant trial, but I don't know how to give such

a pt. to hyper cc that it should get. [This is a common problem in all

of LSrch.]

In spite of all these drifts: OZI seems to work for input maps:
So try to flush out the details of that soln.

Re: SGA: Say we have found a stock grammar ~~that~~ ^{that's "not bad"}:
we can use that grammar to hunt for peaks efficiently. Also if ~~the~~ ^{real world} trials
are expensive, using the stock grammar is much cheaper ~~to~~ ^{way to} find by G
pts.

But how do we then try to improve the grammar? We can spend time
w. fixed data set, hunting for better grammar &/o make well chose
~~the~~ new trials, that help make grammar fit better, &/o ~~select~~
different structural forms &/o other modifications of the grammar.

← Maybe use
"50%
Soln".

In non SGA, **OZI**: I think one doesn't do anything very smart
in the search: After finding a new peak, most of the cleverness is in
modifying the grammar to fit to new hyper G region.

Actually what we want is a default grammar for each G level;
So we have $P(\alpha, G)$ ^{form} Prob is assigned to each α, G pair.

Here, ~~the~~ G can be part of α just another param of α .

so $P(\alpha, G)$ is just another stock grammar: Here
 $P(\alpha, G)$ usually isn't designed so that it's easy to find α 's
they have by expected G .

In non-SGA, **OZI**: It's a rather "local" Hill-climbing Alg. The only
novel part is that the idea of local search is "information distance (conditional
complexity)".

I've been looking for a ^{source for} sufficiently "rich" batches of problems. SM is one
possy. Another is Alg. problems or Game problems from Books or that I derive.

→ Another possey: Re: "SW Agent" that watches a person's work on a computer
~~tries to learn from~~ & learns how to help him do various things.

The person it interacts with can be quite complex, quite "intelligent".

But are the info channels betw. the human & agent "wide" and are the
problems in mode(s) accessible to the agent, potentially "Rich" env?

So, several probs (1) Getting a good problem pool: that returns a
+sp. that can lead to fairly intelligent prob. solns.

(2) Method of (run): straight OZI: Details of the ~~the~~ SGA, how it would
work. Will full OZ hill climbing be needed?

On t. 2 symmetry of distances between UMCS:

Say M_1 & M_2 are 2 UMCS: $\forall x, (M_1(x) = M_2(\alpha_{12}x))$ so α translates from M_1 to M_2 . To (try) to show α_{21} need not be longer than a fixed amount, Δ ,

Plan α_{12} : (Δ depends on M_1 , but not on M_2 or α_{12})

Given M_1 and α_{12} : M_1 tries all machines M_i in order, until it finds

one $\rightarrow M_i(\alpha_{12}x) = M_1(x)$ for a large set of x . It will do this using ϵ -search

over all x and if $M_i(x)$ doesn't converge in $\epsilon^{-1} \cdot |x|$ time $\propto 2^{-|x|}$, it will drop it & try a new x . At a given time T , we will have a large set of strings $[X]^T$ for which we know $M_i(x)$ converges: if we will know $M_i(x)$.

As soon as we find a string x in $[X]^T$, we try it on a ϵ of the machines M_i : $M_i(\alpha_{12}x)$ for time T , each.

For each T , we will have a set $[X]^T$ and a set of machines M_i such that we know $M_i(\alpha_{12}x) = M_1(x)$: $\textcircled{2}$ a set of machines M_i'

" " $M_i'(\alpha_{12}x)$ either failed $= M_1(x)$ or failed to stop in time T , but \rightarrow still a candidate for " $= M_1(x)$ ".

So, for every T we will have a set of machines M_i' $\textcircled{1}$ & $\textcircled{2} M_i'$.

Def Q: Does T exist so that eventually M_2 has to be in $\textcircled{1} [M_i']$?

Also, we are interested in finding $M_1(y)$ for some string y .

Not enuf: M_2 could get into $[M_i']^T$ but $\rightarrow [M_i']^T$ for $\log_2 T$

Essentially, we want to know $M_2(\cdot)$: so we want to know $M_2(y)$ for any y : so consider all $[M_i(y)]^T$ & $[M_i'(y)]^T$.

Some of these will not converge in time T .

Radot Radot Actually, just consider the sum of these 2 sets: $\equiv \text{new } [M_i]$

Anyway, to Q is, ~~with parameters~~ for all $(\log_2) T$, will there be

$M_i \rightarrow M_i'$ ~~either~~ $(\alpha_{12}x)$ either converges correctly or does not converge, but we not know to how different from $M_1(x)$,

and there is at least one $y \ni M_i(y) \neq M_j(y)$ are known to

be different and: M_i & M_j are both in set $\textcircled{30}$

spac
01:25:40

One very simple way to do OZI! At each point in r. search, we move out from the previous peak only. (Noiseless Peak needed!). We simply do Lurch in distance from that peak. As soon as we find a higher pt., we use that as center for ~~the~~ Lurch. This is a very "Greedy" algm., & I imagine that it would go into local max's very easily (But because of the search Matrix being intermedial, I'm not sure that it's bad!).

One way to deal w. noise (i.e. get a less (apparently) Brittle soln.) is to follow ^{apparently} non-peak values - assuming the G's have noise in them. Say the prob. dist. of noise is $f(x)$ - so if we get G for a given pt., the probability ^{density} that it is really $G+x$ is $f(x)$. If we have several (3) micro local peaks, G_1, G_2, G_3 , using the $f(x)$ d.f., we can determine the prob. each is really the peak & use this prob. as a factor to ↓ the PC of excursions from that "local peak". It could estimate $f(x)$ if the noise is temporal - so I can repeat trials & get different G's.

Another ~~effectively noise~~ effect is given by the ^{Actually quite likely} (possy) that G is not a smooth function of "info distance".

➔ However, No Matter what the objections are: If I know the soln. to an IMV. problem, I have a good idea (upper bound) on how long the Lurch will take! So I have an idea of the "Smoothness" of G (or whatever) wrt. the Info distance.

On "S.W. AGENTS" as a poss. "Ingr. Pool", ~~EVOLV~~ Some Q's:

- 1) Just what kinds of agents? What kinds of induction problems do they solve?
- 2) ~~A~~ Is the pool of problems \rightarrow it has an adequate Ingr. sep. in it that is discoverable, \rightarrow that ~~can~~ can lead to by expertise,
 — a ability to solve really hard, interesting problems.

P(α, G) as
a kind of
Stoch Grammar.

One Ingr. Pool
used for a while,
then another, then
another correspond
to "Hybrid Vigour"

On a more general Q of TSPQ's, Prob. Pools:

- 1) Say we train TM on a small prob. pool in one area: then we use a small problem pool in a somewhat different area: would this give ~~us~~ very rapid ~~to~~ (ing: \approx like "Hybrid Vigour")?

The Hybrid Vigour, is more like taking 2 populations developed in 2 different problem areas, then mixing the populations & giving ~~it~~ a new, much harder problem. This last would be a bit like giving a population \approx ~~some~~ some what carefully constructed TSPQ

- 2) Perhaps we can take any set of ^{not necessarily closely related} induction problems \approx that are solvable ab. initio by OZI or GA or SGA, and use them as a esp. "By Problem" ~~time~~ ~~in~~ \rightarrow may ~~be~~ mean a small problem pool. So I ~~could~~ could take various problems that Koza solved using his "Genetic Pump". A dirty w. Koza's solns, hwn, is his ~~initial~~ initial choice of languages, which gives an enormous "Boost" to the process (lots of apriori). I couldn't ~~even~~ simply ~~to~~ train on one then ~~switch~~ switch to another — unless TM has some way to realize that certain ~~parts~~ concs. were more relevant to certain areas.

This last ~~is~~ ^{needs} a kind of "Mate learning" or hour lang. The usual way is that various areas of knowledge have names that have been given to them by the culture ~~and~~ \rightarrow scribbles Establishment: Assoc. in each "Name" is a set of ~~or~~ concs. that the culture has found to be appropriate to that area of knowledge ("Domain"). So TM can be, to some extent,

- .33 "To be" ~~what~~ ~~the~~ ~~it's~~ ~~into~~ i.e. be given ab-initio, the way to recognize
- .34 knowledge areas, along w. the set of appropriate concs for each area.

In some (unortho way) the info of .33-.34 has taken many generations ~~to learn~~ for the culture or sci establishment to learn.

Having been given the info of .33-.34, TM ~~can~~ should then learn to modify it in view of its own experience in prob. solving.

Another thing: TM may not have as sharp allocations of

22597 TM - pop.

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concs. to ~~the~~ domains is most critical, so he will be able to "mix" domains ~~more~~
easier.

On ~~the~~ OZI v.s. GA.

In OZI, one makes a grammar for each ^(fuzzy interval) region of G . $P(\vec{\alpha}) : \vec{\alpha}$ is a string.

More generally, consider to each grammar $P(\vec{\alpha}, G)$. We can think of this as either a different Grammar for each G , or consider G to be a symbol of the string, $\vec{\alpha}$ - possibly a special kind of symbol. $\vec{\alpha}$ can be any kind of object type, but it can have numerous & / or distinct parts. - so making G part of $\vec{\alpha}$ is ~~very natural~~ a very natural General.

In OZI, the way we approximate $P(\vec{\alpha}, G)$ is by fitting grammars for slowly \uparrow values of G - so we don't really try to curve-fit $P(\vec{\alpha}, G)$ along the whole seq. of G values simultaneously.

In SGA, I assumed that in stack grammars it was always easy to find pe's of h by p , by ~~then~~ doing h on the p 's. There are ~~some~~ a couple of cases where this is not true:

(a) In rewrite-rule grammars, w. probabilistic rewriting!



at level a we have 2 choices ~~for~~
~~at~~ At level b we have 2 ~~choices~~ choices for v . top & 100 choices for p . bottom.

If we use h by p we will choose the bottom in level a :

This gives $p \cdot .9 \times .01 = .009$:

But the actual p 's for v . top choice in a gives $.1 \times .9 = .09$,

for h by p of $.01$, the lowest: both are higher than $.009$

(b) In the Grammar that assigns p 's to functions, that I used on the Linear Max: Picking the most likely choice each time, would not always get the most likely function: \uparrow reasoning may be v to v_2 .

In v_2 , if one considers n levels ~~for~~ for making decisions, rather than just 1 level ~~at~~ then one gets a less "Greedy" v_2 - usually much closer to p or v .

01: On problem optimizing TM's ~~algorithm~~ "action algm" fast pass! This is Jovan's approach!

How can we defend it? Say we have many action algms each with its own mean G & its own pc. ("long R"). Which is "best"? should we time share μ_k , μ_m w. various wts? - what wts? : Say μ t. G 's are utilities (linear combination via probs).

02: This looks like the SM problem of evaluating "Signal" systems! - A problem that I've not really solved, - tho it seems easier than SM than

That's: General TM probem does it have this feature may make it easier to solve paradoxically!

in general, because of $R = u$ equations of \square probem is mult. yield in SM.

One has an opinion on desirability of "signal systems" w/o "action algms": but what are these to opinion of? Tho desir. how far out "t" system is in "info space" the order in which one would seek for it.

A way to look at it: Say one chooses a signal system a priori, then tests it on a "stationary" seq. The empirical yield is then an unbiased estimator of t. long term yield of that system.

Suppose one has 3 systems S_1, S_2, S_3 w. a priori μ P_1, P_2, P_3 $P_1 + P_2 + P_3 = 1$

Ratio yields are $Y_{1,2,3}$. $E \geq P_i Y_i$, t. on hand expected yield of t. system $\geq P_i S_i$? Well, $E \geq P_i Y_i \geq t.$ yield of $\geq P_i S_i$

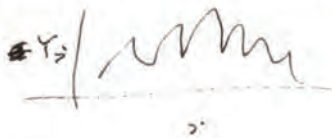
If $\geq P_i S_i$ is regarded as an a priori choice of system, then $\geq P_i Y_i$ is its

unbiased yield estimate.

35.10

This result seems to be Covari's Unbi. portfolg.! - which, to many seems not to learn fastenuf ! No! it doesn't do "rebalancing".

We might pick ϵ_i S_i w. max $P_i Y_i$. Hvr, if there are many ϵ_i & (maybe continuous?) then $P_i S_i$ can be very noisy in ϵ_i .



so picking a peak is not a vfg. way. Drawing a smoothly curve thru Y_i as a func of ϵ_i , a picking its peak, is usually much better.

Still, w. SM, applied, I'd have no good ideas of t. expected future yield of this peak. - on an even worse, I have no idea of its variance, so I don't know how much to bet. Tho in the SM case, the lower

can be one of the system params, & one picks it as part of the choice of peak!

This is beginning to look like an Lsach applied to O2 probs!

So we're looking for $S_i \geq P_i \hat{G}_i$ is max. One way is to do Lsach

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Trials: spend up to ~~10~~ cc; cc; on ~~the~~ trying system S_2 .

162 Hvr, there's another consideration: if more cc one spends on S_2 , the larger of the past corpus, one is able to consider, is the less error there is in the estimate.

So, the Learn problem becomes more complex! This looks very much like a general SM problem of deciding which timeseries to predict & which to use as data, & how much data to use!

It may be possible to get a ruff & dirty soln. to 02 ff that's good

→ Anot for Jurgen's general Schema!

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40 Betty James



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So: The TM₂ evaln. problem in Jurgen's Systems is very much like the evaln. of a "Signal" system in SM (31.08)

.87

117 So the TM project is now in 2 parts:

(1) Devising a good h.c. system for learning: working out all of the details of evaluating ~~it~~, defining the Govc.

(2) Creating, Maintaining, augmenting, combining "Problem pools":

To startoff, for (1) (the h.c. part) I could use any manner / sets & discrete system like G.A., S.G.A., Simanneal, ~~Moza~~GA, Lsuh OZI, LsuhOZ, etc.

Perhaps find good existing PSM.

Workout other details of (1), hvr.

For (2) start with SM because easy to get data: I'm familiar in E. field, & — but continue to work on ways to expand corpus into other domains.

Also, see if I can find ways to annex practically any kind of "problem pool" into the corpus.

Perhaps try to get it to learn to understand English used to desc. word problems in Algebra books. The ultimate goal being to be able to usefully read Algebra books.

Is there any way I can use ~~Levat's~~ "Cyc" info base?

If it learns English for desc'ing SM data: will this be of much use in learning English for Algebra problems? — or for probs in other fields?

On SM as a problem pool: I originally thought that I'd simply start w. ^{many} short sample
 say 100. & look for reg's, strategies. I'm not sure this will work: It will find high yield
 strats explicitly, but low yield strats would not be found. — They would contribute to noiser
 background.

Actually low yield strats are not of immediate practical value. If they are easy to find (A
 since they require long sequences to be evaluated.)
 They don't seem to be useful though they can be of use, however — as an aid for finding
 useful concs.

At ~~first~~ first, say, I'll look for strats in max $\frac{\mu}{\sigma}$ / yr. — w.o. contribution from Bldg, Slippage.
 If its strats, if it is very small & there are no options on it, stock, I ~~can~~ will probably find it
 of no ~~immediate~~ immediate practical value — but high risk value!

13 w. pure random noise as signals, I will find lots of concs w. by $\frac{\mu}{\sigma}$ / yr: k

But if I factor in the pc of trials, I should find a rather small product:

i.e. **If I** do m trials (w. pc of $\frac{1}{m}$ each) on random noise, I will expect
 $n \cdot \text{erfc}(\alpha)$ of them w. $\frac{\mu}{\sigma} > \alpha$. (I remain not "sure" but $Q(\alpha) = \int_{\alpha}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$.

→ What I want is a strat. w. yield that is least likely to be random noise. — $Q(\alpha)$ is related closely to $Q(\alpha)$
 $P(\alpha) = 1 - Q(\alpha)$

So for each strat. I divide the probab. of its occurrence by its probab. of occurrence by a
 random noise source.

But that is not what I'm considering: I already know that if strat. has a

$\frac{\mu}{\sigma}$ of \geq , say: But I have done m trials, so I expect a certain no. w. yields
 that far out, even if it were random noise. I think I want something

like maximum $\left[\frac{1}{m \cdot \text{erfc}(\alpha)} \right]$ i.e. $\max \frac{P(\alpha)}{\text{erfc}(\alpha)}$

Is the "Strong Law of Large Nos." relevant?

Suppose I have $n-1$ strats that are of yield ϕ (w. various σ 's) (say all σ 's = σ)
 I have 1 ~~strat~~ strat. w. $\mu = \mu$ (> 1) & $\sigma = \sigma$; $\frac{\mu}{\sigma} = \alpha$ (≥ 1 say). — for a given n ,

can I defeat it? Or, say all but 1 have $\mu = 0$ & various σ 's. One has $\mu \gg 0$, say.

In general, say we more or less know all the σ 's; but we don't know the μ 's.
 The data seems consistent w. hypothesis that all μ 's (but 1) are ϕ & one is an outlier

or its μ var (ly) is not ϕ . — essentially, what we want to know is: what's the expected

std of yield in the future of "f. Outlier"? — which is the problem ALP addresses!

The ALP is concerned w. Rel. of price (more directly)

⊕ perhaps, (for signal systems): Try to find diff. of near future

say 1 or 2 days as a function of near past.

A few Random ideas: 1) In signal systems: usually they are of the form: if $f(\text{near past}) > \alpha$
 then (buy). A better way to study such systems: What is expected future time response as a
 function of $f(\text{near past})$: In general, the response is $\left\{ \begin{matrix} \alpha \\ \beta \end{matrix} \right\}$ or is the low level characteristic
 of $f(\cdot)$ and σ .
 $\beta \rightarrow$ Time.

I analysed such "Signal Systems" in this way, some time ago. There is an optimum way to bet on — depending on noise level. One wants to complete bet soon, to get most yield per unit time: but B/S slippage tells one to wait for more yield of bet.

It would be well to find that analysis (In SM notes)

2) In 33.13 ff: I can assume a default d.f. — in which case, w. f. p. strat. I got a d.f. of yield about zero, w. a certain var, σ .

That I got a yield far out from this σ , suggests that ϵ f. ~~data~~ data is not t. default d.f. So I may be looking at this routine as rather taxing t.

strategy, or testing t. hyp.. That t. data is R.W. (Rand. Walk)

3) By using $\frac{\text{variation of price}}{\text{no. of trades}} \left(\frac{\text{no. of trades of strategy}}{\text{no. of trades}} \right)$ instead of time as basis, it may be poss. to get more uniform lookover of a strategy, year by year. If a year is very busy we end up w. a greater activity and $\frac{\mu}{\sigma}$ for 6. year should \uparrow . If we ~~not~~ divide

$\frac{\mu}{\sigma}$ by $\left(\frac{\text{no. of trades}}{\text{no. of trades of market}} \right)$ total variation of price 14 yr.

we may get to have that is much more uniform, yr by yr.

This is an. guide, but a better way to do it is to just look at f. seq. of trades & try to predict it — end of any renewal of structure of f. strategy

. 20

4) Actually, what I really want, is to ~~know~~ d.f. of t. yield of t. next trade

This would seem to be a pure prodn. problem in which ALP should be directly applicable.

One way to once approach this problem: I first tried to predict t. data stream — giving pc's for all poss. outcomes. From this, t. expected future yield of any strat. was directly computable.

Avr., ~~if~~ once I had t. near future D.f., I could probly devise a simpler

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strategy to take advantage of it!

It does tell something about t. structure of t. corpus, & should be possible to code that corpus.

Anyways, t. fact that a su. strategy has high yield, is a fact (conclusion) of t. corpus

it can be used to code it. (I most vacantly, did some work on this idea in Germany (in 1990) but also other times! My best Germany.)

One trick I had: say one has this R.W. default d.f. T. strategy makes \times μ / yr.

Def. R.W.

This makes it have a pc. of ϵ , which is rather small! So one knows it's

t. within a small fraction, ϵ of all poss. yr. corpi. This gives an μ pc (over default) of $\times \frac{\mu}{\epsilon}$. from this, one divides out t. pc of decn. of t. strategy itself.

\Rightarrow T. ratio is $\frac{\text{pc of strategy}}{\epsilon}$ is t. effective μ in pc obtained by this strategy, and ϵ (maybe) = its multiplicative yield.

So say one has a continuum of strategies. Should one do rebalancing of them (in dim space!) as Govt's Unli. portfolio? — Or simply give each one a w. not rebalance. (Remaker, covers system puts bad fast as no. of dim q!)

Actually this was more complicated: T. decn. of t. corpus involves both t. decn. of t. strategy & t. decn. of μ , its yield.

of μ for those continuous params. Also (if this is very imp.) one can make ϵ very small, If t. strategy has continuous params, we need to know t. d.f.

2:27:97 TM

If ρ is specificity k very narrow; so we want to know ρ density of t . default corpri at γ yield = k .
 This tells us how many corpri are specified, \Rightarrow function of k 's precision.

One more impl. candidate. Similarity betw. Evaln. of "signal systems" in GM & Evaln. of TM_2 's.
 In evaluating TM_2 's one often can't ~~run all of the data of t . past~~ — so one picks a short corpus: One does a similar thing in evaluating "signal systems" on part of t . data.

Perhaps there is an analog. of "Search/v.s. ~~frack~~ "frack" mod in Radar.
 First a ruff evaln. $\left\{ \begin{array}{l} \text{short corpus} \\ \text{Then a finer evaln. (longer corpus).} \end{array} \right.$

10:31:24 Going back to the "fund. idea" of 3.13-24: We chose a strategy w. prob p_i & we test it on t . corpus. Its yield is \bar{Y}_t — which is an unbiased estimate of its yield; the set $\{p_i\}$ gives the probab that I'd try $\{S_i\}$. So we could try the strat $\sum p_i S_i$ w. yield $\sum p_i Y_i$ — which is an unbiased estimate of its yield.
 Note this \rightarrow is not the rebalanced portfolio, but a simple ~~weighting~~.

If in $\sum p_i Y_i$, one of the $p_i Y_i$ is large and \bar{Y}_t is expected to be a v.g. strat.

\bar{Y}_t is an Honest unbiased evaln.

If $p_i S_i$ are set strats, & they are for futures, one could also automatically update \bar{Y}_t & leverage for ~~the~~ Kelly Betting. — so \bar{Y}_t yield would automatically take var. into account — & this ~~is~~ mult. yield would be an unbiased estimate.

Choose
 choose

This may be an answer to the Φ : In $p_i Y_i$: What is p_i the prob of ?
 It is to spec that one ~~would~~ ^{choose} choose the strat. S_i to test on t . corpus. If S_i has a complex structure, it's less likely to be chosen as a cand.

So, say one has a very large set of strats, S_i / i one tests them on t . corpus: Then $\sum p_i Y_i$ is an unbiased yield est. for $\sum p_i S_i$ & if one of the $p_i Y_i$ is quite large, or a group of strats $\sum p_i Y_i$ is large: then one may want to bet on it. Note that S_i will have expected yield of $\sim p_i Y_i$ not \bar{Y}_t . Usually p_i is quite small so \bar{Y}_t has to be very large for $p_i Y_i$ to be > 1 ! (i.e. Y_i is mult. yield). So \bar{Y}_t would not be a v.g. way to find strats w. mult./yr. of < 2 , say! Unless one had many yrs. of data!

30

A poss. way to deal w. this: we find $\{S_i\}$ w. $\text{rank} \{p_i S_i\}$ ~~...~~ rather low rank.
 In fact — or, say the top $100 \{p_i S_i\}$ ~~...~~ or e. $\{S_i\}$ tested.
 We then look at the structures of these 100 S_i 's, make a low. rank evaln \bar{Y}_t by $p_i Y_i$.
 Use this to generate new S_i 's / i which have tests on a new corpus (new M time) & /o new "strats".
 (30) This is actually the way we in fact deal w. this: ~~...~~ p_i 's are actually used to choose new S_i 's to try.

39

There must be some way to deal w. this $p_i Y_i$: $p_i Y_i$ will ~~be~~ ^{rarely} > 1 .
 Only if we use lots of data & $\text{time} > 2$ (yr. unit) we make it! The actually p_i accounts for by yields by healy betting, i.e. $\exp(\pm \frac{\mu}{\sigma})$, per year.
 But still, we need at least $\frac{1}{\sigma} = 2$ to e^2 /yr or e^{20} for 10 yrs — so we could have 2 pc per strat. $> e^{-20} \approx 2^{-29}$: 29 bits.

Actually 35.30-.35 is rather disturbing! Its really the "Old Faithful" method! (Somewhat refined, but essentially "Old Faithful" - w. all of its ~~weak~~ inefficiency. 38.01

It does support Post covlos' ~~FD~~ TK may be very good; it was developed on ~~a~~ $< 1/2$'s data on 1 commodity: Phon tested on ≈ 20 commodities for 5 yrs.

This is essentially Old Faithful again & I should be poss. to do much better!

T. only thing I've added is loop making step at 35.30-35.

While ALP helps a little here, perhaps RLP would help more: That part of t.

activity of 35.30-35 is designed to reduce search cc. So we use up data in order to ~~save~~ save cc! But I'm not so all about t. mechanics of how its occurring in

30.30-.35. Say I had much more cc available: how would I do this better?

T. first stage of 30.30-.35 finds the "top 100" then makes a good search for them. Now this could have been done a priori - when t. initial $\{p_i\}$'s were assigned - in which case we'd have/have ^{perhaps much} pc assigned to this set of 100.

Ok. maybe we can do this: we have some loop in which we preliminarily assign to $\{p_i\}$'s to the $\{S_i\}$ set. Post hoc, we get the $\{S_i\}$'s and we then try to find better codes for the "top 100" S_i 's: If we find better loop codes, we could have used those codes for the original a priori. So this first stage of "top 100" selection

could be a "coding aid". It points out which S_i are v.g. & suggests that I ~~can~~ could find a way in which t. should have assigned them a much by a priori.

I'm not yet sure of 20 ft, but its an attractive idea! It does get back the main pc (or data) that we traded for cc. - So if we were able to do this, we don't need the extra data that "Old Faithful" normally requires

Essentially, we are doing the recoding of the "TOP 100" post hoc - (after the fact) we can do this only because we have a general approach instead of obtaining a priori - But in journals we don't get the "best possible" a priori, because of cc limitations.

So, in general, if one has a clearly stated, more or less ~~the~~ universal

loop for one's initial a priori, one can legitimately look at the results of tests on a corpus & modify the a priori thru further recoding, if ~~it makes sense~~ if certain regularities in S_i vs. P_i suggest "groupings" of S_i & one can, indeed ~~use~~ devise short codes to implement these "groupings".

Another way of looking at .30: A special case! One has a strategy

(i.e. a loop) a priori. One tries a strat.: Its not bad but could be ^(moderately) improved by modification, which ~~is~~ makes code ^(cheaper) more compact.

T. all over yield. a priori increases. - So its ok.

- Here, it is necessary to have the a priori / loop or some idea of it a priori, or also there is danger

of assigning exactly by pc tot. modulus. of t. strat.

The forgg. is also true of modulus. of scientific theories.

36.35 It's a way of legitimating fudging a strategy. (fudge can ↑ or ↓ apparent yield)

I don't feel that I completely understand v. forgg. discn., but it does seem reasonable.

OZ probs by Lsach

The idea of "compression", is to modify a prior in view of the new data. But it will also involve "improving" the p.d. by finding more/better regys. in the corpus.

How. when we modify the p.d., we also want to do this to realize heuristics:

So the p.d. is modified in a way so that we expect it will help us solve more rapid ΣDG_2 per unit. time ~~(NOT)~~ (SOG is ~~to be~~ a corpus of other several future problems)

If we are working OZ problems, the p.d. will bias toward h.c. methods that were useful in the past, but it will (usually/inevitably) extrapolate to the set of "good hill climbers" by giving proper ~~some~~ weight defining how concs. & giving different wts. to old concs. that have been (or could have been) useful in generating empirically good h.c. methods.

On the "Honesty Principle" This is an idea that if you apriori choose a strategy (or arby function of strings) & apply to do a part of a sequence, then the value of that strat (or function) obtained is an unbiased estimator of the value if that strat. is applied ~~any~~ at a random pt. in time.

Now the forgg. is true if the seq. is stationary. But in general, the seq. is

•27 not stationary: say its v. output of a stoch. source w finite mem. Well, in this case, its an unbiased estimator, but for only that same time-point. Of course, if the seq. is stationary, all time pts are statistically equivalent - so its true for all time pts. From .27 ff it would seem that if the seq. is not stationary unbiased estimator of ^{application stat.} t₁ would not apply to measurements at t₂ unless t₁ = t₂.

•2797 There's some later work on non-stationary "honesty": considers the ensemble from which the corpus come.

01:36.02 On evaluating "strats", ect. One way to get a narrow spread, is to process a lot of data related to the area of interest. Again, this might be done post-hoc. - i.e. after one gets a strat, one can try to find reasons why it was off by a pip, (by looking at errors or parallel data - also by recoding (finding new regys in) ~~existing~~, "known", data.

On the Q. of (in SM, say) ~~then~~ trying strats on short seqs of data, because its fast. Say one finds a strat. w. high yield per unit time, but only 20 data pts. The expected yield should be small, because small size, \therefore large rare, even if yield is high. Hvr., this idea has to be worked out Exactly.

For SM, I ran (perhaps) fast a strat by using it as the default d.p. - to get its default variance. This tells how unlikely that particular yield is. If I divide this yield by the no. of trials to get it, I obtain a ruff figure of merit.

A better way is to ~~use~~ ^{divide} pc of strat by ~~the probability of such a yield~~ \rightarrow the unlikelyness of such a yield.

One (rough) way to get this ratio is to compare yields of actual SM strat. for both default d.f. and tried SM data, simultly. Every time one makes a ~~run~~ ^{on stock} run using the strat, one makes a run using single strat on "default" d.f. data.

If one takes k samples, it will take k times as long as testing strat alone. Hvr., if one ~~uses~~ ^{computes} the default yield proportionally it can take an amt. of time in d.p. of computer length. If units be possible to ~~do~~ ^{do} this analysis automatically, but it is likely to take much longer than 10 parallel default runs, if the total amt. of data ~~is~~ ^{per} run, is small!

It is easy to devise betting schemes w. high variance & zero mean: A classic case: Betting on coin flip: ~~start~~ I start w. \$1 bet. α If I win, have bet is \$1. If I lose, have bet is twice what I lost. loop to α I think the var = 0.0 in this case; Other schemes probly have $< \infty$ var., but very high.

If one uses $k=1$, the ratio of ~~the~~ SM yield to default yield can be very large, because default yield was very small. For $k=1$ this becomes much less likely. The d.f. of the mean of the sum of k Gaussians is well known & d.f.

for $k=1$?

one can get the first & second moments of the d.f. for $k=1$ & then do it for $2k$.

Probly X is betw $X = X + \Delta$ is Δ

" $X^2 = X^2 + 2X\Delta + \Delta^2 = \frac{1}{2} \frac{d}{dx} = \frac{1}{2} \frac{d}{dx}$

$X^2 \dots X^2 = (X + \Delta)^2 = X^2 + 2\Delta X + \Delta^2 = 0$

$\int \Delta^2$ is integrable with Δ^2 .

Any way, using some n value, (say $n=10$) we obtain a bunch of stats w. by apparent $(\rho c / \text{prob of that yield})$. We then $\uparrow n$ a/o \uparrow length of data to \uparrow precision.

.03 ① What is the best way to choose an initial k & continue the investigation?

② How do we choose a reasonably good stock, long, for this?

③ After we have been running for a while on the stock long of ② - how does mobility of long, in view of our experience?

.06 ④ How much cc. ~~of~~ was spread on ③? (50% ~~was~~ is a default value).

.03-.06 is a rough outline of the problems!

Before even ① The long, for generating stats, must be able to specify 1 Time Series \Rightarrow then specify 0 or more other T.S.'s to be used to help predict δ - first/1: (As well as δ - algorithm for how to relate δ data to δ bars to be made.)

.20: 34.26 A sort of optimum way to evaluate stats! First do an analysis of the t.s. being bet on - obtain a $\dots \Rightarrow$ 34.20-26 That rather than "strategies" getting the d.f. of the future of δ sep. would be about as good, & probably a lot more tractable, statistically! Other than the fact that a strategy is a kind of prediction, I'm interested in "signal systems" because ~~they are~~ similar to the general problem of evaluating "courses of action" - which is a very general, very common problem occurring in R.W. In SM, it is somewhat simplified because of the existence of the random walk default d.f.

.28 SM The default d.f. in t.s.-M. isn't really that clever i.e. the size of the δ can vary: One way to simulate it! At every tick, the default signal is δ has same Δ as true δ 's, but is of random sign. We can do this w. $\Delta T = 1 \text{ day } 1 \text{ wk or } 1 \text{ yr. also.}$ I don't exactly know how good a default model this is: It has some "total random" δ & s. SM. The model of .28 would be easy to make into a M.C. default for 38.20 ff.

Here, Note that the general "course of action", ~~is~~ knowing yield is a constant on corpus, is a "code" or "partial code".

.34 Actually the SM model of 38.20 is not entirely unrelated to other action functions evaluations in non-SM. Just as a strategy having a certain yield, may be regarded as a constant on the corpus, it is a part of a "code". Similarly a known value of an "action function" can be regarded as a "partial code" used to obtain probs. \Rightarrow perhaps some direct info about expected values of the "action function" in the future.

One class of algos I wite consider is, not prediction of a single stock from data of many stocks; but rather a switching system. This Mite be equit. to being able to predict, say 10 stocks & merely switching to ~~one~~ one w/ max expected \uparrow for the time period of interest.

Preprocessing: pick stock of hy varc. (or ~~for~~ not funds of hy varc.)

Hvr., for a near-term goal, just good predictions for next day, say, would be fine. say we want max $\frac{\mu}{\sigma}$ per bet; w. yield/yr. of $\frac{\mu}{\sigma} e^{\frac{1}{2}(\frac{\mu}{\sigma})^2}$ /yr. (n = no. bets/yr.)

Perhaps mult. this by ρ of strategy (\equiv predia. method) to get score \leftarrow ?

Actually, $e^{\frac{1}{2}(\frac{\mu}{\sigma})^2}$ is probly of an excursion of A , if true mean is zero... so that's fine! so $\rho / e^{\frac{1}{2}(\frac{\mu}{\sigma})^2} = \rho \cdot e^{-\frac{1}{2}(\frac{\mu}{\sigma})^2}$ \rightarrow just fine!

Well, .12 is a bit strange! I put that it would give relative distribution of various stocks, but it seems to depend on "n": Say we have several stocks in the same time length for their data: so same n for each. Then if we let unit of time be

2 yrs instead of 1, n doubles, with which changes relative effect of σ .

2 factors n .12 = : $\rho \cdot \left(e^{\frac{1}{2}(\frac{\mu}{\sigma})^2} \right)^n$

~~prob~~ probably leave out the n factor!

$e^{\frac{1}{2}(\frac{\mu}{\sigma})^2}$ is ρ probly of that random

excursion from mean of zero: indep of n: then ρ consider ρ cost of strategy

So criterion should be $\rho e^{\frac{1}{2}(\frac{\mu}{\sigma})^2}$

Here we do a "simplified version" of 38.20ff, perhaps: The ρ is obtained, ~~or~~ by observing the strat. in action, rather than observing the strat. operating on the default corpus.

I guess .12 is best for ~~the~~ short data strings. For the very long data strings .26 is not bad, but an analytical substitute: column of $\rho \cdot \rho^2$ of ρ stat. on default d.f. is then perhaps fastest (because its cc. is indep of corpus length).

Actually .27 is not quite right. .26 is not a simplified version of 38.20:

It's a somewhat different way to estimate ρ future d.f. of ρ future or ρ significance of observed deviation.

One way to look at ρ criterion .26: If one has several items (how big is ρ wt.

≥ 55 ac. w. each. We ~~must~~ must also consider the default item: It's of zero mean & ρ best with ~~with~~ ρ for ρ data. If a item has much more wt. than the default item, ρ drop the default item. If not, then we must consider default item in making wtd.

predns of all items tested.

Hvr. in .26: $\frac{\mu}{\sigma}$ for ρ (non-yearly) data string tells us how likely it is to have $\mu > 0$; But we also really want into yearly μ . for a $\frac{\mu}{\sigma}$ obtained in T time T yrs:

w'd like $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ as something of interest. Trouble is, if $T \ll 1$ yr, $\frac{\mu}{\sigma}$ will be $\ll 1$

SAC 46.02

3.11.97: TM : "Signal System Evolv (more generally "course of Action Evolv")

spec of 39.40 Back to the problem of "Signal" system evolv: or more generally "course of Action" evolv,

I think the "coding" idea of 39.34-40 is v.g.; I think I have a good genl way to deal w. practically all practn problems. If I can map the essential features of the "course of action evolv" problem into the coding problem, I may solve it in a very useful way.

Consider that finding a hy word strategy is equivalent to a kind of code. One has, always, to a priori / ^{p.d. out.} sequences. A single (or > 1) code of that sequence, constrains it.

Well: a single code actually falls what to sep. is! If it's an ALP-type code, random runnings of it give a p.d. for the future.

So, in this sense, a code for part of a corpus, gives a new p.d. out. rest of the ensemble.

A hy word strat, in this sense, then, is also a constraint on the entire ensemble.

3139: It's like to find that work & did in Seeb: ~1990: on signal systems; part of the idea was some thing like not really coding a corpus, but coding an "ensemble" of corpus.

Probably would be a good idea to state just exactly what the "Sig Sys evolv" prob. is:

Various approaches that have been made: Which ~~is~~ sounds unsuccessful & why (w. refs. if any) promising is why find them.

WAGNER 1) $\sum_{i=1}^{\infty} \sqrt{p_i} p_i$ may not converge.

T. proof of 44.10 fl. clarifies term. a lot of gas is extra power

2) Its the fastest way to do a M. Carlo search: i.e. if p_i is the prob of success on trial i ,

Then say we chose i w. prob $f(p_i)$. The expected time to success is

ca $T = \sum_i p_i \cdot f(p_i)$ *wrong!*

Say both $\sum p_i$ & $\sum f(p_i)$ converge & = 1.

We want T to be minimum. Invariance Is there a solution?! It may be that in general, we have to know the form of f_i as a func of i before we can find $f(p_i)$ for any one i ! One approach is χ calc. of variations:

$f(p_i) \rightarrow f(p_i) + \epsilon h(p_i) \Rightarrow h(p_i) = 0$ is my constraint $h \rightarrow$ so it will not

Then $T = \sum_i p_i (f_i(p_i) + \epsilon h(p_i))$ *disturb* Normaliz. of $f(p_i)$.

$\frac{\partial T}{\partial \epsilon} = 0$ for all $h(p_i)$ satisfying

so $\sum p_i h(p_i) = 0$ for all $h(p_i) \Rightarrow$ long as $\sum h(p_i) = 0$.

That tells us nothing about $f(p_i)$!

So how did I get $f(p_i) \propto \sqrt{p_i}$?

20 \rightarrow $T = \sum_i \frac{p_i}{f(p_i)}$ *wrong!*

Try $\epsilon h(p_i)$ again: $T = \sum_i \frac{p_i}{f(p_i) + \epsilon h(p_i)}$

w. $\sum \epsilon h(p_i) = 0$

$\frac{\partial T}{\partial \epsilon} = \sum_i \frac{p_i \cdot h(p_i)}{f(p_i)^2 + \epsilon h(p_i)^2} = 0$ for all $h(p_i)$

Is it true for $f(i) = \sqrt{i}$? $\sum \frac{p_i h(p_i)}{\sqrt{p_i} - h(p_i)} = 0 \Rightarrow \sum \frac{h(p_i)}{\frac{1}{\sqrt{p_i}} - \frac{h(p_i)}{p_i}} = 0$

from 20 it becomes more wants $f(p_i)$ to be large for

If $f(p_i) = p_i$ then $T = \sum_{i=1}^n 1 = n$

If $f(p_i) = \sqrt{p_i} \cdot k$ we get $T = \sum_i \frac{p_i}{k \sqrt{p_i}} = \sum_i \frac{\sqrt{p_i}}{k} = \frac{1}{k^2}$

$\sum \sqrt{p_i} = \frac{1}{k} \left(\sum p_i = 1 \right)$

So, is $k \geq 1$? \forall all; $\sqrt{p_i}$ is always $\geq p_i \Rightarrow \sum \sqrt{p_i} > 1$ so $\frac{1}{k} > 1$

So $k < 1$. \therefore so $\frac{1}{k} > 1$ and $\therefore T > 1$ *! T > 1 anyway even if $\alpha = 0$, $T = \frac{n}{k_0 \cdot k_1} = \frac{n}{1}$ is $T < n$, not $T < 1$*

but for any power of p_i , $\sum p_i^\alpha$ is finite $\forall \alpha < 1$, then

TV \Rightarrow M. calculus, if $\alpha > 0$, $T < n$

39 say $f(p_i) = k_\alpha p_i^\alpha \Rightarrow \sum \frac{p_i}{k_\alpha p_i^\alpha} = \sum \frac{p_i^{1-\alpha}}{k_\alpha} = \frac{1}{k_\alpha \cdot k_{1-\alpha}}$

$\sum p_i^\alpha = \frac{1}{k_\alpha} = \frac{n}{k_\alpha} \Rightarrow k_\alpha = \frac{n}{k_\alpha}$

for $\alpha = 1 = \frac{1}{k_1 \cdot k_0} = \frac{n}{1}$
" $\alpha = 0 = \frac{1}{k_0 \cdot k_1} = \frac{n}{1}$

If $\alpha < 1$, $k_\alpha > 1$ If $0 < \alpha < 1$, then k_α is $k_{1-\alpha}$ and $k_\alpha < 1$

" $\alpha > 1$ $k_\alpha < 1$ So $\frac{1}{k_\alpha \cdot k_{1-\alpha}} > 1 \therefore T > 1$

" $\alpha = 1$ $k_\alpha = 1$ If $\alpha > 1$ $k_\alpha < 1$, $k_{1-\alpha}$ is? $1 - \alpha < 0$.

" $\alpha = 0$ $k_\alpha = 0$ - could be $\frac{1}{0}$ or 0 if $\alpha = 0$

So maybe $k_\alpha < 0$ if $k < 0$ or $k > 1$.

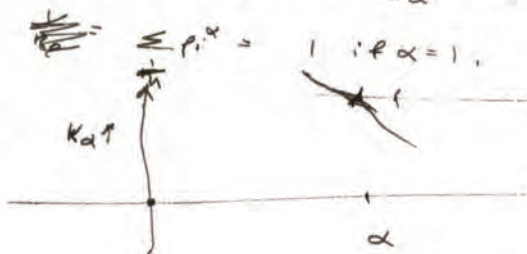
Say $P_i = \frac{1}{2^i}$ $i = 1, 2, \dots$ $P_i^\alpha = 2^{-i\alpha} \cdot k_\alpha$ $i = 1/\alpha$

$\sum_{i=1}^{\infty} (2^{-i})^\alpha = 2^{-\alpha} \sum_{i=1}^{\infty} (2^{-\alpha})^i = 2^{-\alpha} \cdot \frac{1}{1-2^{-\alpha}} = \frac{1}{2^\alpha - 1} = \infty, \text{ if } \alpha = 0!$

Clearly a ugly problem; but I do seem to indicate that I was ~~wrong~~ very wrong about this matter!

No! See 26

From 42.35: $\frac{1}{k_\alpha} = \sum P_i^\alpha$; If $\alpha = 0$ $\sum P_i^\alpha = \sum_{i=1}^n \frac{1}{2^i} = 1 - \frac{1}{2^n} \rightarrow 1$ as $n \rightarrow \infty$
 $\frac{1}{k_\alpha} = \infty$



If $\alpha < 0$? $\sum (2^{-i})^\alpha$ diverges.

So I guess $\frac{1}{k_\alpha \cdot k_{1-\alpha}} > 0$ unless

I think I still have this error: $T = \frac{1}{k_\alpha \cdot k_{1-\alpha}}$ it

$T = 1$ at $\alpha = 0$ or $\alpha = 1$; it's likely to have a min for $0 < \alpha < 1$! for α outside $(0, 1)$, it diverges; it's likely to have a min for $0 < \alpha < 1$!

In p2 $\frac{1}{k_\alpha} = \frac{1}{2^\alpha - 1}$; $\frac{1}{k_\alpha \cdot k_{1-\alpha}} = \frac{1}{2^\alpha - 1} \cdot \frac{1}{2^{1-\alpha} - 1} = \frac{1}{2^\alpha - 1} \cdot \frac{1}{2 \cdot 2^{-\alpha} - 1}$

$= \frac{2^\alpha}{2 - 2^\alpha} \cdot \frac{1}{2^\alpha - 1}$ ← empirically, this seems to have a min. at $\alpha = .5$!

Well, it would have to be at .5 since f. funct. is symmetric about $\alpha = .5$!

so all is needed is to show that f. funct. at .5 is < than near by ~~points~~ It's rather hard when $0 < \alpha < 1$.

In 29, $\frac{1}{k_2} = 1$ but that was not the problem!

T. q. is was it $< n$!

So, did I prove this for all forms of P_i ?

$\sum P_i^\alpha \cdot \sum P_i^{1-\alpha} \equiv \max$ perhaps we had to be better. $0 < \alpha < 1$.

Actually, the divergence of ~~$\sum P_i^\alpha$~~ $\sum P_i^{\frac{1}{2}}$ is not so badly diverging.

If we use $\alpha = 1$ or $\alpha = 0$ (as is normally done) we are concerned w.

the divergence of $\sum_{i=1}^n \frac{1}{P_i} = \sum_{i=1}^n 1 \rightarrow \infty$. Which is much worse!

3.14.97 T.M.

from 42.34: $T = \frac{1}{\sum_{i=1}^n p_i^{-\alpha}}$ which is symmetric about $\alpha = .5$, so if $\alpha > .5$ $\geq \min$, it must be at $\alpha = .5$: This will be true, indep of how p_i is a function of $\{z_i\}$.

So if we can find $f(p_i)$ to powers of p_i , p_i^{\pm} is a viable soln.
 But are there other functional forms possible?

Look at 42.34: say we have $\sum p_i = 1$

$\sum f(p_i) = 1$ $\sum \frac{p_i}{f(p_i)} = \epsilon?$

Say we had function: use Lagrange's multi.

so, ~~test~~ looser: use an arbitrary f_i (not really a function of p_i)

Roller

$\sum p_i = 1$ so $\sum \frac{p_i}{f_i} = \min.$

$\frac{\partial S}{\partial p_i} + \lambda \frac{\partial f_i}{\partial p_i} = 0$ $-\frac{p_i}{f_i^2} + \lambda = 0$ so $\lambda f_i^2 = -p_i \therefore f_i \propto \sqrt{p_i}$!

So this proves it for all forms of p_i is a function of z_i ! So the theorem is stronger than one I was trying to prove.

Great! [It's quite General]

So: (1) T. theorem is true if $\sum_{i=1}^n p_i^{\pm}$ converges: or if n is finite

(2) T. convergence of $\sum_{i=1}^n p_i^{\pm}$ is a weaker condition than $\sum_{i=1}^n 1 = n$, which is needed for the "usual" method to converge - i.e. "choosing z w. prob p_i "

Paradox: While .10 ff seems to prove optimality for (r) , clearly,

choosing the z in p_i order is better, or at least as good. ~~T. theorem is~~

Reason of "paradox": ~~.10~~ assumes probabilistic soln. .25 is a non-probabilistic soln.

Also: no replacement.

Just 10% tax
 & 20% in 16% tax

Much work on this on 11.01.98:

Again an imp. idea: When I was in College, I would often find interesting, potentially important problems that I wanted to work on, but I couldn't spare time at that pt.

- So I wrote them down, but didn't work on them then.

Again, I have a situation: I often find interesting problems, but I more often work on them - ~~not~~ to the detriment of the Main Goal (TM).

At the present, I am even more pressed for time than when I was at Hage, but I haven't really taken it, ^{or} adequate seriousness!

Present tentative emphasis:

1.12 1) Without details of Lark for 11 input Muxer: I want to see if I think there may be some imp. theoretic. diffys that I have to look at.

1.15 2) Without the details of Lark for the 11 problem pool! In particular, how the corpus length of code & w. time spent. Also in 11 w. this, one looks at ~~shorter~~ ~~corpi~~ ~~for~~ ~~loss~~ likely stacks or corpus. of stacks. [stuff leading to 40.26 is reasonable!]

Next: the Q. of how to incorporate observed ragys into it. ~~code~~ "compression" of the p.d. - i.e. How do we modify the Ref. in view of the observed ragys thus far?

3) Think about J. Schmitku's ideas. He is being less el. than I am - can I use his ideas at all? Also, it may be that J. is rather naive about probty, & is still using random search & calling it "Lark".

4) 3.17.98 is a kind of summary of probs in 1.12 & 1.15

SN) Re: Lark for 11 input muxer: Putting and, or, not in the function pool will make a difference in the first equation or 2, but after that they will be won low pc's because the "If" will be almost entirely used.
3 input muxer $P(3 = 1+2)$

SN₂) On the problem of Lark's putting MANY solutions of previous problems in Memory & then getting low pc's for those solns, because there are so many: Humans normally associate each prob. solving method with a certain area of science - this is a lot. We can extend the relevance of a prob. soln. to other areas by using analogy & other coding constructs.

Also people learn to apply (ps. methods / Theorems) rather narrowly: They devise methods to suggest which terms to try: The "methods" can be based on Gross statistics or

.01 Can be based on very carefully designed "descriptors" → (50.01 for how to take advantage of this)

.02: $\frac{\mu}{\sigma}$ spac: ~~is~~ $(\frac{\mu}{\sigma} \ll 1)$ so even tho $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ may be large, there's a v.g. chance that ϵ observed μ is really not $\Rightarrow 0$.

I think tho $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ & "T" params are quite diffnt. Just what Gore do I want?

If $T \gg 1$, then ~~is~~ $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ is a good Gore. If $T \ll 1$, I'm not sure what ϵ Gore is. I guess something like Expected value of $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$. — But how to

Obtain this?

.08 One way: assume an aprpd for $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$.

Quick impress: Unless $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ somewhat > 1 , one can't tell much. — since $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ is say 1 or < 1

Then I think T must be $\gg 1$ before one has much confidence in its being > 0 .

Note $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}} \sim .25$ can still be very useful! $\frac{1}{2} (\frac{\mu}{\sigma})^2 = \frac{1}{2T}$: so you'd need $\sim \frac{16}{.25^2}$ of these

to get a yearly yield of $e^{.5}$ but only 8 gives yield of $e^{.25} = 1.284/\text{yr}$.

T. above remarks about $\frac{\mu}{\sigma}$ for SM are also true for more general "Action items" Evalns. i.e. in Non-SM applies

.20 Still, I need a Gore for $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ $T \ll 1$. — I wish to recall to get it using ϵ aprpd of $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ (.08).

What is wanted! $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ for small T. or as about as small a T as one

would ever use! for SM, this would be ϵ smallest T \Rightarrow

.25 $\left[\frac{1}{2} \frac{\mu^2}{\sigma^2} \frac{1}{T} \right]$ (to yield) is useful. I actually want ϵ value of $\frac{\mu}{\sigma} \cdot \frac{1}{\sqrt{T}}$ (.25)

The value can be obtained by assuming a aprpd for μ .

easy to have an aprpd for σ or σ^2 also? Probably: we can know μ w.

much accuracy w.o. knowing σ . Usually, hrs. Its easier to get σ^2 w. usable accuracy than μ (when μ is centered about ϕ — as is usually the case.)

So our wish assume σ is known fairly well & one wants to know ϵ ~~aprpd~~

ϵ distribution of μ in view of ϵ data = ϵ aprpd of μ .

Say ϵ aprpd is normal(ly) $\mu \sim \phi$, var = σ^2 . Suppose we actually know ϵ σ^2 of ϵ d.f. of which μ is $\hat{\mu}$ estimate of mean. So we are only interested in ϵ σ^2 of μ .

Maybe not so good to assume σ^2 is known; This screws up the discovery,

Since μ is normally given by ϵ aprpd; for a given value because that value \downarrow its var!

So to start w. d.f. for μ is σ^2 . At first let σ be indep. — let μ be

Gaussian: let σ^2 be distributed like $x e^{-x}$ or $x^2 e^{-x}$ or $x^{100} e^{-x}$

Another tack: $\frac{\mu}{\sigma}$ is some aprpd. Its most likely that $\mu \approx \phi$ & σ is fairly large.

In this case $\frac{\mu}{\sigma}$ is not normally distributed. (There is some talk about ϵ d.f. for x^2 having an integrable singularity at ϵ origin.)



Say we are doing prediction: The mean of f_i of $d.f.$ is about f_i of $d.f.$ of f_i signal itself. So, with an assumed σ^2 , we have to predict for the ensemble. So we have this approx for μ, σ .

SM SN If the SM trades are not Poisson in time, then the time between successive trades is correlated. This is an easy hypothesis to test. We can do a scatter in 2 or 3 dims of T_n as a function of T_{n-1} : also a function T_{n+1} to T_n . In fact, we can study the differences between trades as a t.s. & try to find its dimension using that oldie psm, from LeBaron et al!

We can also do this by plotting no of trades per 5 or 10 min / bar. (directly available from "tick data" s.w.)


Or to times of successive "trade sets".

The relationship of latest at (say) μ can be treated as linear to start with, then try simple non-linear (Taylor series), then neural nets ModelQuest, Fractal etc.

The goal here is to try to predict var, & \therefore optimal values.

"Start of day" and "end of day" would have to be treated in special ways

- One way: do constraint T_n w/ T_{n-1} as a function of time since start of day. So one gets one value each day. Then one does some "pooling" - depending on how yr. data looks.

Actually, the DF for σ^2 is not critical: probably one could use a gaussian!  We may ~~cut off~~ $\sigma^2 < 0$; re-normalize.

- x. TRB subtraction
- y. Div subtraction
- z. TXI subtraction

Use whatever is mathically most tractable!

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So approx:

$$P_\mu(\mu) \cdot P_{\sigma^2}(\sigma^2) \cdot \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{\sum (u-s_i)^2}{2\sigma^2}\right)$$

$\mu = \bar{s}_i$
 $\sigma^2 = \text{MS error w/ } \mu = \bar{s}_i$
 $\neq \text{not } \frac{\sum (s_i - \bar{s}_i)^2}{N-1}$

$$\sum (u-s_i)^2 = \sigma_{OB}^2 + \sum (u-M_{OB})^2$$

$M_{OB} = \frac{1}{N} \sum s_i$
 $\sigma_{OB}^2 = \frac{1}{N} \sum s_i^2 - M_{OB}^2$

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$$\frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{\left(\sigma_{OB}^2 + \frac{(u-M_{OB})^2}{2\sigma^2}\right)}{2\sigma^2}\right\}$$

if $\mu = M_{OB}$; the peak w.r.t σ is $\sigma = \sigma_{OB}$
for $\mu \neq M_{OB}$; the peak w.r.t σ is M_{OB} .
If $\mu \neq M_{OB}$; the peak w.r.t $\sigma^2 = \sigma_{OB}^2 + \frac{(u-M_{OB})^2}{2}$

with μ not set $M_{OB} = 0$. The d.f. w.r.t μ is

Gaussian Gaussian w/ var = σ^2 .

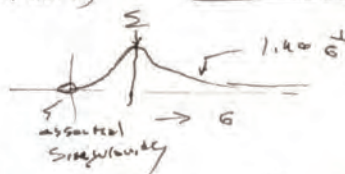
of the d.f. μ is at $\sigma_{OB}^2 + M^2$ (if $M_{OB} = 0$):

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look at d.f.

$$\frac{1}{\sigma} \exp\left(-\frac{x^2}{2\sigma^2}\right)$$

it's width $\propto \sigma$.
it's height $\propto \frac{1}{\sigma}$.
See 49.09



can be found by numerical analysis of the curve $\frac{1}{x} e^{-\frac{x^2}{2\sigma^2}}$

or $x e^{-\frac{x^2}{2}}$. we can find second moment about $x=1$ also find first order moments.

3 1797 TM

first 2nd moments can be found analytically.

Not so easy! They all diverge at $x \rightarrow \infty$ (including ϕ_{22} moments) 48

is $\int_0^{\infty} \frac{1}{x} e^{-\frac{1}{2x^2}} dx$ convergent?

say $u = \frac{1}{x^2}$ $dy = -\frac{dx}{x^2}$

$dx = -x^2 dy = -\frac{1}{y^2} dy$

so $S = \int_0^{\infty} \frac{1}{y^2} e^{-\frac{y^2}{2}} dy = ? \int_0^{\infty} \frac{1}{y} e^{-\frac{y^2}{2}} dy$
 diverges at 1-origin

OH! The S does of 47.37 does diverge: $\frac{1}{x} e^{-\frac{1}{2x^2}} \rightarrow \frac{1}{x}$ as $x \rightarrow \infty$,
 Hm, maybe ϕ_{22} d.f. for σ^2 doesn't diverge.

But one could just approximate this by using "5 power pts" - where prob goes down by $\frac{1}{2}$, say,

$S = \sigma^2 \quad ds = 2\sigma d\sigma \quad d\sigma = \frac{dS}{2\sigma} = \frac{dS}{2\sqrt{S}}$

$\frac{1}{\sigma} \exp\left(-\frac{S^2}{2\sigma^2}\right) d\sigma = \frac{1}{\sigma} \exp\left(-\frac{S^2}{2\sigma^2}\right) \frac{dS}{2\sigma} = \frac{1}{2} \frac{1}{\sigma^2} \exp\left(-\frac{S^2}{2\sigma^2}\right) dS$
 $= \frac{1}{2} \frac{1}{S} \cdot e^{-\frac{S^2}{2S}} dS$ which also diverges at $S \rightarrow \infty$!

Thus:

Grill Seas food
 Beach etc.
 2000.
 Mithras to dif.

Well oh: the $\phi_{22}(\sigma^2)$ factor can take care of the divergence of d or d^2 at ∞ .

We could approx. the d.f. by its width in x is d^2 increments! Then vice σ^2 down. Multivariate d.f.

in a direction x var is σ_{ob}^2
 " d^2 " " is $\propto \sigma_{ob}^2$. some fixed fraction of σ_{ob}^2 .
 y. d.f. of .19 - .20 is modified by the apnd $\phi_{\mu}(\mu) \cdot \phi_{\sigma^2}(\sigma^2)$ of 47.26.

$\propto \exp\left[-\frac{1}{2} \left(\frac{M-\mu_{ob}}{\sigma_{ob}}\right)^2 + \frac{(\sigma^2 - \sigma_{ob}^2)^2}{\alpha(\sigma_{ob}^2)^2}\right]$

say $\phi_{\mu}(\mu) \cdot \phi_{\sigma^2}(\sigma^2) \approx \frac{1}{\sigma_{app}} \exp\left(-\frac{\mu^2}{2\sigma_{app}^2}\right) \cdot \frac{1}{\sigma^2} \exp\left[-\frac{1}{2} \frac{(\sigma^2 - \sigma_{ob}^2)^2}{(\sigma_{ob}^2)^2}\right]$

α is some numerical constant < 1 which we will choose/determine.

So the problem is to get σ^2 moments $i \geq 2$ moments

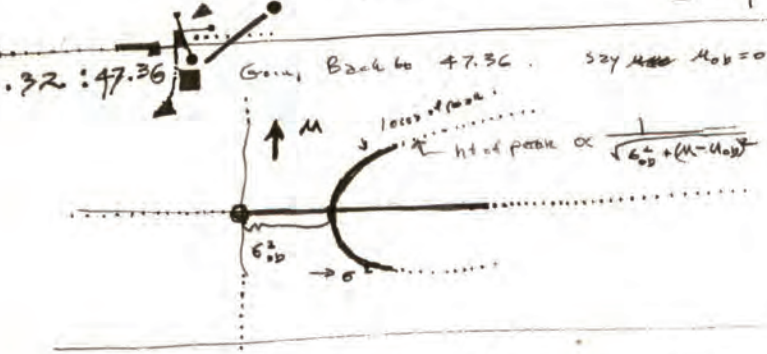
σ_{ob}^2 moments can be observed σ^2 of the signal w.r.t ϕ .

it is not so easy to guess: try various values α & see how things look! I may

try σ^2 large values, then reduce it in view of empirical results.

T. output param are to mean, μ & its width

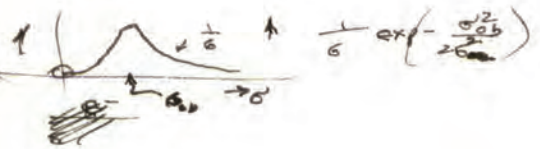
The σ^2 of the ensemble



then, on 47.37: d.f. $\frac{1}{\sigma} \exp\left(-\frac{S^2}{2\sigma^2}\right)$ diverges at $\sigma^2 = \infty$
 if $y = \sigma^2$; $\frac{1}{\sqrt{y}} \exp\left(-\frac{y}{2}\right)$
 this diverges even worse at $y \rightarrow \infty$:
 At σ or $y = 0$ there is no divergence because of essential singularity at origin

This sounds very strange! w.o. strong Apriod. the σ^2 p.s.p for σ or σ^2 is meaningless!

How this surface looks: along the line $\mu = 0$



Along line $\sigma = 0$ curve is Gaussian of width σ
 I don't know what = value contours are.

Since all σ moments of σ^2 are infinite, the mean var of σ^2 is to every impf. extent, determined by the spread!

Could it be poss. to find $\frac{M_1}{M_0} = \frac{M_2}{M_0^2} = \frac{M_2}{M_1^2} = \left(\frac{M_1}{M_0}\right)^2 \dots$ ($M_2 = 2^{\text{nd}}$ moment)

$M_0 \rightarrow \sigma$ as $\frac{1}{\sigma}$; $M_1 \rightarrow \sigma^2$ as $\frac{1}{\sigma^2}$; $M_2 \rightarrow \sigma^4$ as $\frac{1}{\sigma^4}$

$\frac{M_2}{M_1^2} \rightarrow \text{constant as } \sigma \rightarrow \infty$: The equality $\frac{M_2}{M_1^2}$ may converge.

$M = \frac{M_1}{M_0}$; $\sigma^2 = \frac{M_2}{M_0^2} - \frac{M_1^2}{M_0^2} = \frac{M_2}{M_1^2} - \frac{M_1^2}{M_1^2}$ which does converge!

$\frac{1}{\sigma} e^{-\frac{x^2}{2\sigma^2}}$ for large $\sigma = \frac{1}{\sigma}$ indep of σ^2 so this would be $-\frac{1}{2}$: Not at all reasonable!

$M_1 \rightarrow \sigma$; $M_1^2 \rightarrow \sigma^2$; $M_2 \rightarrow \frac{\sigma^2}{2}$ so this $= -\frac{1}{2}$: Not at all reasonable!

Important Grosser! The eq. should be $\propto \frac{1}{\sigma^n} e^{-\frac{x^2}{2\sigma^2}}$ (for n data pts).

for $n \geq 2$, M_0 converges

for $n \geq 2+2$, M_2 converges. So for $n \geq 4$ M_2 converges.

$\int_0^\infty \sigma^{-n} e^{-\frac{x^2}{2\sigma^2}} dx$ is probably a known well known S. = $\int_0^\infty \sigma^{-n} e^{-\frac{x^2}{2\sigma^2}} \frac{2x dx}{2\sigma}$

$\sigma^{-2} = x$ $\sigma^2 = \frac{1}{x}$ $2\sigma dx = -\frac{dx}{x^2}$ = $-\int_0^\infty x^{\frac{n}{2}} e^{-\frac{x}{2}} \frac{dx}{2x}$

$\sigma = \frac{1}{\sqrt{x}}$ $\int_0^\infty \sigma^{-\frac{n}{2}} \cdot \frac{dx}{2\sigma} e^{-\frac{1}{x}}$ $\left[\frac{1}{2\sigma} = \frac{1}{2}\sqrt{x} \right]$ $\int_0^\infty \frac{1}{2} x^{\frac{n+1}{2}} e^{-\frac{1}{x}} dx$

$\int_0^\infty x^n e^{-x} dx = n!$ p. 255 B.S. \leftarrow no!

$\int_0^\infty x^m e^{-x^2} = -\frac{1}{2} \Gamma\left(\frac{m+1}{2}\right)$ works! for $m > -1$; $n > 0$ B.t. $-2 < 0$

$m = n$, $n = -2$ This may be nice anyway! perhaps check it.

For $n = -2$, it may work: It should be $n \leq 2$

$\int_0^\infty \frac{1}{\sigma^n} e^{-\frac{x^2}{2\sigma^2}} dx = \int_0^\infty \frac{dx}{\sigma^n} e^{-\left(\frac{x}{\sigma}\right)^2}$ $\frac{\sigma^2}{2} = \beta^2$ $\frac{\sigma}{\beta} = x$ $\sigma = x\beta$

$\sigma = \beta\sqrt{x}$ $dx = dx \cdot \beta$

$= \int_0^\infty \frac{dx \beta}{x^n \beta^n} e^{-x^2} = \beta^{n+1} \int_0^\infty \frac{dx}{x^n} e^{-x^2}$ $\frac{1}{x^2} = y$ $-\frac{2}{3} \frac{dx}{x^3} = dy$

$= \beta^{n+1} \int_0^\infty y^{\frac{n}{2}} e^{-y} \cdot \frac{1}{2} x^{\frac{2}{3}} dy$ $x = y^{-\frac{1}{2}} \mid dx = -\frac{1}{2} y^{-\frac{3}{2}} dy$

$\frac{\beta^{n+1}}{2} \int_0^\infty y^{\frac{n-3}{2}} e^{-y} dy = \frac{\beta^{n+1}}{2} \left(\frac{n-3}{2}\right)!$ $x^2 = y^{-1}$

0.4. $1! = 1$ for $n \geq 2$ it converges, for $n = 1, 0, -1, \dots$ it diverges

for $n = 1$, $\frac{1-2}{2} = -\frac{1}{2}$ $(-\frac{1}{2})!$ diverges as it should $n = 2$ $\frac{2-3}{2} = -\frac{1}{2}$ $(-\frac{1}{2})!$ converges as it should

$n = 3$ $\frac{3-3}{2} = 0$ $0!$ converges $n = 4$ $\frac{4-3}{2} = \frac{1}{2}$ $(\frac{1}{2})!$ converges

$= \frac{\beta^{-n+1}}{2^{\frac{n-1}{2}}} \cdot 2 \cdot \left(\frac{n-3}{2}\right)! = \frac{\beta^{-n+1}}{2^{\frac{n-1}{2}}} \left(\frac{n-3}{2}\right)!$

$= \frac{\beta^{-n+1}}{2^{\frac{n-1}{2}}} \left(\frac{n-3}{2}\right)!$

$(-\frac{1}{2})! = \sqrt{\pi}$ $\Gamma(n+1) = n!$

$+\frac{1}{2}! = \frac{1}{2}\sqrt{\pi}$ $\Gamma(n+\frac{1}{2}) = 1 \cdot 3 \cdot \dots \cdot (2n-1) \cdot \sqrt{\pi}$

$+\frac{3}{2}! = \sqrt{\pi} \cdot \frac{3}{4}$ $\Gamma(n-\frac{1}{2}) = (n-\frac{1}{2})!$

37 $M_0 = \sigma$

38 $M_1 = \sigma^2$ for $n=1$

$M_2 = \frac{\sqrt{2}}{2} \sigma^2 \left(\frac{n-1}{2}\right)! \cdot \frac{\sigma^2}{2} \left(\frac{n-3}{2}\right)! = \frac{\sqrt{2}}{2} \sigma^2 \left(\frac{n-3}{2}\right)!$

50.09 spec

7:46.01 (SN) Part: Perhaps just have TM work induction problems in various fields! Say it starts in one area, learns to solve a lot of probs; then we give it probs in a new field that it learns to solve. — And so on w. many fields!

At first the cones for γ : Various fields are kept separate, but as time goes on, correspondences, analogies, are found betw. γ various fields \rightarrow so γ : p.d.'s for γ : various fields get "Mixed".

Hvr. It would be well to choose problem fields that were related closely, so as to encourage "Mixing" & actually speed up lmg.

09:49.40: So Now, I want to get exact equis. for Gork for search, w. ideas on just how to run the search; ~~to do~~ how fast do we & corpus (anyth U.S. try new corpi?)

Going back to 47.26: I expression should be:
$$\left(\frac{1}{\sqrt{2\pi}} \cdot \frac{1}{\sigma}\right)^n e^{-\frac{1}{2} \left(\frac{x - \mu}{\sigma}\right)^2}$$

$$= \left(\frac{1}{\sqrt{2\pi}} \cdot \frac{1}{\sigma} e^{-\frac{1}{2} \frac{\sigma^2 \cdot (n-1)k^2}{\sigma^2}}\right)^n \leftarrow \text{within } \mu, \text{variance is } \frac{\sigma^2}{n} \text{ should be } \left(\frac{n-1}{2}\right)! \text{ 2 k lines } \approx 90 pp.$$

$$\frac{M_1}{M_0} = \frac{2 \sum_{k=1}^n \frac{(n-1)!}{(n-2k)!} \sum_{k=1}^{n-1} (-1)^{k+1}}{\left(\frac{n-3}{2}\right)! \sum_{k=1}^{n-1} (-1)^k} = \sqrt{2} \frac{\sum_{k=1}^n \frac{(n-1)!}{(n-2k)!}}{\left(\frac{n-3}{2}\right)!}$$

$$\frac{M_2}{M_0} = 2 \sum_{k=1}^n \frac{(n-5)!}{(n-2k)!} = 2 \sum_{k=1}^n \frac{(n-5)!}{(n-2k)!} = (n-5) \sum_{k=1}^n \frac{(n-5)!}{(n-2k)!}$$

Go back to 18.32:
$$\frac{\beta^{1-n}}{2} \left(\frac{n-3}{2}\right)! = M_0$$

$$\rightarrow \text{which is } M_0 \rightarrow M_2 = \frac{\beta^{1-n-2}}{2} (n-2-3)!$$

$$M_2 = \frac{\beta^{-1-n}}{2} (n-1)!$$

$$M_2 = \frac{\beta^{-1-n+2}}{2} (n-2-3) = \frac{\beta^{2-n}}{2} (n-5)!$$

$$\frac{M_2}{M_0} = \beta^2 \frac{(n-5)!}{n-5} = \frac{\beta^2}{n-5}$$

$$M_1 = \frac{\beta^{1-n+1}}{2} (n-1-3)! = \frac{\beta^{2-n}}{2} (n-4)!$$

$$\frac{M_1}{M_0} = \beta \cdot \frac{(n-4)!}{(n-3)!} \approx \left(\frac{\sum_{k=1}^n}{\sqrt{n-4}}\right) \leftarrow 57.25$$

$$\Gamma\left(n+\frac{1}{2}\right) = (n-\frac{1}{2})! = \frac{1 \cdot 3 \cdot 5 \cdots (2n-1)}{2^n} \sqrt{\pi} = \frac{(2n-1)!}{2^n n!} \cdot \sqrt{\pi} = \frac{(2n-1)!}{n! 2^{2n}} \sqrt{\pi}$$

$$= \frac{(2n-1)!}{(2n-1)!} / \left(\frac{2(n-1)-1}{2}\right)!$$

$$\frac{(n-4)!}{(n-3)!} \text{ case 1: } n \text{ is even! } n = 2m \quad \frac{2m-4}{2}! = (m-2)!$$

$$\frac{2m-3}{2}! = (m-1-\frac{1}{2})!$$

$$(m-1-\frac{1}{2})! = \sqrt{\pi} \frac{(2(m-1)-1)!}{2^{m-1} (m-1)!} = \frac{(2m-3)!}{(m-1)! 2^{2m-2}} \sqrt{\pi} \text{ damn}$$

$$\frac{(m-2)! (m-1)! 2^{2m-2}}{(2m-3)! \sqrt{\pi}} = \frac{1}{\sqrt{\pi}} \frac{2m-2}{2m-3} \frac{2m-4}{2m-5} \dots$$

$$\frac{(n-4)!}{(n-3)!} = \sqrt{\frac{n-4}{n-3}} \frac{(n-4)!}{(n-3)!}$$

But for use stilling: $X \approx \sqrt{2\pi} X \cdot e^X$



$$\frac{x+\frac{1}{2}}{x!} = \sqrt{\frac{x+\frac{1}{2}}{x}} \cdot \frac{(x+\frac{1}{2})^{x+\frac{1}{2}}}{x^x} \cdot \frac{e^{x+\frac{1}{2}}}{e^{x+\frac{1}{2}}} = e^{\frac{1}{2}}$$

$$\frac{(x+\frac{1}{2})^{x+\frac{1}{2}}}{x^x} = \left(\frac{x+\frac{1}{2}}{x}\right)^x \cdot (x+\frac{1}{2})^{\frac{1}{2}} = \left(1+\frac{1}{2x}\right)^x \cdot \sqrt{x+\frac{1}{2}}$$

for large x: $\sim e^{\frac{1}{2}} \cdot \sqrt{x}$

$$\sqrt{1+\frac{1}{2x}} \cdot \sqrt{x}$$

$$\sqrt{4.5}$$

$$\sqrt{\frac{5}{x+1}} = \frac{\sqrt{5}}{\sqrt{x+1}}$$

$$\frac{x+\frac{1}{2}}{x!} = \frac{(x+\frac{1}{2})^{\frac{1}{2}}}{x^{\frac{1}{2}}} \cdot \frac{(x+\frac{1}{2})^x}{x^x} \cdot \frac{e^{x+\frac{1}{2}}}{e^{x+\frac{1}{2}}} = \frac{(x+\frac{1}{2})^{\frac{1}{2}}}{x^{\frac{1}{2}}} \cdot \left(\frac{x+\frac{1}{2}}{x}\right)^x \cdot e^{-\frac{1}{2}}$$

$$= \frac{(x+\frac{1}{2})^{\frac{1}{2}}}{x^{\frac{1}{2}}} \cdot e^{-\frac{1}{2}} \cdot \frac{(x+\frac{1}{2})^x \cdot (x+\frac{1}{2})^{\frac{1}{2}}}{x^x} = e^{-\frac{1}{2}} \frac{x+\frac{1}{2}}{x^{\frac{1}{2}}} \cdot \left(\frac{x+\frac{1}{2}}{x}\right)^x$$

$\sqrt{x+1}$ is better

$$\frac{(x+\frac{1}{2})!}{x!} \approx \frac{x+\frac{1}{2}}{x^{\frac{1}{2}}} = \sqrt{x + \frac{1}{2\sqrt{x}}}$$

P 272 B.S. for tables of $\Gamma(x), \Gamma(x+1/2)$

$\Gamma(5) = 24$	$\Gamma(5.5) = 52.3427$
$4! = 24$	$4.5! = 52.3427$
ratio = 2.1909	$\sqrt{4 + \frac{1}{2\sqrt{4}}} = 2.25$: not needed!

9! $\Gamma(10) = 3.6288 \times 10^5$; $9.5! = 3.1232 \times 10^5$; $\sqrt{9 + \frac{1}{2\sqrt{9}}} = 3.166$ better

9.5! $\Gamma(10.5) = 11.3327 \times 10^5$

50.25 : $\frac{M_1}{M_0} = \beta \frac{(\frac{n-4}{2})!}{(\frac{n-4}{2} + \frac{1}{2})!} \approx \beta \sqrt{\left(\frac{n-4}{2}\right)^{\frac{1}{2}} + \frac{1}{2\left(\frac{n-4}{2}\right)^{\frac{1}{2}}}}$

25. $\frac{M_1}{M_0} = \beta \sqrt{\frac{2}{n-4}} = \frac{\beta}{\sqrt{n-4}}$ since $\beta \sqrt{2} = \frac{M_1}{M_0}$

ITM: $\frac{\beta}{\sqrt{n-2}}$ is better $\rightarrow 52.02$

well: $\int_0^{\infty} \frac{1}{\sigma^n} e^{-\frac{x^2}{2\sigma^2}} dx$

as $n \uparrow$: σ gets smaller and smaller



because σ magnitude for σ is smaller.

$$\frac{M_1}{M_0} - \left(\frac{M_1}{M_0}\right)^2 = \frac{22}{n-5} - \left(\frac{2}{\sqrt{n-2}}\right)^2 = \frac{22}{n-5} - \frac{4}{n-2}$$

$$\rightarrow \frac{22}{n-5} - \frac{4}{n-2} = \frac{22(n-2) - 4(n-5)}{(n-5)(n-2)} = \frac{22n - 44 - 4n + 20}{(n-5)(n-2)} = \frac{18n - 24}{(n-5)(n-2)}$$

Which is the value of σ in the direction

32 doesn't seem right! The first moment, $\frac{M_1}{M_0}$ should $\rightarrow \sigma$ as $n \rightarrow \infty$.

also, I left out $\left(\frac{1}{\sqrt{\pi}}\right)^n$ factor

also, use $\frac{x+\frac{1}{2}}{x!} \sim \sqrt{x+1}$ is better approx (slightly better). $\sqrt{x+1} - e$ actually

Well! One error: 14 49 29 " Σ " should be " $n \leq 2$ ", also

That last seems to fix it up just fine! replace \leq w. $\sqrt{n} \leq$!

02:51:25 $\frac{M_1}{M_0} = \sqrt{\frac{n}{n-2}} \leq$ good! it shall approach \leq as $n \rightarrow \infty$.

04:51:26 $\frac{M_2}{M_0} = \dots$ I think this is wrong: it leads to estimate of $\leq 2 \rightarrow 3 \leq 2$ as $n \rightarrow \infty$.

The expression for $\frac{M_2}{M_0}$ is probably wrong

3:30

49:32 should be $\frac{\sqrt{2n}}{2} \beta^{-n} \cdot \frac{\beta}{2} \left(\frac{n-3}{2}\right)!$ we get values for $n, n-1$ in $n-2$

Then in M_1/M_0 & M_2/M_0 we substitute $\beta^2 = \frac{n \cdot 2}{2}$

Not quite! The $\sqrt{2n}$ factor doesn't change for the 3 cases — so it cancels, as does: $\frac{\beta}{2}$

so $\frac{M_1}{M_0} = \beta \cdot \frac{(n-1)!}{2} / \left(\frac{n-3}{2}\right)! = \frac{\beta}{2} \cdot \frac{n-1}{2} = \frac{2\beta}{n-3}$

$\frac{M_1}{M_0} = \beta \cdot \frac{(n-1)!}{2} / \left(\frac{n-3}{2}\right)! = \frac{\beta}{\sqrt{\frac{n-2}{2}}} = \frac{\sqrt{2} \cdot \beta}{\sqrt{n-3}} = \frac{\sqrt{2} \sqrt{n}}{\sqrt{n-3}} = \sqrt{\frac{n}{n-3}} \cdot \sqrt{2}$

$\frac{M_2}{M_0} = \beta^2 \cdot \frac{(n-5)!}{2} / \left(\frac{n-3}{2}\right)! = \frac{\beta^2}{\frac{n-3}{2}} = \frac{2\beta^2}{n-3} = \frac{2 \cdot 2}{n-3}$

Very bad: wrong $\frac{M_2}{M_0} = \left(\frac{M_1}{M_0}\right)^2 = \beta$.

3:04:22

$\frac{x!}{(x-1)!} - \left(\frac{x!}{(x-\frac{1}{2})!}\right)^2 = ?$ so how much does $\left(\frac{x!}{(x-\frac{1}{2})!}\right)^2$ differ from x ?

from 5:15 ff: $\frac{x+\frac{1}{2}}{x!} = \frac{(x+\frac{1}{2})^x \cdot (x+\frac{1}{2})^{\frac{1}{2}}}{x^x} \cdot \frac{e^x}{e^{x+\frac{1}{2}}} \cdot \frac{\sqrt{x+\frac{1}{2}}}{\sqrt{x}} \cdot \frac{(x+\frac{1}{2})^{x+\frac{1}{2}}}{x^{x+\frac{1}{2}}} \cdot e^{-\frac{1}{2}}$

$= \left(1+\frac{1}{2x}\right)^x \cdot e^{-\frac{1}{2}} \left(\sqrt{x}\sqrt{1+\frac{1}{2x}}\right) \cdot \sqrt{1+\frac{1}{2x}}$
 $= \left[1+\frac{1}{2x}\right]^x \cdot e^{-\frac{1}{2}} \cdot \sqrt{x} \cdot \sqrt{1+\frac{1}{2x}}$

So $\sqrt{x+\frac{1}{2}}$ but correction: empirically, $\frac{x+\frac{1}{2}}{x!} \approx \sqrt{x+\frac{3}{4}}$

$\ln(1+\epsilon) = \epsilon - \frac{\epsilon^2}{2} + \frac{\epsilon^3}{3} - \dots$

$\ln\left(1+\frac{1}{2x}\right) = \frac{1}{2x} - \frac{1}{8x^2} + \frac{1}{24x^3} - \dots$

$\left(1+\frac{1}{2x}\right)^x = e^{x \left(\frac{1}{2x} - \frac{1}{8x^2} + \frac{1}{24x^3} - \dots\right)}$

So correction to $\sqrt{x+\frac{1}{2}}$ is $x \left(1 - \frac{1}{8x}\right)$!! which is in wrong direction!

$x+\frac{3}{4} / x+\frac{1}{2} = \frac{1+\frac{3}{4x}}{1+\frac{1}{2x}} \approx 1+\frac{1}{4x}$ so it's not too wrong sign!

1:9
1:20

1:30

1:36

It would be poss. to check the Algebra step by step, for $n=4$, say, since the approxns. are very good & close!

$$\frac{(x+\frac{1}{2})^{x+\frac{1}{2}}}{x^{x+\frac{1}{2}}} \cdot (x+\frac{1}{2})^{\frac{1}{2}} \approx \left(\frac{x+\frac{1}{2}}{x}\right)^x \cdot \frac{x+\frac{1}{2}}{\sqrt{x}}$$

$$\begin{aligned} & \sqrt{x} \cdot \left(1+\frac{1}{2x}\right) \\ & \sqrt{x} \cdot \left(1+\frac{1}{2x}\right) \\ & \approx \sqrt{x} \sqrt{1+\frac{1}{2x}} \\ & \approx \sqrt{x+1} \\ & \approx \sqrt{x} \left(1+\frac{1}{4x}\right) \\ & x \left(1+\frac{1}{4x}\right) = \\ & \sqrt{x} \cdot \left(1+\frac{1}{2x}\right) \\ & \sqrt{x} \cdot \sqrt{1+\frac{1}{2x}} \end{aligned}$$

0.4. so Proof's it! The error was on 52.30!

$\sqrt{x} \cdot \sqrt{1+\frac{1}{2x}}$ instead of $\sqrt{x} \cdot \left(1+\frac{1}{2x}\right)$: The first order approx:

$1+\frac{1}{2x} \approx \sqrt{1+\frac{1}{2x}}$: so $\sqrt{x+1} \approx \sqrt{x} \sqrt{1+\frac{1}{2x}}$

The hyperbolic approximation .36 is $x \exp(-\frac{1}{2x}) \approx 1 - \frac{1}{2x}$

or $x \sqrt{1-\frac{1}{2x}}$ $\sqrt{x+1} \cdot \sqrt{1-\frac{1}{2x}} = \sqrt{x+\frac{1}{2}-\frac{1}{4}}$
 $= \sqrt{x+\frac{1}{4}-\frac{1}{4x}}$ which is $\approx \sqrt{x+\frac{1}{4}}$ to first order!

It seems to have less much precision somewhere, but especially its use.

18 $\frac{x+\frac{1}{2}}{x!} \approx \sqrt{x+\frac{1}{4}}$ → Sec 79.15 for moment formula:
 is \approx w. fun $\approx 2\%$ for $x=2$ and rapidly gets better for $\log x$.
 5% error for $x=1$ error $\frac{1}{2} - 1$; $\frac{1}{2} - 1\frac{1}{2}$; $\frac{1}{2} - 1\frac{1}{4}$
 2% error for $x=0$ $\frac{1}{2} - 1\frac{1}{2}$
 1% " " $x=15$ &

20 52.19 $\frac{M_2}{M_0} = \beta \frac{(n+1)!}{(n-3)!} = \frac{\beta}{\frac{n}{2} - \frac{3}{4} + \frac{1}{4}} = \frac{\beta}{\frac{n}{2} - \frac{3}{4}} = \frac{\sqrt{2}\beta}{\sqrt{n-\frac{3}{2}}} = \frac{\sqrt{n}\beta}{\sqrt{n-\frac{3}{2}}}$

52.20 $\frac{M_2}{M_0} = \frac{n-2}{n-3}$

$\frac{M_2}{M_0} - \left(\frac{M_1}{M_0}\right)^2 = \frac{n-2}{n-3} - \left(\frac{n}{n-3}\right)^2 = \frac{n-2}{n-3} - \frac{n^2}{(n-3)^2} = \frac{n-2}{n-3} \cdot \frac{n-3}{n-3} - \frac{n^2}{(n-3)^2} = \frac{(n-2)(n-3) - n^2}{(n-3)^2} = \frac{n^2 - 5n + 6 - n^2}{(n-3)^2} = \frac{-5n + 6}{(n-3)^2}$

$\frac{M_2}{M_0} = \frac{n-2}{n-3} \cdot \frac{1}{2(n-3)}$ ← ratio ↓ from $\frac{n-2}{n-3}$

$\frac{M_1}{M_0} \approx \sqrt{\frac{n}{n-\frac{3}{2}}}$
 $\approx \left(1 + \frac{1}{4} \cdot \frac{1}{n}\right)$ for large "n".

Is $\frac{n}{n-5/2}$ consistent? should not be an estimator of mean but biased? perhaps not!
 It's just a normalized first moment! The first moment Mult by $\sqrt{\frac{n-5/2}{n}}$ would perhaps be unbiased.

for large n, χ^2 d.f. converges to a narrow peak in the μ, σ plane:

width or direction is $\approx \sqrt{\frac{n}{n-5/2}} \left[\frac{1}{\sqrt{2(n-3)}} \right]$

width in μ " " $\approx \frac{1}{\sqrt{n-1/2}}$

so for large n, its width in μ term of σ direction by $\sqrt{2}$ factor.

For smaller n, we want $\geq \frac{\mu}{\sigma}$, which is known for Gaussian & 2 dim d.f.s.
 The Gauss. in μ, σ direction can cause trouble w. $\sigma = \phi$ (maybe, I'm not sure about integrability).
 The d.f. Γ on using has an essential zero (singularity) at $\sigma = \phi$, so this would cause no trouble in the case, but the integration may maybe suitably off!

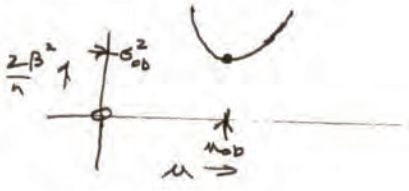
38 $\int_{\phi=0}^{\infty} \int_{\mu=-\infty}^{\infty} \frac{M}{\sigma} \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^n \exp\left[-\frac{n(\phi_0^2 + (\mu - \mu_0)^2)}{2\sigma^2}\right] d\mu d\sigma$

factor $\frac{1}{\sigma}$ causes unchanging int. expression. M seems to change things a lot.

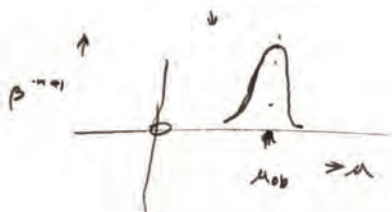
we can easily integrate in the direction.

02 $\frac{2\beta^2}{n} = \frac{(\sigma_{ob}^2 + (\mu - \mu_{ob})^2)}{n}$; $53.38 = \int_{-\infty}^{+\infty} d\mu \mu \cdot \beta^{-n+1}$ ~~...~~ $\left| \frac{1}{2(\sqrt{2\pi})^n} \cdot \left(\frac{n-2}{2}\right)!\right|$

via 49.32
should be β^{-n} . See 55.20-23



since the function is symmetric about $\mu = \mu_{ob}$, the integration is trivial! - it's as if all mass were concentrated at μ_{ob} ! Not so trivial! [See 58.01-08 for Exact Argument]



The first moment is related to μ_{ob} .
Zeroth moment by factor μ_{ob} .

But we still have to find the zeroth moment.

Just set u to ϕ in integral $\int_{-\infty}^{+\infty} \beta^{1-n} d\mu$ ($\frac{2\beta^2}{n} \rightarrow (\sigma_{ob}^2 + \mu^2)$)

$\int_{-\infty}^{+\infty} \left(\frac{1}{\sigma_{ob}^2 + \mu^2}\right)^m d\mu = \frac{\sigma_{ob}^2}{(\sigma_{ob}^2)^m} \int_{-\infty}^{+\infty} \left(\frac{1}{1 + \frac{\mu^2}{\sigma_{ob}^2}}\right)^m \frac{d\mu}{\sigma_{ob}}$

probly unknown S.

20

23

62R
P 245
3.252.2
Also 1294
3.249.1

26

28

Note $\int_{-\infty}^{+\infty} \frac{dx}{(x^2 + c^2)^n}$ $c > 0$

$= \frac{(2n-3)!! \pi}{(2n-2)!!} \frac{1}{c^{2n-1}}$ $\approx \frac{\pi}{c^{2n-1}} \frac{1}{\sqrt{\pi(n-1)}}$

Approxly: $\exp\left(\frac{1}{2} + \frac{1}{2} + \frac{1}{2} + \dots + \frac{1}{2}\right) = \exp\left(\frac{1}{2} \ln j + \delta\right) = k \cdot j^{\frac{1}{2}}$

$\frac{1}{2} \cdot \frac{3}{4} \cdot \frac{5}{6} \cdot \frac{7}{8} \dots \frac{2j-1}{2j} \cdot \sqrt{\pi j} \rightarrow 1$

samey note bottom of p 190 of Jolly

It could be obtained from Stirling's, but not easily.

It may be that $\frac{1}{2} \pm \frac{1}{2} \cdot \frac{3}{2} \cdot \frac{5}{2} \dots$ is related to $(2n-3)!!$

$\int_{-\infty}^{+\infty} \frac{dx}{(x^2 + c^2)^m} \approx \frac{2\sqrt{\pi}}{\Gamma(m-1)} \cdot \frac{1}{c^{2m-1}}$ for $m > 1$

for $m=1$ probably wrong

NO!

2	4	6	8	10
1	3	5	7	9
3	4	6	8	10
1	3	5	7	9
1	3	5	7	9
1	3	5	7	9

I don't understand the σ_{ob} factor $n=20$
actually β .

$-m = -n+1$; $m = n-1$

$\int_{-\infty}^{+\infty} \frac{d\mu}{(\mu^2 + c^2)^{n-1}} \approx \frac{2\sqrt{\pi}}{\Gamma(n-2)} \frac{1}{c^{2n-2-1}} = \frac{2\sqrt{\pi}}{\Gamma(n-1)} \frac{1}{c^{2n-3}}$

Telix
C/Simul
K P 41

No! I expect it to be of dimension 1. (See dimensionless) No! $\int \frac{d\mu}{\mu} \rightarrow \ln \mu$

Dimension $\frac{1}{\sigma}$, so when I multiply by μ , I get $\dim \frac{\mu}{\sigma}$

Soln. 55.07

C/B/T/Y, Bat

Look at 53.38 w/o. $\frac{1}{\sigma}$ factor; Aproly dimensional form:

$\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \frac{1}{\sigma^n} \exp\left(-n \frac{(\sigma_{ob}^2 + (\mu - \mu_{ob})^2)}{2\sigma^2}\right) d\sigma d\mu$

T. integrand is in form of product:

$\prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(\sigma_i - \mu)^2}{2\sigma^2}}$

38

3.20.97 TM

Eq. 54.38 R would seem to be dimensionless $\frac{1}{\sigma^n}$: when integrated w. or direction; $\frac{1}{\sigma^{n-1}}$

when integrated w. $d\theta$ & $d\mu$, dimension $\frac{A}{\sigma^{n-1}}$ or $\frac{1}{\sigma^{n-2}}$, (since θ & ϕ have same dimension)

Consider a simpler problem! A variable has zero mean and unit σ^2 & is gaussian.

$\{S_i\}$ are observations; σ is 2 poss. s.d.

Prob of S_i if σ is true s.d. is: **woops!** prob of observation betw. S_i & $S_i + \Delta$ is $\int_{S_i}^{S_i+\Delta} \dots$

So, we have this factor Δ^n whenever we make n observations! That explains the dimensionality!

The $\frac{1}{\sigma^n}$ factor gives probability density in an n dim space (of n observations),

O.k. so lets get back to problem of 54.02 & 53.38

$$\int_{-\infty}^{+\infty} \beta^{1-n} d\mu = \beta = \left(\frac{n}{2} (\sigma_{00}^2 + \mu^2) \right)^{\frac{1}{2}} \quad \text{--- (i.s. } \mu_{00} = 0)$$

$$\left(\frac{n}{2} \right)^{\frac{1-n}{2}} \int_{-\infty}^{+\infty} \frac{d\mu}{(\sigma_{00}^2 + \mu^2)^{\frac{n-1}{2}}} = \left(\frac{n}{2} \right)^{\frac{1-n}{2}} \cdot \frac{2\sqrt{\pi}}{\sqrt{\frac{n-2}{2}}} \cdot \frac{1}{\sigma_{00}^{n-2}}$$

$n = 54.28 \quad m = \frac{n-1}{2}$
 $m+1 = \frac{n-2}{2}$
 $2m-1 = n-2$

I'm not sure about dimensions again: may be off by factor of σ^n .

In 54.02 I. dim after integration by $\frac{d\theta}{\sigma}$ was $\frac{\beta}{\sigma^n}$: or σ^{-n+1}
I think in I forgot the $\frac{1}{\sigma}$ factor: I. S should be of dim $\frac{1}{\sigma^{n+1}}$, $d\theta$ or $\frac{1}{\sigma^n}$
The ~~factor~~ ^{interand of (w.o.m)} 54.02 is ~~of~~ σ^{-n+1} so I ^{did} forget the $\frac{1}{\sigma}$ factor.

the β factorial part is correct, but only if exponent of β is wrong. should be $\beta^{-(n+1)+1} = \beta^{-n}$

So the integrand is $\beta^{-n} = (\beta^2)^{-\frac{n}{2}}$

in 54.28 $n = \frac{m}{2}$
 $m-1 = \frac{n-2}{2}$
 $2m-1 = n-1$

So $\int_{-\infty}^{+\infty} \beta^{-n} d\theta = \int_{-\infty}^{+\infty} (\beta^2)^{-\frac{n}{2}} d\mu$
 $= \left(\frac{n}{2} \right)^{\frac{n}{2}} \cdot \frac{2\sqrt{\pi}}{\sqrt{\frac{n-2}{2}}} \cdot \frac{1}{\sigma_{00}^{n-1}}$ \leftarrow Multiply $\frac{1}{2(\sqrt{\pi})^n} \cdot \left(\frac{n-2}{2} \right)!$ to get 54.02

for the final result, we multiply by Δ^n . No other expected values of $\frac{1}{\sigma^n}$ (!)

Clearly there is something wrong! The dimension is $\frac{A \cdot \sigma}{\sigma^n}$:
with integrand of dim $\left(\frac{A}{\sigma} \right) \frac{1}{\sigma^n}$ we integrate in $d\mu$ & $d\theta$ directions to get $\frac{A \cdot \sigma}{\sigma^n} \rightarrow \frac{A}{\sigma^{n-1}}$

But: we need a result w. dim zero: it's $\frac{A}{\sigma}$. Well; multiply by Δ^n to get the prob of being in the Δ^n cube assoc. w. n observations, $\{S_i\}$.

Maybe for some small Δ size, the integrand of 54.38 gives the relative distribution of σ, μ in the σ, μ plane. Its values do not sum to 1. The values are $\propto \Delta^n$, Δ being the cube side.

3.20.97 TM.

So maybe we have to normalize by integrating μ via 53.38 w.p. $\frac{\mu}{\sigma}$ factor
 This new integral relation is only slightly different from previous one:

T. integral corresponding to 54.02 (via 49.32) is 53.38 (w.p. $\frac{\mu}{\sigma}$);

.04

$$\int_{-\infty}^{\infty} dx \beta^{-n+1} \cdot \left(\frac{n-3}{2}\right)! \cdot \frac{1}{2} \cdot \left(\frac{1}{2\pi}\right)^{\frac{n}{2}}$$

evaluated at 55.10719

$$\left(\frac{n}{2}\right)^{\frac{1-n}{2}} \cdot \frac{2\sqrt{\pi}}{\sqrt{\frac{n-3}{2}}} \cdot \frac{1}{\sigma_0^{n-2}}$$

I want 55.30 divided by 56.04

.12

do it factor by factor

$$\frac{\left(\frac{n}{2}\right)^{-\frac{n}{2}}}{\left(\frac{n}{2}\right)^{\frac{1-n}{2}}} \cdot \frac{\sqrt{\frac{n-3}{2}}}{\sqrt{\frac{n-2}{2}}} \cdot \frac{\sigma_0^{n-2}}{\sigma_0^{n-1}} \cdot \frac{2(\sqrt{2\pi})^n \left(\frac{n-2}{2}\right)!}{2(\sqrt{2\pi})^n \left(\frac{n-3}{2}\right)!}$$

$$\left(\frac{n}{2}\right)^{\frac{1}{2}} \cdot \sqrt{\frac{n-3}{n-2}} \cdot \frac{1}{\sigma_0} \cdot \sqrt{\frac{n-1}{2}} = \frac{1}{\sigma_0} \sqrt{\frac{n-3}{n-2}} \cdot \sqrt{\frac{n-1}{2}}$$

$$= \frac{1}{\sigma_0} \cdot \sqrt{\frac{(n-3)(n-1)}{(n-2)n}}$$

which $\rightarrow \frac{1}{\sigma_0}$ as $n \rightarrow \infty$

So the expected value of $\frac{\mu}{\sigma}$ is $\frac{\mu_0}{\sigma_0} \cdot \sqrt{\frac{(n-3)(n-1)}{(n-2)n}}$ which is always < 1 ;
 $\approx \frac{\mu_0}{\sigma_0} \cdot \left(1 - \frac{1.25}{n}\right)$ for $n \gg 6$

Try it out for $n = 1, 2, 3, \dots, 20$

.20

TM 7-56. Box

See 56.6 for output
 $n = 1/40$

n	$\frac{\mu}{\sigma}$	approx $1 - 1.25/n$
4	.559	.69
6	.75	.79
10	.86	.875
15	.911	.916
20	.931	.934
25	.94	
30	.957	
35	.963	
40	.968	

is very very close: see 56.5

for large n:

$$\sqrt{1 - \frac{3}{n} - \frac{3}{2n} + \frac{1}{n}}$$

$$1 - \frac{3}{2n} = 1 - \frac{1.25}{n}$$

$$\approx 1 - \frac{1.25}{n}$$

$1 - \frac{1.25}{n}$

is much closer
 since for $n=5$ it's off by 8.0064%
 6.4%
 for $n=10$ it's off by 0.85% ; for $n=6$, it's exact.
 for $n=100$ it's off by .0025% ; for $n=20$, it's off by .013%

T. for μ results may be much modified by $\frac{\mu}{\sigma}$ spread. T. for σ assumes uniform.

spread for σ & μ .

That + $\frac{\mu}{\sigma}$ \in value into μ is not exactly $\frac{\mu_0}{\sigma_0}$ means it may not be "unbiased",

— so possibly unbiased divide it by $\frac{\mu}{\sigma}$ to "unbias" it ($>$).

n	$\sqrt{\frac{(n-2)(n-\frac{1}{2})}{(n-2)n}}$	$1 - 1.25/n$
4	.559017	.6875
5	.68313	.75
6	.75	.7916667
7	.792825	.8214285
8	.8228507	.84375
9	.8451542	.8611111
10	.8624094	.875
11	.8761716	.8863636
12	.887412	.8958333
13	.8967696	.9038461
14	.9046836	.9107143
15	.9114654	.9166667
16	.9173428	.921875
17	.9224859	.9264706
18	.9270248	.9305556
19	.9310602	.9342105
20	.9346716	.9375
21	.9379228	.9404762
22	.9408652	.9431818
23	.9435408	.9456522
24	.9459843	.9479167
25	.9482248	.95
26	.9502867	.9519231
27	.9521905	.9537037
28	.9539536	.9553571
29	.9555912	.9568965
30	.9571162	.9583333
31	.9585398	.9596774
32	.9598719	.9609375
33	.9611211	.9621212
34	.9622946	.9632353
35	.9633995	.9642857
36	.9644413	.9652778
37	.9654254	.9662162
38	.9663566	.9671053
39	.9672388	.9679487
40	.968076	.96875

99	.9872666	.9873737
100	.9873950	.9875

N.B. $1 - \frac{1.25}{n-1}$ is a much better approx. —

Its error at $n=6$ is error = .15% at $n=10$
 " = .002% at $n=100$.

at $n=1000$ error = $2.18 \times 10^{-5}\%$
 $= x(1 + 2.18 \times 10^{-7})$

Looks like error is like $\frac{1}{n^2}$.

A Review of 46.20 - 56.40:

T. problem: I have a lot of data strings & a lot of strategies for predicting them. I try the strategies on form strings & % of sub-sets of strings &, for various sizes (string lengths). For each, I get

1.07 $\mu = \frac{1}{n} \sum_{i=1}^n S_i$ and $\sigma_{ob}^2 = \frac{1}{n} \sum_{i=1}^n (S_i - \mu)^2$.

Here S_i is the "yield" of i trials. I'm interested in finding stats that have a large $\frac{\mu}{\sigma}$ per bet. This is $\frac{\mu}{\sigma} \cdot \sqrt{\text{no. of bets/yr.}}$ or

Max $\frac{\mu}{\sigma} \Big|_{\text{time } T} = \frac{1}{\sqrt{T}}$ (see 46.20-23 for discn).

SN This (46.20-56.40) investigation is a bit different from prodn. Its involved with "yield" which is perhaps more generally useful.

Consider a particular data sequ. $[S_i]$ of yields. If the generator was a Gauss d.f. of mean μ & s.d. σ , the probab of obtaining n data

in a n -cube $(S_i, S_i + \Delta)^n$ (so to speak) is

1.20 $\prod_{i=1}^n \left(\frac{1}{\sqrt{2\pi}\sigma} \cdot e^{-\frac{(S_i - \mu)^2}{2\sigma^2}} \right) = \frac{1}{(2\pi)^{\frac{n}{2}}} \frac{\Delta^n}{\sigma^n} e^{-\frac{n(\sigma_{ob}^2 + (\mu_{ob} - \mu)^2)}{2\sigma^2}}$ [see defn for σ_{ob}^2, μ_{ob}]

This considers the a priori of μ, σ to be uniform. for a post non-uniform a priori: see 48.25.

1.22 If we let $\beta^2 = \frac{n}{2} (\sigma_{ob}^2 + (\mu_{ob} - \mu)^2)$ we get

1.23 $(2\pi)^{-\frac{n}{2}} \frac{\Delta^n}{\sigma^n} e^{-\frac{\beta^2}{\sigma^2}} = f(\mu, \sigma) \cdot \Delta^n$

Given that n data occurred in the n -cube Δ^n , the Bayesian distribution on μ, σ is given by $f(\mu, \sigma)$

This is not a normalized d.f.; only gives relative densities in the μ, σ plane.

To get the expected value of $\frac{\mu}{\sigma}$:

1.29 $\frac{\mu}{\sigma} \Big|_{\text{per bet}} = \frac{\int_{-\infty}^{\infty} \int_0^{\infty} \frac{\mu}{\sigma} f(\mu, \sigma) d\mu d\sigma}{\int_{-\infty}^{\infty} \int_0^{\infty} f(\mu, \sigma) d\mu d\sigma}$ ← integral in denom. is indep of μ , so set it to 1 in denom.

Since the $(2\pi)^{-\frac{n}{2}}$ factor occurs in both numerator & denominator, we will not consider it henceforth.

33 Consider denominator: $\int_0^{\infty} d\sigma \int_{-\infty}^{\infty} f(\mu, \sigma) d\mu = \int_0^{\infty} \frac{d\sigma}{\sigma^n} e^{-\frac{\beta^2}{\sigma^2}} = \frac{\beta^{-n+1}}{2} \left(\frac{n-3}{2}\right)!$ see 49.29-32 for this last derivation.

34 next, consider $\int_{-\infty}^{\infty} d\mu \beta^{-(n-1)} = \int_{-\infty}^{\infty} d\mu (\beta^2)^{\frac{1-n}{2}}$

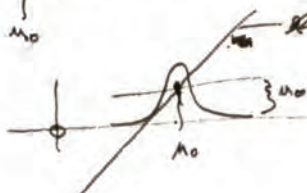
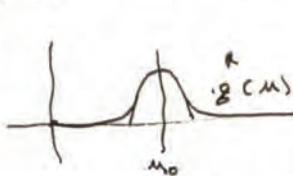
35 $\left(\frac{n}{2}\right)^{\frac{1-n}{2}} \int_{-\infty}^{\infty} d\mu (\sigma_{ob}^2 + \mu^2)^{\frac{1-n}{2}} \approx \frac{2\sqrt{\pi}}{\sqrt{\frac{n-3}{2}}} \cdot \frac{1}{\sigma_{ob}^{n-2}} \cdot \left(\frac{1}{2} \left(\frac{n-3}{2}\right)!\right)$

2 most cancells cancel 49.33R

36 $\left. \begin{aligned} m &= \frac{n-1}{2} \\ m-1 &= \frac{n-3}{2} \\ 2m-1 &= n-2 \end{aligned} \right\}$

→ 58.09

01 **SN** Root ~~method~~ arguments of $57.20 \rightarrow .20$ on integrating $(\beta^2)^{\frac{n-1}{2}}$!



$g(u)$ is symmetric about u_0 . I want $\int g(u) du$.

It is $\int_{-\infty}^{\infty} g(u) du$

$$\int_{-\infty}^{\infty} g(u) du$$

the reason is!

Say $l(u)$ is a str. line thru $(u_0, g(u_0))$ with slope 1. Then the \int of $l(u) \cdot g(u)$ is indep of u_0 . So let slope = 1. This is true.

08
09 : 57.40! Next, look at numerator of 57.29:

first do $\int_0^{\infty} du \frac{1}{G} F(u, G) = \int_0^{\infty} \frac{du}{G^{n+1}} e^{-\frac{\beta^2}{G^2}} \beta^{-n+1} \left(\frac{n-1}{2}\right)! = \beta^{-n} \cdot \frac{1}{2} \left(\frac{n-1}{2}\right)!$

$$\int_{-\infty}^{\infty} du \cdot u \beta^{-n} = M_{ob} \int_{-\infty}^{\infty} \beta^{-n} du = M_{ob} \int (\beta^2)^{-\frac{n}{2}} du = \int_{-\infty}^{\infty} du \cdot u \beta^{-n}$$

with $M_{ob} = 100$

See 58.01-08 for why this is true.

$$= M_{ob} \frac{1}{2} \left(\frac{n-1}{2}\right)! \int du (G_{ob} \cdot u^2)^{-\frac{n}{2}} = \frac{2\sqrt{n}}{G_{ob}^{n-1}} \cdot \frac{1}{2} \left(\frac{n-1}{2}\right)!$$

19 $(10) = \text{entire numerator} = \left(\frac{n}{2}\right)^{-\frac{n}{2}} \frac{2\sqrt{n}}{G_{ob}^{n-1}} \cdot \frac{1}{2} \left(\frac{n-1}{2}\right)! \cdot M_{ob}$

58 19
57 38 $\approx \left(\text{see 56.12 for actual calculus} \right) \approx \frac{M_{ob}}{G_{ob}} \left(\frac{n}{2}\right)^{\frac{n}{2}} \sqrt{\frac{n-3}{n-2}} \cdot \frac{\left(\frac{n-1}{2}\right)!}{\left(\frac{n-3}{2}\right)!}$

25 $\approx \frac{M_{ob}}{G_{ob}} \cdot \left(\frac{(n-3)(n-\frac{3}{2})}{(n-2) \cdot n}\right)^{\frac{n}{2}}$ This is the expected value of $\frac{1}{G}$ for n observations. See 61.25 for comments on validity.

26 $\approx \frac{M_{ob}}{G_{ob}} \left(1 - \frac{1.35}{n}\right)$ for $n > 10$: see (56.20) for value as a function of n for (23) is for approx of (26).
 $\approx \frac{M_{ob}}{G_{ob}} \left(1 - \frac{1.25}{n-1}\right)$ is a much better approx. here.

N.B. The approx. of $\frac{(X+\frac{1}{2})!}{X!}$ is very good even for X as small as 2 & improves much as $X \uparrow$.

The approx of $\int \frac{du}{\beta^n}$ of 57.28 is also v.g. even for m as small as 2, it's only off by ~ 6%. ; 2% error for $m > 6$.

The S 's diverge for $n=1$ or 2; for $n=3$ + formula for $\int_{-\infty}^{\infty} \frac{du}{(G_{ob}^2 + u^2)^m} = \frac{2m-3!!}{2^{m-2} (m-2)!} \frac{\pi}{G_{ob}^{2m-1}}$ is incorrect: $m=3$ gives $m=1$!

The correct form when $n=3$ or eqn. 25 is $\int_{-\infty}^{\infty} \frac{du}{(G_{ob}^2 + u^2)} = \frac{\pi}{G_{ob}}$

$$\frac{M_{ob}}{G_{ob}} \cdot \left(\frac{n-3}{n}\right)^{\frac{n}{2}} = \frac{M_{ob}}{G_{ob}} \cdot \sqrt{\frac{1}{2}}$$

for nice way to deal w. $n=3$! Also correction for 58.23

See 65.01

Another Q one might ask is: If the "true" μ is σ (or μ is σ), what are the expected values of μ_0 & σ_0 for n observations? What is the d.f. of these estimates?

Well, the expected value of μ_0 is μ . I think $\frac{\sigma^2}{n}$ is expected value of σ_0^2 is $\frac{\sigma^2(n-1)}{n}$

Fidelity Criterion.

T. result of 58.25 was obtained assuming uniform spread for σ : What would it look like if we assume uniform spread for σ^2 ?

$d(\sigma^2) = 2\sigma d\sigma$. It has the effect of using $n \rightarrow n-1$ in the data of $f(\mu, \sigma)$ of 57.23

So the integrands in both forms $d\mu$ of 57.23 are mult by 20.

The result should be easy to compute, using the (almost) exactly same eqns. as 57.01-58.40:

~~But I'm not sure it is~~ But my impression is that this is not the main present problem.

At first glance it seems to do $n \rightarrow n-1$ in 58.25

1.13
$$\sigma_0 \approx \frac{\mu_0}{\sigma_0} \left(\frac{(n-1)(n-\frac{1}{2})}{(n-2)(n-1)} \right)^{\frac{1}{2}}$$
 But this should be checked, more carefully, if it is to be a check.

for "large" n it is $\frac{\mu_0}{\sigma_0} \left(1 - \frac{1.25}{n} \right)$ - same as 58.25, but n must be larger than 5. 2.5

becomes a good approx. It may be shown an impl. distance say for $n=5$. -1-1.5

Look at table of 56.20: r. formula of .13 amounts to $n \rightarrow n-1$

T. distance betw. 58.25 & 59.13 is best studied by looking at $1 - \frac{1.25}{n}$ v.s. $1 - \frac{1.25}{n-1}$:

No dramatic distance! See listing of 56.2

If formula (.13) is wrong, it is most likely to be of form of changing certain n values (in both num & denom) by 1 - producing little change ~~approx.~~ is no change approx. $1 - \frac{1.25}{n}$

Actually, what I probably got in 58.25 was the result of finding the expected value of $\frac{1}{\sigma}$.

Not quite! In deriving wrt. μ did produce another $\sqrt{\frac{(n-3)}{(n-2)}}$ factor:

I'm not sure this is meaningful: to get $E \frac{1}{\sigma}$, one has to integrate wrt. μ , in any way

So I guess the result of finding $E \frac{1}{\sigma}$ would be identical to 58.25. Then with μ by μ_0

OK. Returning to log neutrons: I want to find a stat. w. by $\left(\frac{\mu}{\sigma}\right)^2 \frac{1}{T}$: this is exponentially yield per unit time. So perhaps I should have computed $E \frac{\mu^2}{\sigma^2}$.

Most likely, it will look like of form $\left(\frac{\mu_0}{\sigma_0}\right)^2 \cdot \left(\frac{(n-2)(n-b)}{(n-c)(n-d)}\right)^{\frac{1}{2}}$

Now: say all systems considered have 1 bit/day: so I want stat. w. max

$$\left(\frac{\mu_0}{\sigma_0}\right)^2 \cdot \left(\frac{(n-2)(n-b)}{(n-c)(n-d)}\right)^{\frac{1}{2}} \cdot \frac{1}{n}$$
 limit μ_0 & σ_0 will be around 4 times σ of true, resp.

In general, $\sigma_0^2 \cdot n$ will be about constant, so for various values of n , this expression will

not change much: this part will be a bit w. n .

So from this discussion, there would seem to be little motivation to use larger n values

in this case! - not at all reasonable!

3.21.97 T.M.

Perhaps what I want is a D.F. for $(\frac{1}{n} \cdot \frac{M^2}{\sigma^2})$ induced by t. data.

To get this, integrate $f(\mu, \sigma)$ (57.23) w.r.t. σ (me $\frac{M}{\sigma} = k = \text{constant}$)

for a first approx: Get σ_{ob}^2 using some reasonable n ; then we have M centered about M_{ob} w. var $\sim \frac{\sigma_{ob}^2}{n}$. - or M_{ob} w. var $\sim \sigma_{ob}^2$.

So $\frac{M^2}{\sigma^2}$ has \approx mean $\frac{M_{ob}^2}{\sigma_{ob}^2}$ w. var $\frac{4}{(\sigma_{ob}^2)^2} \cdot \frac{1}{n} \approx \frac{1}{n}$.

\therefore mean $(\frac{M_{ob}}{\sigma_{ob}})^2 \approx \frac{1}{n}$

I suspect that $(\frac{M_{ob}}{\sigma_{ob}})^2 \div (\frac{1}{n}) = n (\frac{M_{ob}}{\sigma_{ob}})^2$ has to be > 1 for one can get much expected yield out of it.

One would $E M^2$ is always > 0 . The actual yield is not M^2 , but $\frac{1}{n}$ times what one thinks it is. For small M , this product can be > 0 or < 0 ; On t. average, for small M , it's close to 0. So the yield may be $\propto M_{ob} \cdot M$. I think n has to be large enough so that one is fairly sure of t .

Size of $M(!)$ - I think this is $\approx n (\frac{M_{ob}}{\sigma_{ob}})^2 > 1$ i.e. n must be large enough so that $n (\frac{M_{ob}}{\sigma_{ob}})^2 \approx > 1$.

Which suggests that I can't learn much from small n ! In general, it would seem

on t. that I'd get $(\frac{M_{ob}}{\sigma_{ob}})^2 \cdot n > 1$ for $n < 2$ yrs!

One possy. is that w. $n < 2$ yrs, I get $\frac{M_{ob}}{\sigma_{ob}}$ on of some same sign. for

a very large no. of stocks.

Woops! t. Idea of .15-.95 is wrong!

even if $\frac{M_{ob}}{\sigma_{ob}}$ is very small,

t. expected value of $M_{ob} \cdot M$ is still M_{ob}^2 !

Assuming mean is M_{ob} & betting one way, very return

has neg. yield, - but very often, yield is $>> M_{ob}^2$ (i.e. $M > M_{ob}$). - on t. average,

yield is $\approx M_{ob}^2$. For a low $\frac{M}{\sigma}$, bet fully bet is $\propto \frac{M}{\sigma}$; yield is $1 + \frac{M^2}{\sigma^2}$

or $\exp(\frac{M^2}{\sigma^2})$ per bet. If M has 2 poss

Actually if M_{ob} is mean of broad Gauss d.f., one bets an amount $\propto M \cdot e^{-\frac{(M-M_{ob})^2}{2\sigma^2}}$ (assuming $\sigma_0 = \sigma$ = actual sigma).

Def: $M_0 \equiv M_{observed} \equiv M_{ob}$; $\sigma_0 \equiv \sigma_{observed} \equiv \sigma_{ob}$.

total bet = $\int_{-\infty}^{\infty} M e^{-\frac{(M-M_0)^2}{2\sigma_0^2}} dM$



$\propto M_0$.

So it would seem that value of t-stret was $\propto (\frac{M_0}{\sigma_0})^2$ (part bet = "padding")

If $(\frac{M_0}{\sigma_0})^2$ was obtained w. large n , t. fact that M_0 is known to hyper precision, does not seem to \uparrow yield!

01 Seems Unreasonable! Say x_i data is random. We will have a ^{broad} distrib. of M 's for t . Various strategies, ~~but we will think~~ Over "expected" ~~yield~~ yield will be $\frac{1}{n}$ maybe. [i.e. we will think M^2 is larger when n is small.]

S: in case $(.01)$ how would we realize that there were no winning stats? —

Well, we would notice that t . no. of apparently winning stats is consistent w/ hypothesis of randomness of data, & that this hypothesis continues to fit for larger values of n .

This way of looking at things is easy to justify if t . "Stat" is a ~~pen~~ Pen — in which case t . random hypothesis is ~~not~~ of by a priori, & gives t . same code for ~~the~~ (or no longer) than other models give.

Well, actually, if we multd. 57.20 (≈ 57.23) by t . a priori, we would have a rite answer.

t . fact is, a uniform prior is unlucky & "unreasonable", & it's not surprising that it gives "Unreasonable" results.

On t . other hand, w/ a reasonable prior to multd. by, 57.20 is correct, & unarguably so. Bayes idea is a "Theorem" ~~not~~ not a "suggestion".

So, a reasonable prior is the only way out of the "Unreasonableness" of $(.01)$

We can start out by obtaining a prior via large n for maybe 10 trials: Have n large enough so that $n \left(\frac{M_0}{60}\right)^2 > 1$ — say $n = 2$ or 4.

After one has begun to get useful results, t . a priori will change! it will be a funct.

0.24 OP 6: params used to define t . strategy.

2.5 SN Formula 58.25 is clearly wrong for $n=2$ or 3; for $n=1$ it gives imaginary values \therefore wrong. However, for $n=1, 2, 3$, t . Bayesian treatment gives reasonable results. The trouble is that t .

approx: $\frac{(x+\frac{1}{2})!}{x!} = \sqrt{x+\frac{3}{4}}$ works o.k. for $x=2$ or even $n=1$

$n=1$ $\frac{(1+\frac{1}{2})!}{1!} = \frac{3}{4}\sqrt{\pi} \approx 1.32934$ (5% error)
 $\sqrt{1+\frac{3}{4}} = \sqrt{\frac{7}{4}} = 1.322$

$n=2$ $\frac{(2+\frac{1}{2})!}{2!} = \frac{5}{2} \cdot \frac{3}{2}\sqrt{\pi} = 6.646$
 $\div 2! = 3.32335$

$\sqrt{2.75} =$
 $\left(\frac{(0+\frac{1}{2})!}{0!}\right) = \frac{1}{2}\sqrt{\pi} = .886227$
 $\sqrt{.75} \approx .8660$ — 2% off

From 58.26: $\frac{(2n-3)!!}{(2n-2)!!} \sim \frac{1}{\sqrt{\pi(n-1)}}$ for $n=1$ $-1!!$ may be \emptyset or $(2-2)!!$ means 1

for $n=2$: $\frac{1}{2} = .5$; $\frac{1}{\sqrt{\pi}} = .564$; which is only 13% off.

for $n=1$: $\frac{-1!!}{1!} = 1$; $\frac{1}{\sqrt{\pi-1}} \approx 1.2$

66.01
66.20
Spec

$-\frac{1}{2}! = \sqrt{\pi}$
 $\left(\frac{3}{2}\right)! = \frac{3}{2} \cdot \frac{1}{2}! = \frac{3}{4}\sqrt{\pi}$
 $\frac{1}{2}! = \frac{\sqrt{\pi}}{2}$

I. S converges for $m=1$.

T. approx ≤ 4.28 is clearly wrong for $m=1$. — Tho + S would diverge for $m=1$.

I have note that $(-1)!! = 1$; and of course $(2n)!! = 2^n n!$

↑ I don't know where I got this!

$$\int_{-\infty}^{\infty} \frac{dx}{(x^2+a^2)^n} = \frac{(-1)!!}{0!!} \cdot \frac{\pi}{2^n} = \frac{\pi}{2^n} \quad \text{So the approx is wrong & diverges for } n=1.$$

But for $n=2 \text{ \& } 3$ it's o.k. — so 58.25 must be wrong!

In 58.25 the $\left(\frac{n-3}{n}\right)^{\frac{1}{2}}$ factor(s) are o.k. for $n \geq 1$. $\frac{n-3}{n-2}$ comes from $\int_{-\infty}^{\infty} \frac{dx}{(1+x^2)^{n-1}}$ ($n=1,2,3, \dots$)

for $n=1$: in 57.34. $\int du (\beta^2)^{\frac{n-1}{2}}$ diverges for $n=0,1,2$; but converges for $n \geq 3$.

The main problem is the denom. of 57.29 ("Herman factor").

$$\iint \frac{dudv}{\sigma^n} e^{-\frac{u^2+v^2}{2\sigma^2}} = k \cdot \frac{1}{\beta^{n-1}} \quad \beta = \left(\sqrt{\sigma_0^2 + u_0^2} \cdot \sqrt{\frac{\pi}{2}}\right); \beta^2 = \frac{\pi}{2} (\sigma_0^2 + u_0^2)$$

let $u_0 = 0$.

$$\int du \frac{1}{\beta^{n-1}} \text{ diverges for } n=1 \text{ or } n=2, \text{ but converges for } n \geq 3.$$

(SN) The express for the ratio of $\int \beta^{n-1}$ for ~~numerator~~ / denom. is obtained in an approx. way! But

Actually, if I write out the exact $\frac{2!!}{0!!}$ factors for num & denom, they would mainly cancel!

out if I make a very simple approx ~~numerator~~ process!

$$\text{say } \frac{\text{Den}}{\text{Num}} = \frac{2n-3!!}{2n-2!!} \left(\approx \frac{\sqrt{\pi}}{\sqrt{n-1}} \right) \quad n = \frac{n}{2} - 1$$

$$\approx \text{same} \quad n \approx \frac{n}{2}$$

$$\text{Num} = \frac{2n-5!!}{2n-4!!}$$

$$\text{numerator} \int \frac{1}{(\beta^2)^{\frac{n-1}{2}}} = \frac{\cancel{2n-3!!}}{\cancel{2n-2!!}} = \frac{n-1-3!!}{n-1-2!!} = \frac{n-4!!}{n-3!!}$$

$$\text{denom} \int \frac{1}{(\beta^2)^{\frac{n}{2}}} = \frac{n-3!!}{n-2!!}$$

So nothing cancels!

If the numerator were the expected value of $\frac{1}{\sigma^2}$, they would cancel a lot. I don't know what happens — by \int a factor like $\frac{n-3}{n-2}$; but by doing expected value of $\frac{1}{\sigma^2}$, I say for $\frac{1}{2}$ way, so $\sim \sqrt{\frac{n-3}{n-2}}$.

Looking at 57.33 & 34: for $n=1,2$, $\int \frac{1}{\beta^{n-1}}$ diverges;

for $n=3$, $\int \frac{1}{\beta^{n-1}}$ converges, but the approx. I'm using is invalid for $n-1=2$ (it's ok if $n-1=3,4,5, \dots$)

for $n=3$ make special calcn.

$$\text{denom} \int_{-\infty}^{\infty} \frac{dx}{(c^2+x^2)^{\frac{n}{2}-\frac{1}{2}}} \quad ; \quad \int_{-\infty}^{\infty} \frac{dy}{(c^2+y^2)^{\frac{n}{2}}}$$

$$\text{for } n=3 \quad \int_{-\infty}^{\infty} \frac{dx}{(c^2+x^2)} = \frac{\pi}{c} \quad \left(\text{for } n=3 \text{ numerator } \int_{-\infty}^{\infty} \frac{dy}{(c^2+y^2)^{\frac{3}{2}}} = \frac{\sqrt{\pi}}{\sqrt{2}} \cdot \frac{1}{c^2} \right)$$

$$= \frac{3-3!!}{3-2!!} \cdot \pi \cdot \frac{1}{c^2} = \frac{1}{1} \cdot \frac{\pi}{c^2}$$

So, for $n=3$ the $\left(\frac{n-3}{n-2}\right)$ factor is replaced by 1 .

Maybe bad error on 54.28:

$$n=3 : \frac{2n-3}{2n-2} = \frac{6-3}{6-2} = \frac{3!!}{4!!} = \frac{3}{8}$$

was

Actually, the formula I should be using is

05 Turnout $\frac{n-1!!}{n!!} = \frac{1}{\sqrt{\pi}} \frac{(\frac{n-1}{2})!}{\frac{n}{2}!}$ exactly! (probably for n integer or half integers).

so $\hookrightarrow \approx \frac{1}{\sqrt{\pi}} \frac{1}{\sqrt{\frac{n}{2} + \frac{1}{2}}}$

13 $\int_{-\infty}^{\infty} \frac{dx}{(x^2+c^2)^m} = \frac{\pi}{c^{2m}} \frac{2m-3!!}{2m-2!!} \approx \frac{\pi}{c^{2m-2}} \frac{1}{\sqrt{\pi}} \frac{1}{\sqrt{2m-2+\frac{1}{2}}} \approx \frac{\sqrt{\pi}}{c^{2m-1}} \frac{1}{\sqrt{m-\frac{3}{4}}}$

Contrast w. 54.28 in which we have $m-1$ instead of $m-\frac{3}{4}$.

How good is the approx when $m=1$? $\frac{2m-3!!}{2m-2!!} = 1 + \text{rand.}$

17 $\frac{-1!!}{0!!} \approx \frac{1}{\sqrt{\pi}} \frac{1}{\sqrt{\frac{1}{2}}} = \sqrt{\frac{2}{\pi}} = 1.128$ not great but tolerable. (13% error).

$n=2$: $\frac{1!!}{2!!} = \frac{1}{2}$ $\frac{1}{\sqrt{\pi}} \frac{1}{\sqrt{1+\frac{1}{2}}} = .5096$

20 So change 58.25 to $\frac{4_{ob}}{8_{ob}} \cdot \sqrt{\frac{(n-2.75)(n-1.5)}{(n-1.75)(n)}}$

This formula gives imaginary or ∞ for $n=0, 1, 2$:
for $n=3, 4$ gives a real value — but small

.375

I think I did that conversion wrong. 58.25 was originally

TM7-62.05

The correction is $\sqrt{\frac{\frac{n-3+\frac{1}{2}}{2} \cdot \frac{n-1.5}{2}}{\frac{n-2+\frac{1}{2}}{2} \cdot \frac{n}{2}}}$

$\rightarrow \sqrt{\frac{n-3+\frac{1}{2}}{n-2+\frac{1}{2}} \cdot \frac{n-1.5}{n}} = \sqrt{\frac{n-2.5}{n}}$ quite simple, if true!

29

$\sqrt{\frac{n-2.5}{n}} \approx 1 - \frac{2.5}{n}$ would $(1 - \frac{1.25}{n-1})$ be better?

$(1 - \frac{2.5}{n})^2 = 1 - \frac{5}{n} + \frac{1}{4} (\frac{2.5}{n})^2$

$(1 + \frac{1.25}{n-1})^2 = 1 + \frac{2.5}{n-1} + \frac{3.9}{16} \cdot \frac{1}{(n-1)^2}$

$\frac{2.5}{n-1} = \frac{2.5}{n} \left(\frac{n}{n-1}\right)^2 \frac{1}{(1-\frac{1}{n})^2} \approx 1 + \frac{1}{n} + \frac{1}{n^2} + \frac{1}{n^3}$

$= \frac{2.5}{n} \left(\frac{1}{n} + \frac{1}{n^2} + \frac{1}{n^3}\right) + \frac{2.5}{16} \left(\frac{1}{n-1}\right)^2$

Re formula: $\int_{-\infty}^{\infty} \frac{dx}{(x^2+c^2)^n} = \frac{2n-3!!}{2^{n-1} \cdot 2^{n-1} \cdot c^{2n-1}} \cdot \pi$ is true only if n is an integer

If it's a $\frac{1}{2}$ integer (i.e. not an integer), formula is wrong by a factor $\frac{\pi}{2}$ (!).

Heuristics: Now: $\int_0^{\frac{\pi}{2}} \sin^n x dx = \int_0^{\frac{\pi}{2}} \cos^n x dx$

$$= \frac{n-1!!}{n!!} \cdot \frac{\pi}{2} \text{ if } n \text{ is even}$$

$$= \frac{1}{2} \cdot \frac{\pi}{2} \text{ if } n \text{ is odd}$$

$$= \frac{1}{2} \sqrt{\pi} \frac{\Gamma(\frac{n}{2} + \frac{1}{2})}{\Gamma(\frac{n}{2} + 1)}$$

so, probably $\int_{-\infty}^{\infty} \frac{dy}{(x^2+c^2)^n} = \frac{\pi}{c^{2n-1}} \sqrt{\pi} \frac{\Gamma(\frac{n}{2} + \frac{1}{2})}{\Gamma(\frac{n}{2} + 1)}$

for even and odd n .

is probably any n .

say $\int_{-\infty}^{\infty} \frac{dy}{(x^2+c^2)^n} = g(n)$: $g(n+1) = \frac{2n-1}{2n} g(n)$

$$= \frac{1}{2} \left(\frac{2n-1}{2n}\right) g(n)$$

for integral n .

Most likely: My formula for $g(n) \approx \frac{\sqrt{\pi}}{\sqrt{n-0.75}}$ is probably good for both integer values of n .

in integral is $\frac{1}{2}$ integer values of n .

$g(\frac{1}{2})$

57.37 ff: for $n=3$ we want $\int dx (c^2+x^2)^{-3}$ which =

$\int_0^{\infty} \frac{dx}{(x^2+1)^3} = \frac{1}{2} \tan^{-1} x \Big|_0^{\infty} = \frac{\pi}{2}$ $\therefore \int_{-\infty}^{\infty} \frac{dy}{(1+x^2)^3} = \pi$ (See also 62.03)

$n=1$ compare w. 63.13 $\frac{\sqrt{\pi}}{\sqrt{1-0.75}} = \sqrt{\frac{\pi}{0.25}} = \sqrt{4\pi} = 2\sqrt{\pi} = 3.5499 = \pi \times 1.1284$ (See 63.17)

13% off. (high)

This is about as bad as it gets, Avr. not divisible by $\frac{\pi}{2}$.

$\int_{-\infty}^{\infty} \frac{dx}{(x^2+1)^{\frac{3}{2}}} = \frac{0!!}{1!!} \cdot \pi = \pi$ | $\int_{-\infty}^{\infty} \frac{dx}{(x^2+1)} = \pi$ also!

This is impossible! $(x^2+1)^{\frac{3}{2}}$ is ~~stump~~ always $>> (x^2+1)$

so maybe $= \pi \cdot \frac{2}{\pi} = 2$.

$\frac{\sqrt{\pi}}{\sqrt{1.5-0.75}} = \sqrt{\frac{\pi}{0.75}} = 2.0466$

only 2% off high

$n=1.5$
 $n=2$

$\int \frac{1}{(x^2+1)^2} = \frac{1!!}{2!!} \cdot \pi = \frac{\pi}{2}$

$\frac{\sqrt{\pi}}{\sqrt{2-0.75}} = \sqrt{\frac{\pi}{1.25}} = 1.5708$

1% off (high)

1.73

1.73 instead of 1.75

first 8.5% error on $n=1$ is 2.3% error on $n=1.5$

2 1.28 " " $n=2$

$\sqrt{\pi}$ v.s. $\frac{\pi}{2}$

$\frac{\sqrt{\pi}}{\pi} \cdot 2$

-1!! seems to be 1

on the other hand, it = ϕ from 63.05

10:

$\frac{-1!!}{0!!} = \frac{1}{\sqrt{\pi}} \frac{(\frac{1}{2})!}{0!}$

$-\frac{1}{2}! = \sqrt{\pi}$

so if $0!! = 1$

$-1!! = 1 > 1/2!$

$\frac{\sqrt{\pi}}{1.25} \cdot \frac{2}{\pi} = \sqrt{\frac{\pi \cdot 4}{1.25 \cdot \pi}}$

$= \sqrt{\frac{4}{1.25}} = \sqrt{\frac{16}{5}}$

$= 4 \sqrt{\frac{1}{5}}$

Feeling w. x to give better fit at $n=1$ & 1.5 seems pointless, since I will probably never use n that small.

ol: 58.40: The result of the review P57-58 is \approx correct, but formula at 58.25 is a bit off.

should be expected value of ~~$\frac{\pi}{2}$~~ $\frac{1}{5}$ is $\frac{16b}{60b} \sqrt{\frac{n-2.5}{n}}$

The work for this is ^{correction} 63.25 - 63.29:

The error in previous work resulted from the formula

$$\int_{-\infty}^{+\infty} \frac{dx}{(1+x^2)^m} = \frac{2^{m-3}!!}{2^{m-2}!!} \pi$$

being true only if m was an integer. If m was of form $m = \frac{1}{2}$, the formula had an extra $\frac{\pi}{2}$ factor!

for $\frac{1}{2}$ integral it's $\frac{2^{m-3}!!}{2^{m-2}!!} \cdot 2$

More generally $\int_{-\infty}^{+\infty} \frac{dy}{(1+y^2)^m} = \sqrt{\pi} \frac{(m-\frac{1}{2})!}{(m-1)!}$ is exact for all $m > \frac{1}{2}$ for $m \leq \frac{1}{2}$ the integral diverges.

This means I can use my formula good approx for $\frac{x+\frac{1}{2}}{x!} \approx \sqrt{x+\frac{1}{2}}$ (good for any x , in half ints or whatever).

So $\int_{-\infty}^{+\infty} \frac{dx}{(1+x^2)^m} = \frac{\sqrt{\pi}}{\sqrt{m-\frac{1}{2}}}$ This formula is $\left. \begin{array}{l} 13\% \text{ high for } m=1 \\ 2\% \text{ " " } m=1\frac{1}{2} \\ 1\% \text{ " " } m=2 \end{array} \right\} \text{ 69.20ff.}$
 (instead of $\frac{\sqrt{\pi}}{\sqrt{m-1}}$)

In 57.38 and 58.19 we get factors $\sqrt{\frac{n-3}{2}}$ & $\sqrt{\frac{n}{2}-1}$, resp:

the final formulae, 58.25 has their ratio; $\frac{\sqrt{\frac{n-3}{2}}}{\sqrt{\frac{n}{2}-1}}$; 65.18 changes it into

$$\sqrt{\frac{\frac{n}{2}-\frac{3}{2}+\frac{1}{4}}{\frac{n}{2}-1+\frac{1}{4}}} = \sqrt{\frac{n-3+\frac{1}{2}}{n-2+\frac{1}{2}}} = \sqrt{\frac{n-2\frac{1}{2}}{n-1\frac{1}{2}}}$$

cancel out and we get $\frac{16b}{60b} \sqrt{\frac{n-2.5}{n}}$

In 58.25, the two "n-1/2" factors $\frac{n-2.5}{n-1.5} \cdot \frac{n-1.5}{n} \rightarrow \sqrt{\frac{n-2.5}{n}}$

d: 61.24:

(SN) Actually, the GS: 02 term shouldn't work for $n=3$ any way; this is because while

ol $\int \frac{d\mu}{(\sigma^2 + \mu^2)^2}$ does converge, $\int \frac{\mu d\mu}{(\sigma^2 + \mu^2)^2}$ diverges: (The μ & σ^2 's cancel out)

what, we will want $E \frac{\mu^2}{\sigma^2}$ and well, no: the divergence problem occurs in both Num. & denom of 57.29 for the same values of n .

In the denom, divergence is for $n=0,1,2$:

In numerator, when we integrate w $\frac{\mu}{\sigma}$, we increase exponent \uparrow (0.01) search by $\frac{1}{2}$ so it converges better, then mult by μ so it converges worse. The σ^2 effects leave the fact of convergence invariant: so it converges for $n \geq 3$.

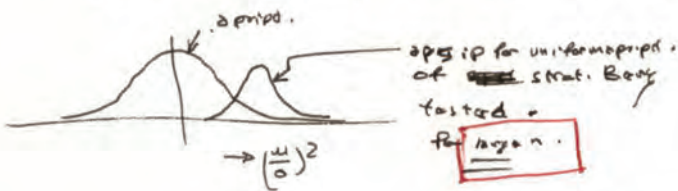
Similarly, when we want $E \frac{\mu^2}{\sigma^2}$; the effects of μ^2 & $\frac{1}{\sigma^2}$ on convergence, cancel out, so it converges for $n \geq 3$.

Note, here, one must integrate wrt to then $d\mu$. If one does it in the inverse order, I don't know what happens! Actually, it may be a much easier integration! Or, on the other hand, it begins to look about the same!

1: spec
2: 61.24

Say we decide on an μ prior for $(\frac{\mu}{\sigma})^2$: we could then find the $E(\frac{\mu}{\sigma})^2$ of a particular stat. If the "true" $(\frac{\mu}{\sigma})^2$ were large (say \gg the s. d. of μ).

μ prior for $(\frac{\mu}{\sigma})^2$ then, as $n \uparrow$, the $E(\frac{\mu}{\sigma})^2$ will \uparrow , because its d.f. gets narrower



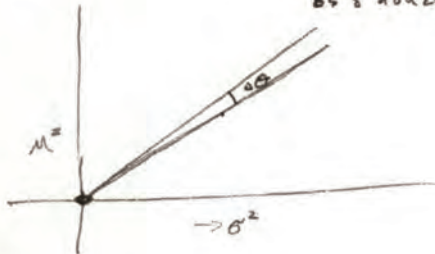
While $E(\frac{\mu}{\sigma})^2$ is of importance, we are also much interested in the d.f. of $(\frac{\mu}{\sigma})^2$ — its width of this d.f.

The reason is: If we are selecting the "best" of a bunch of stats, we need

to know the width to know how likely it is that our "best" is spurious.

One approach:

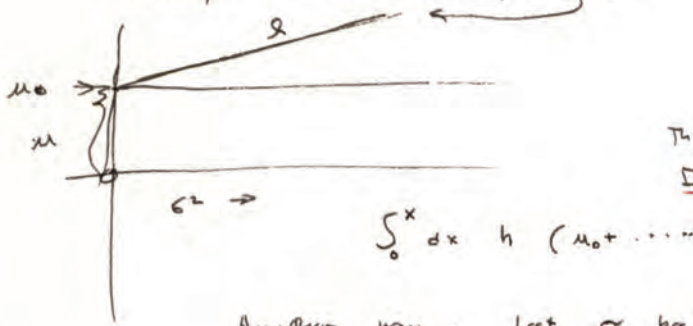
as a function of μ^2 & σ^2 plot in 2 dimensions, the area of density of μ^2, σ^2 , as a function of $[S_i]$. Then integrate in the narrow angle $\Delta \theta$.



Say $f(\mu^2, \sigma^2)$ is the prior. $k < \frac{\mu^2}{\sigma^2} < k + \Delta$
 Then $\int_{x=0}^{\infty} \int_{y=0}^{\infty} h(\mu^2, \sigma^2) dx$
 $\int_{k=0}^{\infty} dk \left(\int_{x=0}^{\infty} h(kx, x) \cdot x \cdot dx \right)$
 $= \Delta \cdot \int_{x=0}^{\infty} h(kx, x) \cdot x$
 we want this as a function of $(k = \frac{\mu^2}{\sigma^2})$.

~~4.2~~ $h(\mu, \sigma^2) = 57.20$; not so easy for us, since in 57.20, μ^2 isn't expressed as such.

A way out: Integrator along t , line e^{-t}



Essential part of h :

$$h(\mu, \sigma^2) = \frac{1}{\sigma^n} e^{-\frac{(\sigma_0^2 + (\mu_0 - \mu)^2)}{2\sigma^2}}$$

That would make \int integration easy, but it's not what I want!

Another way, let σ be the integration param. $\mu = k\sigma$, but we have

to multi. \int integrand by something because of $\frac{\sigma^2}{\mu^2} \leftrightarrow \frac{d}{d}$ conversion. This may not change form of integrand, but.

Well, if I had to d.f. for $\frac{\mu}{\sigma}$, I could get to d.f. for $(\frac{\mu}{\sigma})^2$ from it directly, so get to d.f. of $\frac{\mu}{\sigma}$.

$$h(n, \mu_0) = \int_0^\infty \sigma \frac{d\sigma}{\sigma^n} e^{-\frac{n(\sigma_0^2 + (\mu_0 - k\sigma)^2)}{2\sigma^2}}$$

Hor. axis

integrates, as σ , is

rather simple.

$$= \int_0^\infty \frac{d\sigma}{\sigma^{n-1}} e^{-\frac{n(\sigma_0^2 + \mu_0^2) - 2n\mu_0 k \sigma + n^2 \sigma^2}{2\sigma^2}}$$

constant factor

See 78.15 for a practical resolu. of this problem! just set $\mu_0=0$ (!)

No!

Hor, for $k \geq \sqrt{2}$ the integral diverges at $h(\sigma)$.

$k=2$ means $(\frac{\mu}{\sigma})^2 \geq 2$: Its utility

that will diverge near that region! $\frac{d}{d}$ is for a simple but!

for any k , \int converges for $n > 3$ or 4 , say,

$$\int_0^\infty d\sigma \sigma^{-n} e^{-\frac{A}{\sigma^2} + \frac{B}{\sigma}}$$

as a func. of z . ($0 < z < \infty$)

$$x = \frac{z}{y}$$

say we let $y = \frac{z}{x}$ that may do it!

$$\int_0^\infty dx x^{-n} e^{-\frac{A}{x^2} + \frac{B}{x}}$$

$$dy = -\frac{z}{x^2} dx \quad \text{we may also pop up w/}$$

$$dx = -\frac{x^2 dy}{z} \quad x = \frac{z}{y}$$

$$\int_0^\infty y^m e^{-2xz + bx} \quad \text{which is simply}$$

$$dx = \frac{z^2 \cdot z dy}{y^2 \cdot z} = \frac{z}{y^2} dy$$

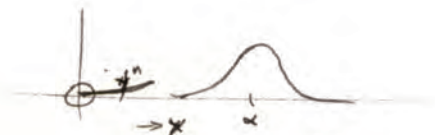
a moment of an offset Gauss. \int is probably well known.

$$-\int_0^\infty = \frac{z}{y^2} \cdot y \cdot \frac{y^n}{z^n} e^{-\frac{A y^2}{z^2} + y} dy$$

$$= z^{1-n} \int_0^\infty y^{n-2} e^{-\frac{A y^2}{z^2} + y} dy$$

$$\text{so we need } \int_0^\infty y^m e^{-\frac{(y-x)^2}{2\sigma^2}}$$

$$\int_0^\infty y^m e^{-\frac{(y-x)^2}{2\sigma^2}} dy$$

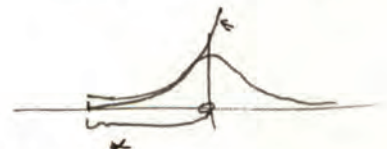


$$z = y - \alpha$$

$$y = z + \alpha$$

I think I usually want $x > 0$.

$$g(m, x) = \int_{-\infty}^\infty (z+\alpha)^m e^{-\frac{z^2}{2}} dz$$



$$\frac{\partial g(m, x)}{\partial x} = \int_{-\infty}^\infty m (y+\alpha)^{m-1} e^{-\frac{y^2}{2}} dy = m g(m-1, x)$$

$g(m, \alpha) = \int_0^{\infty} dx x^m e^{-\frac{(x-\alpha)^2}{2}}$; $f(0, \alpha) = \int_0^{\infty} e^{-\frac{(x-\alpha)^2}{2}} dx$ is like $\text{erf}(\alpha)$

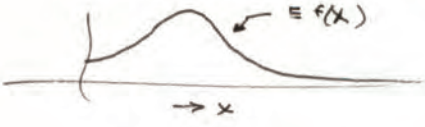
$\frac{\partial}{\partial \alpha} g(m, \alpha) = \int_0^{\infty} dx x^m (+ (x-\alpha)) e^{-\frac{(x-\alpha)^2}{2}}$
 $= \int_0^{\infty} dx x^{m+1} e^{-\frac{(x-\alpha)^2}{2}} - \alpha \int_0^{\infty} dx x^m e^{-\frac{(x-\alpha)^2}{2}}$ which seems to differ from 67.40! $\rightarrow 69.09$

03

Condon, Odishaw p. 1-141 col 1 seems relevant: 'Relation of moments about origin' "central moments"
 Some expansion of moments see ibid p. 1-140 col 2.

I think $X_1^{a_1} X_2^{a_2} \dots X_n^{a_n}$ means $X_1^{a_1} \cdot X_2^{a_2} \dots X_n^{a_n}$.

For this integral, the d.f. that I'm obtaining the moments of is $e^{-\frac{(x-\alpha)^2}{2}}$ for $x > 0$ but is 0 for $x < 0$



It is assumed so $\int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^{\infty} f(x) dx = 1$.

m in cond 2.06; m is mean of d.f. I think.

p. 1-141 col 1 eq 12.72, 12.73; In 12.72: shouldn't $\binom{n}{j}$ be binomial coeffs; to they are sym. in $j \rightarrow n-j$? (see ibid p. 1-10 col 1.)

$M_2 = \binom{2}{0} \alpha_0 m^2 + \binom{2}{1} m \alpha_1 + \binom{2}{2} m^2 \alpha_2$
 $= \alpha_0 m^2 + 2 m \alpha_1 + m^2 \alpha_2$
 so 2 terms cancel partly cancelled.

Maybe some terms are zero.

$M_3 = \binom{3}{0} \alpha_0 m^3 + \binom{3}{1} 3 \alpha_1 m^2 + \binom{3}{2} 3 \alpha_2 m + \binom{3}{3} m^3 \alpha_3 = 3 \alpha_1 m^2 + 3 \alpha_2 m + m^3 \alpha_3$

so it's D.K. In general, the terms in 12.72 are the result of combination;

I think that 2 terms always combine: i.e. $\alpha_0 m^n$ and $n \alpha_1 m^{n-1}$ giving $\pm (n-1) m^n$.

So in 12.72, remove last term on r.h.s.; the other terms are the first

$n-1$ "binomial coeff" terms.

1	1	1
2	1-3	3-1
3	1-4+6	4-1
4	1-7+13	partly cancel

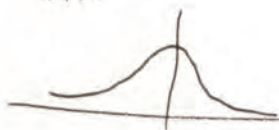
So, D.K.:

I need m , the first moment which is $m(\alpha) = \int_0^{\infty} x e^{-\frac{(x-\alpha)^2}{2}} dx$ which is sort of diversity integrable.

$\frac{d}{dx} e^{-\frac{(x-\alpha)^2}{2}} = -(x-\alpha) e^{-\frac{(x-\alpha)^2}{2}}$
 $= \int_0^{\infty} (x-\alpha) e^{-\frac{(x-\alpha)^2}{2}} dx + \alpha \int_0^{\infty} e^{-\frac{(x-\alpha)^2}{2}} dx$
 $= \alpha \int_0^{\infty} e^{-\frac{(x-\alpha)^2}{2}} dx + \alpha \text{erf}(\alpha)$
 (with $e^{-\frac{x^2}{2}} + \alpha \text{erf}(\alpha)$)

So I can compute $m(x)$:

next



$$\int_0^{\infty} x^n e^{-\frac{x^2}{2}} dx \quad \text{is probably well known.}$$

$$\int_0^{\infty} x^n e^{-\frac{x^2}{2}} dx \quad \text{I'm not so sure how well known.}$$

any way cond. of p 1-141 col I eq. 12.73 tells how to get Moments from ~~central~~ moments:

To get M.M. moments, I need all central moments up to & including m .

v.09 try solving 67.90 = 68.03: $\frac{\partial}{\partial \alpha} g(m, \alpha) = +m g(m-1, \alpha) = -g(m+1, \alpha) + \alpha g(m, \alpha)$

so $g(m+1, \alpha) + m g(m-1, \alpha) = \alpha g(m, \alpha)$ or $g(m+1, \alpha) = \alpha g(m, \alpha) - m g(m-1, \alpha)$

if this is true, we have here $g(m, \alpha)$ for all m , if we have

$g(m+1, \alpha) = \alpha g(m, \alpha) - m g(m-1, \alpha)$

$g(0, \alpha), g(1, \alpha)$ (or any 2 consecutive (may be not really consecutive) $g(m, \alpha)$)

$g(0, \alpha) = \int_0^{\infty} e^{-\frac{x^2}{2}} dx$

$= \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx \sim \text{erf}(x)$



$m g(m-1, \alpha) = -g(m+1, \alpha) + \alpha g(m, \alpha)$

$g(m+1, \alpha) = \alpha g(m, \alpha) - m g(m-1, \alpha)$

$g(1, \alpha) = \int_0^{\infty} (x-\alpha) e^{-\frac{(x-\alpha)^2}{2}} dx$

$= \int_0^{\infty} x e^{-\frac{(x-\alpha)^2}{2}} dx - \alpha \int_0^{\infty} e^{-\frac{(x-\alpha)^2}{2}} dx$



$\int_{-\infty}^{\infty} (x+\alpha) e^{-\frac{x^2}{2}} dx = \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx + \alpha \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} dx$
 $\sim \text{erf}(x)$
 $= g(0, \alpha)$

so $g(0, \alpha)$ & $g(1, \alpha)$ are expressible in terms of α erf(α) & $e^{-\frac{\alpha^2}{2}}$.

I still don't know sign of α .

$g_{m+1} = \alpha g_m - m g_{m-1} \quad ; \quad g_m = A^m$

$A^{m+2} = \alpha A^{m+1} - m A^m$

$A^2 = \alpha A - m \quad ; \quad A^2 - \alpha A + m = 0$

$A = \frac{\alpha \pm \sqrt{\alpha^2 - 4m}}{2}$ for fixed α^2 , m will eventually be $> \alpha^2$ so A is complex.

$g(1, \alpha) = \int_0^{\infty} (x-\alpha) e^{-\frac{(x-\alpha)^2}{2}} dx$

$= \int_0^{\infty} e^{-\frac{(x-\alpha)^2}{2}} dx - \alpha \int_0^{\infty} e^{-\frac{(x-\alpha)^2}{2}} dx$

$= + e^{-\frac{\alpha^2}{2}}$

$\frac{d}{dx} e^{-\frac{(x-\alpha)^2}{2}} = (x-\alpha) e^{-\frac{(x-\alpha)^2}{2}}$

Eqv. 12R seems unlikely to have $g > 0$ for all large m !

I got the signs in .03 wrong!

Should be $\frac{\partial}{\partial \alpha} E(m, \alpha) - \alpha g(m, \alpha) = m g_{m-1}$

So $\beta_{m+1} = \alpha \beta_m + m g_{m-1}$ — much more reasonable!

From 69.33 $A^2 = \alpha A + m$ $A^2 - \alpha A - m = 0$

$A = \frac{\alpha \pm \sqrt{\alpha^2 + 4m}}{2}$ so one positive A, one negative.

Still, if .03 is correct, g_m will be always > 0 since g_0 is positive.

Well, the best thing to do is evaluate 67.20 by this recursion method.

For 67.20 $\frac{\partial h(n, M_0)}{\partial M_0} = \int_0^\infty g \frac{dg}{g^2} \left(-n \frac{(M_0 - K G)}{G^2} \right) e^{-\frac{h(G_0^2 + (M_0 - K G)^2)}{2 G^2}}$
 $= -n M_0 h(n/2, M_0) + \frac{n K}{G} h(n+1, M_0)$

$\frac{\partial}{\partial M_0} \frac{-n(G_0^2 + (M_0 - K G)^2)}{2 G^2}$
 $= \frac{-n}{2 G^2} \frac{\partial}{\partial M_0} \frac{(M_0 - K G)^2}{\frac{1}{2}(M_0 - K G)}$
 $= \frac{-n M_0}{G^2} + \frac{n K}{G}$

change exponent to $(K G - M_0)^2$

$M_0 \rightarrow K G = z$; $K G = M_0 = z$
 $K G - M_0 = z$; $K G = z + M_0$
 $\sigma = \frac{z}{K} + \frac{M_0}{K}$
 $d\sigma = \frac{dz}{K}$
 $= \int_{-M_0}^\infty \frac{1}{K} \frac{dz}{\left(\frac{z}{K} + \frac{M_0}{K}\right)^{n-1}} \cdot e^{-\frac{n(G_0^2 + z^2)}{2\left(\frac{z}{K} + \frac{M_0}{K}\right)^2}}$

- C ✓
- E ✓
- 67.1
- 8 G 2.1
- SE (100) ✓
- Self amp
- ES est. for "m" in Pub 505

This isn't working out well!

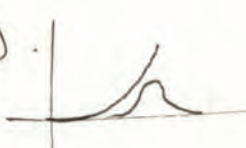
Whoops! .03 is not a linear eq. in m ^{Simple!}
 $\beta_{m+1} = \alpha \beta_m + m g_m$ ^{non-constant coeff. (B + linear)}
So this is not a simple eq.
If $\alpha = 0$ $\beta_m = (m-1)!$ ($\beta_m = (m-1)!$)
Nvr, it is linear: If X_m & Y_m are solns,
 $X_m + Y_m$ is a soln.

in
 $X_{m+1} = m X_m$
 $X_m = (m-1)!$ if $X_1 = 1$
 $X_1 = 1$; $X_2 = 1$
 $X_2 = 1 \cdot X_1$; $X_3 = 2$; $X_4 = 6$

Since m is usually large, g_n increases rather rapidly w. m !
Even if α had a large neg value, β_m would be large m but $m!$ (!).

This surprises me; I would think that β_n would not increase as fast!

look at β_n if $\alpha = 0$; $g_1 = 0$; $g_0 = 1$
 $\beta_2 = 1$; $\beta_3 = 0$; $\beta_4 = 1$



If $\alpha = 0$ $\beta_{m+1} = m \beta_m$ ^{say $\beta_0 = 1$}

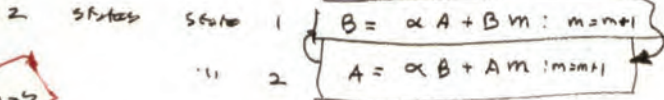
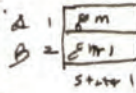
say $\beta_0 = 1$ $\beta_2 = 2$; $\beta_3 = 2 \cdot 4$; $\beta_4 = 2 \cdot 4 \cdot 6 = (4-1)!!$

similarly if $\beta_0 = 1$; $\beta_2 = 1$; $\beta_4 = 1 \cdot 3$; $\beta_6 = 1 \cdot 3 \cdot 5 = (6-1)!!$

0 1 6
1 0
2 1
3 2
4 $\alpha^2 + 3$
5 $\alpha^3 + 3\alpha + 4$; $\alpha^3 + 3\alpha$
6 $\alpha^4 + 3\alpha^2 + 4\alpha + 6$; $\alpha^4 + 3\alpha^2 + 5\alpha + 15 = \alpha^4 + 12\alpha^2 + 15$
7

to P.M.: $P(m) = \alpha g_m + m g_{m-1}$

use 2 consecutive files.



$A = \dots ; B = \dots ; m = 1$

Start w. $A=1, B=0$
Then $A=0, B=1$

$B = \alpha A + B * m ; m = m + 1 ; \text{print } m, B$

$A = \alpha B + A * m ; m = m + 1 ; \text{print } m, A ; \text{if } m < 40 \text{ then } \dots$

P.M. seems to work reasonably

$m=1: 1.3 E+25$

$m=2: 2.1 E+25$

$m=3: 3.55 E+24$

$m=4: 3.27 E+26$

$A=0 B=1 AL=1.5$

$A=1 B=0 AL=1.5$

$A=1 B=0 AL=0$

$A=1 B=0 AL=-1$

for $AL > 0$ its faster than $m!!$

for $AL = 0$ its $m-1!!$

output alternates in sign

A is always > 0 or contrary with

B is < 0

If α is -1.01 alternate values are negative.

Why? For $\alpha = 0, A = 1 ; B = 0$

- 1 1
- 2 0
- 3 2
- 4 0
- 5 8

If both A & B start at 1 w. $\alpha = -1.01$ then remains > 0 (!).

$$\int_0^{\infty} x^n e^{-\frac{x^2}{2}} dx = - \int_0^{\infty} x^{n-1} (-x e^{-\frac{x^2}{2}}) dx = \int_0^{\infty} x^{n-1} \frac{d}{dx} f(x) dx$$

$$\int (x^{n-1} d f(x) + (n-1) x^{n-2} f(x)) = x^{n-1} f(x)$$

so a recursion relation.

for $p = 3/2$ $\int_0^{\infty} x^{2n} e^{-p x^2} = \frac{2n-1!!}{2(p)^n} = \frac{2n-1!!}{2}$ for $p = 1/2$

$= \frac{n-1!!}{2}$ for even n ; which is observed in 70.35

I don't see why it should \uparrow very rapidly w. n — but maybe \uparrow in n is strange

but so that \uparrow effect of $e^{-\frac{x^2}{2}}$ is dealt w. strongly!

ALTV

< to B direction

Count starts
direction

1/4 \rightarrow 43/50

1 132 col.

ALTV using
Branon's M.

or F3

Deby \leftarrow direction

ginter / List 9th order
pat

Well, using the method up to 71.40 I probably could get the expected d.f., but it would be an enormous amt. of work, & fraught w. error. I might do some numerical runs to confirm my conjectures on what the soln. looks like

Conjectures:

- 1) The mean of the $(\frac{M}{\sigma})^2$ d.f. is about $(\frac{M_0}{\sigma_0})^2$.
- 2) Its var is about $\frac{1}{n}$ (E.f. no. of data pts). || Assumes uncorrelated bars. \rightarrow see 78.20: its exactly $\frac{1}{n}$ for small $\frac{M}{\sigma}$ (which is what we want, because it's $\frac{M}{\sigma}$ per bar but that we're after
- 3) If I assume $(\frac{M}{\sigma})^2$ has a Gaussian d.f. & the approx of $(\frac{M}{\sigma})^2$ is also Gaussian, but centered about zero; Then its easy to combine two d.f.s. to get a new Gaussian.

$$(\mu_1, \sigma_1^2); (\mu_2, \sigma_2^2) \rightarrow \left(\frac{\mu_1}{\sigma_1^2} + \frac{\mu_2}{\sigma_2^2} \right) / \left(\frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right); \left(\frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2} \right)^{-1}$$

now μ
now σ^2

The idea is that w. each σ_i^2 we associate a size of $\frac{1}{\sigma_i^2}$, so the results follow easily.

- 4) ~~ONE~~ ONE Q is: What size (σ^2) do we assign to the approx? We assign size (n) to the $\{S_i\}$ data, ~~in units of 1~~ If we assign $\text{Var} = \sigma_{sp}^2$
 to the approx $\frac{M}{\sigma}$ data. Then its approx size is $\frac{1}{\sigma_{sp}^2}$.

Say the typical stat has a yield of $\frac{M}{\sigma}$ ~~of 10%/yr~~ (half barrels); so $\frac{1}{2} \frac{M^2}{\sigma^4} = .1$

or $\frac{M^2}{\sigma^4} = .2$ or $\frac{M}{\sigma} \sqrt{.2} = .44$ We can use this as a rough "Ball Park" estimate.

After we've made lots of studies w. by n , we will have a better idea of what the d.f. is.

.25

- 5) This size of $(\frac{M}{\sigma})^2$ enters in following way: If we have lots of stats, each w. its own $\text{Gorc} (\frac{M}{\sigma})^2$ & a var for each Gorc . Then we pick the best Gorc . If there are many Gorc s to pick from - all w. about the same Gorc , & about same S.D. = σ_{amp} for that Gorc , then the peak will tend to be about Gorc higher than its true value.

If we have only 1 or only a few stats anywhere near the peak, then the peak has a better chance of being near an "unbiased estimate" of the Gorc .

In general, a conservative estimate of the Gorc of a stat that was picked as a peak will be that peak minus its S.D. If there are many "best stats" are chosen out of (few or many) stats, the correction will be $(\text{smaller or larger})$ than the S.D.

I once did a study of $\text{Re}/\text{inversion}$ ~~max~~ ^{expected} of a set of n indep normal

vars, as a function of n . no. of vars \rightarrow I don't know where it is, but \rightarrow 73.01

.30

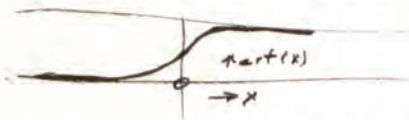
The results are on 75.27: A big problem is 74.18: How to estimate "how many" stats were around the peak one chose, due to correlation betw. stats. \rightarrow 74.22 discusses Recursive Soln.

3.25.97 TM

01: 72.36: Say we have n random vars of zero mean & var=1. What is p.d. that ti max of ti is X?

Qr. What is E. expected value of ti max?

define erf(x) as ti. probty that one of ti vars will be < x: erf(-∞)=0; erf(+∞)=1, erf(0)=1/2.



f(x) = 1/sqrt(2*pi) * e^(-x^2/2)

probty that peak is < x is n * (erf(x))^(n-1) * f(x) = d/dx (erf(x))^n

E(n) = E expected value = integral from -infinity to +infinity of x * n * (erf(x))^(n-1) * f(x) dx

perhaps find a recursion relation via int by parts. Mite be poss. for normal factor,

integral from -infinity to +infinity of n * (erf(x))^(n-1) * f(x) dx

d/dx (erf(x))^n = n f(x) (erf(x))^(n-1)

so ti. normal factor is 1.

g(x) =

integral from -infinity to +infinity of x * d/dx g(x) dx

integral from -infinity to +infinity of g(x) * d/dx dx = (x g(x)) | from -infinity to +infinity = infinity

Well, I could simply do it numerically — but its certainly better to have an analytic expression!

I already have a pen that may do something like integral from 0 to x of e^(-x^2/2) dx. I could decide on n in advance, then have it do ti. integration for all values up to n; similarly, — would take much less time than doing it from sequentially for larger n.

Strange! As n ↑ erf(x)^n gets smaller | so integral of x erf(x)^n f(x) dx should ↓!

on ti. other hand, how come

integral from -infinity to +infinity of n erf(x)^(n-1) f(x) dx is indep of n!

oh! integral of erf(x)^(n-1) f(x) = 1/n

IF I do it numerically, I can fast ti. pen by having it do ti. pen will be same as actual pen but (X < 1.)

07 gives E(1) = 0 since integral of x erf(x)^0 f(x) dx = 0. so it should.

I could just do a pretty dirty "Riemann" S. Do it for an adjustable dx, so

I can totally fast ti. pen for large & X then do ti. low runs w. small dx.

Also adjustable "Max n".

shouldn't bother w. dx = .01, X = -5 / .01 to +5 (~ 5'6)

start w. x from -36 to +36.

Test ti. pen on 10 first. = .25

The starting value of erf(x) at -36 or +36 can be obtained from some asymptotic formula

B.S. Or, I can just look it up. its a value you insert before running pen!

say erf for -5, -4, -3,

Pen has A = .00015 (TLU for erf -5), B = .001, C = .01. we set starting value for information at A, B, or C.

$$\int \frac{d}{dx} (n e^{nx} f(x)) = \int n(n-1) e^{nx} f(x) + \int n e^{nx} f'(x)$$

without S of n's

$$f = \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}}$$

$$f'(x) = -x e^{-\frac{x^2}{2}}$$

$$d \int n e^{nx} f(x) = \int n(n-1) e^{nx} f(x) + \int n e^{nx} f'(x)$$

$$\text{so } \int n e^{nx} f'(x) = \int n(n-1) e^{nx} f(x) \cdot dx$$

= what + what

Well, ok. say I know $E(n)$:

Another way easy (not so accurate) way would be strength mind cards. I have this program that generates Gaussian Vars. Just take the pair of 2 of them, square and sum.

Maybe reset random every 10,000 times by output from typewriter. So I type 10k characters. This gives 1 say 1k random nos. Just I use as 2 seeds.

Another way is to use just big "1 time per" program to put random nos from zip files by XORing many of them.

A Diff Q: To estimate "how many" "uncorrelated" histograms

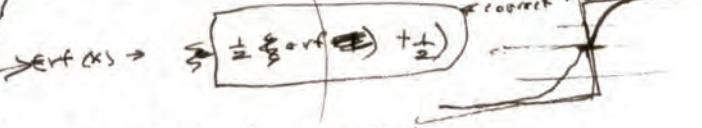
over in t. pool one pick t. peak of!
 There is also a case of an ∞ of stats (like in linear regression) in which one pick a peak. In this case, \equiv MaxM may fall what t. expected in broad expected error is.

Say one has 10 stats: look at var of "sum" of stats to get idea of how correlated they are. A formal soln: We assume that joint p.d. for all stats is known - then 83.30 - 84.15 has good theoretical soln.

Maybe try Maple: ① to try to do symbolic integrations ② To do numerical integrations

$$\int_{-\infty}^{\infty} \frac{e^{-x^2}}{\sqrt{\pi}} dx = 1$$

If I use $\text{erf} + \frac{1}{2}$ I got $\frac{61}{80}$ for result! I expected $\frac{1}{5}$.



Ok. I got $\frac{1}{5}$ as I should

With x^4 \int got .4332921579
 with using exponent \int got .6994711402

n=1 ~~putting~~ putting x in p into a. n=0 gave result ϕ as it should

n=2	(.6994711402) x 2 = .3989422804	~ 9" for calcn.
n=2.5	.511368	
n=3	.5984134206	40
4	.7278783067	60
5	.82234006	80
7	.9561345	100
10	1.08806	150
15	1.22747	
20	1.37084	

18

22

36

3.25.97 TM Int_n
 Evalf((10 * ~~...~~ (erf(x)/2 + 1/2) ^ M-1 * exp(-x*x) / Pi^(1/2), x=-infinity..infinity));

Evolve/Int
 ↳ resplot

TM 7-75.ms

This was saved as **TM 7-75.ms** ~~...~~ c:\mople\V2\bin

7" for n=4
 13" " n=80

One can get it by asking for TM 7-75.ms. ~~...~~
 Works ok for upto N=0, but N=150 Banks it!
 7" to do n=4

Note: results at 74.36 - .40 etc etc $e^{-x^2} = e^{-\frac{x^2}{\sigma^2}}$

i.e. s.d. = $\frac{\sqrt{2}}{2} = .707$, $\sigma^2 = \frac{1}{2}$.

So results should be mult by $\sqrt{2}$ for $\sigma=1$!

This gives ~ 1 for n=4. ~ 1.4 for n=10. 1.86 for n=20. 2.15 for n=40
 ~ 2.5 for n=100

It would be interesting to know if continued doubling eventually gives ~ same \uparrow in Y. S.

	0	0	.4	.2	.03
4	25	1.16	.38	.06	.00
	5	1.54	.06	.03	
	10	1.86	.03	.01	
double	20	2.15	.27		
double	40	2.42			
	80	2.5			
	100				

TM 7-75.ms

TM 7-75a.ms Pretty printing of formula.

The system may print about as easy as Basic: It has a simple do, for, and loop structure.

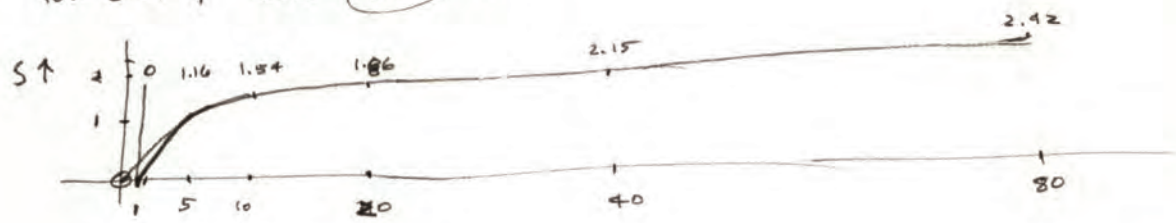
No sign of statement labels!

Also its interpreted lang.
 The "sum" operator simplifies many for new leaves.

27

O.k., so I evaluated 73.07 w. ~~...~~ $\sigma^2 = \frac{1}{2}$, $\sigma = \frac{1}{\sqrt{2}} = .707$
 erf in Maple is $\frac{2}{\sqrt{\pi}} \int_0^x e^{-t^2} dt$: its -1 for $x \rightarrow -\infty$; +1 for $x \rightarrow \infty$.

To get my out of 73.07 I did $\frac{1}{2} (Erf(x)+1)$: (still want σ , here).
 The values for 73.07 that I got were 74.36: Multiplying my sum by $\sqrt{2}$ to get f. results for $\sigma=1$; See 75.16 ff.



See **82.01 - 82.40** for $n = 1, 2, 4, 8, 16, 64, 128$: giving both mean and σ of "overestimate of σ "

3.26.97: TM

Geo Maple to Plot 3D:

TM7-76.m3

$$\left(\frac{1}{\sqrt{2\pi} \sigma} \exp - \frac{\sigma_0^2 + \mu^2}{2\sigma^2} \right)^n = \left(\sqrt{2\pi} \sigma \exp \left(\frac{\sigma_0^2 + \mu^2}{2\sigma^2} \right) \right)^{-n}$$

Say $\sigma_0^2 = 1$.

$$\left(\exp \left\{ \frac{(1+\mu^2)}{2\sigma^2} \right\} \right)^{-n}$$

Axis ~~fixed~~

Boxed

No parafactor.

5 (

so E plotted $\left(\sqrt{2\pi} \cdot \sigma \exp \left(\frac{1+\mu^2}{2\sigma^2} \right) \right)^{-n}$

-1.55 is μ $\frac{1}{2}$ width.

1.1 width μ .

n=5

26 μ $\frac{1}{2}$ values.

3.65 to ~ 1

2.65 width σ

wrap 2/0 acceptance box 77.01-25

n=10

$\theta = 90$ $\phi_{ij} = 180$

= top view.

Lost it!
run saved!

Get it to plot the funct for $\mu=0$ for various n:

n=1 $\left(\frac{1}{\sigma} e^{-\frac{1}{2\sigma^2}} \right)^n$ peaks indep of n; its at:

$$\frac{1}{\sigma^2} \left(z e^{-z^2} \right) \quad z = \frac{1}{\sigma} \quad \Rightarrow \frac{1}{\sigma^2} = e^{-z^2} \cdot (1 - z^2)$$

so peaks at $z=1$

1	$e^{-1} = .367879$	$z = 1$	$z e^{-z^2}$
1.01	$e^{-1.01^2} = .364$		
1.01	$\rightarrow .364$		
.99	$.371$		

$$1 \cdot e^{-z^2} + z \cdot -2z e^{-z^2} = 0$$

$$1 - 2z^2 = 0$$

$$z = \frac{1}{\sqrt{2}} = .707$$

$\sqrt{5}$	$\rightarrow .42879$
11+.01	$.42879$
11-.01	$.42879$

So peaks at $z^2 = .5$
 $\sigma^2 = 2$.

$$\frac{1}{\sigma} e^{-\frac{z^2}{2\sigma^2}} = z e^{-z^2}$$

$$1 \cdot e^{-\frac{z^2}{2}} + (z - 2) e^{-z^2} = 0$$

$$1 - z^2 = 0 \quad z = 1, \sigma = 1$$

$z e^{-\frac{z^2}{2}}$	so 1's peak.
1	.60653
1.01	.6064021
.99	.606469



$$(1-x^2)^n$$

$$(1-x^2)^n = .5$$

$$1-x^2 = .5^{1/n}$$

$$x^2 = 1 - .5^{1/n}$$

$$x = \sqrt{1 - .5^{1/n}} = \sqrt{1 - \frac{1}{2^n}}$$



z^2 to 1/2 on

$$e^{\ln z \cdot \frac{1}{n}} = \left(e^{\frac{1}{n}} \right)^{\ln z}$$

$$z^{\frac{1}{n}} = 1 + \frac{\ln z}{n}$$

$$z^{-\frac{1}{n}} = 1 - \frac{\ln z}{n}$$

30

34

$\frac{1}{2}$ width $\sqrt{\frac{\ln 2}{n}}$

101. Well that clears up the problem of the d.f. of $(\frac{1}{\sigma})^2$ to a large extent!

The d.f. for μ, σ (w.o. error) is (47.30)

7(77)

$$\left(\frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{\sigma^2 + (\mu - \mu_0)^2}{2\sigma^2}} \right)^n$$

The peak is ind. p of n . the μ peak is μ_0 ; of var σ^2

The σ " " σ_0 of var $\frac{1}{2} \sigma_0^2 \sim (2.5)^2 \sigma_0^2$ (from 76.09-10)

The var in both directions are $\propto \frac{1}{n}$ ($> 6.30 - 34$)

The s.d. $\frac{1}{\sigma}$ for σ is ~ 2.5 times that for μ : The d.f. for μ is sym - Gaussian.

d. d.f. for σ is

sharp on forward, slowly on beyond.

Simplify to $\frac{1}{\sigma} e^{-\frac{1+\mu^2}{2\sigma^2}}$ at $\mu=0$ i.e. at $\frac{1}{\sigma^2}$ at $\sigma=1$.

= $\frac{d}{d\sigma} \frac{1}{\sigma} e^{-\frac{1+\mu^2}{2\sigma^2}}$ or just use this take $\frac{1}{\sigma^2}$ of μ for both σ & μ

$$f = -\ln \sigma - \frac{1+\mu^2}{2\sigma^2} \quad \left. \begin{array}{l} \frac{\partial f}{\partial \mu} = -\frac{\mu}{\sigma^2} \\ \text{at } \sigma=1 = 1 \end{array} \right\} \quad \begin{array}{l} \frac{\partial f}{\partial \sigma} = -\frac{1}{\sigma} + \frac{1+\mu^2}{\sigma^3} \\ = -\frac{1}{\sigma} + \frac{1}{\sigma^3} = 0 \text{ at peak.} \end{array} \quad \left| \begin{array}{l} -\frac{1}{2} \sigma^{-2} \\ + \frac{1}{\sigma^3} \end{array} \right.$$

since f peaks at $\sigma=1, \mu=0$;

this is $1-3 = -2$ so σ peaks only $\sqrt{2}$ times

as wide as μ peak: not 2.5 times as wide! Woops! its s. a few way around!

to large $\frac{1}{\sigma^2}$ at σ peak means narrow peak! ~~Peak~~ of $\frac{1}{\sigma^2}$ is $\sqrt{2}$ times narrower than μ .

Plot of $\frac{1}{\sqrt{2\pi}\sigma} \exp^{-\frac{1+\mu^2}{2\sigma^2}}$ width at level .2 (part 13. 242) $\frac{32}{32} = 1$

" $\frac{1}{\sqrt{2\pi}} \exp^{-\frac{\mu^2+1}{2}}$ " " $\frac{50}{40} = 1.25$

so ratio is 1.25 rather than 1.414. I think as we go down the curve, $\frac{1}{\sigma}$ width in σ direction \uparrow more rapidly than in μ direction: Even going from .242 down to only .2, $\frac{1}{\sigma}$ error in σ is very asymmetrical.

I got 2 second derivs from Maple! wrt. μ

wrt. σ I got

~~sets~~ σ may, indeed have $\frac{1}{2}$ var of μ as in 120

These values over σ

±

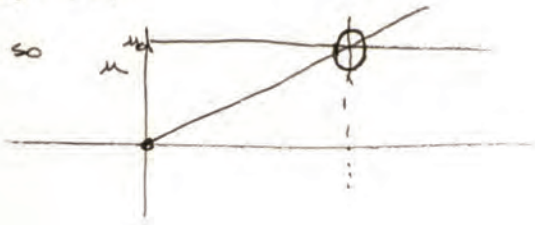
See TM 7-76. m) for how I was finally able to get these derivs.

evalf(subs(x=1, diff(e^{-x²/2}, x, x))); done second deriv!

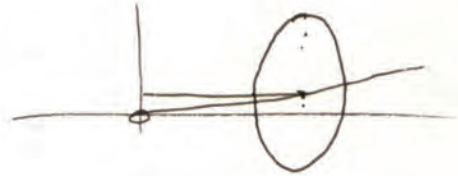
$$- .2419707244 \quad \mu$$

$$- .4839419988 = \text{exactly twice } \sigma \text{ part of } \mu$$

$$\begin{aligned} \sigma &= \frac{\sqrt{2}}{\sqrt{\pi}} e^{-\frac{1}{2}} \left(\frac{3}{2} \frac{5}{2} + \frac{1}{2} \right) = \frac{-\sqrt{2}}{\sqrt{\pi}} e^{-\frac{1}{2}} \\ \mu &= -\frac{1}{2} \frac{\sqrt{2}}{\sqrt{\pi}} e^{-\frac{1}{2}} \end{aligned}$$

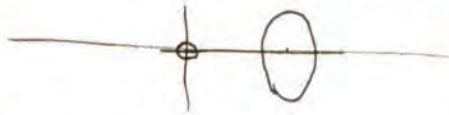


or maybe like:



In fact, we would probably get a v. prob. by setting $\frac{d}{g}$ to d and studying f .

Variation of S w. angle as a funct of n .



Apparently, this method (analytical) integrates easily!

The stuff from 66.20 ff is on this: I think I got a way to solve it: i.e. at 70.03: But the algebra is formidable!

By assuming $u_0 > 0$, I think it becomes very simple!

.15 I think S is 67.20: $\int \frac{d}{g} e^{-n \frac{(g_0^2 + (u_0 - kg)^2)}{2g^2}}$ set u_0 to 0, set g_0 to 1:

$$\int \frac{d}{g^{n+1}} e^{-n \frac{(1+k^2z^2)}{2g^2}} = \int \frac{d}{g^{n+1}} e^{-\frac{n}{2g^2} - \frac{nk^2}{2}}$$

which makes f . S indep of k , which is ridiculous! What is ridiculous is that S is indep of k in a trivial way! i.e. $\propto e^{-\frac{nk^2}{2}}$ i.e. of $\text{vanc} = \frac{1}{n}$ for $(\frac{d}{g})^2$ d.f. Which is strange since its indep of g_0^2 ! (not so strange! - see 72.09)

So check on expected values of u_0 & P for vanc.

$E u$ is u_0 (because of symmetry).

$E \frac{u}{g}$ was obtained on 65.02 ($\frac{u_0}{g_0} \sqrt{\frac{n-2.5}{n}}$ $n \geq 3$.)

$E (\frac{u}{g})^2$ maybe a bit different than 6. sp. of \dots

n	$n! / ((n-1/2)! \cdot \sqrt{n+1/2})$
1	1.009253008
2	1.003003704
3	1.001459433
4	1.000858136
5	1.000553896
10	1.000229072
15	1.000148
16	1.000067
20	1.000058139
	1.000038

So error $\sim \frac{1}{n^2}$
 $2 + \frac{1}{2}$

$\frac{1}{n^2}$
1.0152
1.0092
1.006

$\frac{.006}{n^3}$

$n \rightarrow 2n$
error $\neq e^{-\frac{1}{4-2^n}}$

$\frac{.0152}{n^2} = \frac{.006}{n^3}$

$\frac{e}{n^2}$

I got it to do a simple pm di pinhook that's around here.

There may be a way to get it to do tables: see index.

.0038

T. way it looks! What else been observing it part of error $\propto \frac{1}{n^2}$, part $\propto \frac{1}{n^3}$ or $\frac{1}{n^4}$

Re $\frac{1}{n^2}$ part is $\sim \frac{1}{n^2}$ for doubling!

1	} 3
2	
4	} 3.5
8	
8	} 3.75
16	
16	} 3.875

$\frac{1}{n} \neq \frac{1}{n^2}$ T. formula!

$\frac{.0152}{n^2} = \frac{.006}{n^3}$

This no. is neither not so accurate, but is not bad!

$\frac{1}{n^2} = \frac{1}{n^3}$

16 64

$\frac{x_{16}}{x_{32}} = 3.937088 \quad \Delta = .063$

$\frac{x_{32}}{x_{64}} = 3.9685560186 \quad \Delta = .032$

$x_1 = 100.9253008$

$x_{32} = .000015021 \quad x_{32}^2 = .015381504$
 $x_1 = .009253008$

$= 6.128496 \text{ E-}3$

so try $\text{arr} \approx \frac{.015381504}{n^2} - \frac{.006128496}{n^3}$

x_{16} should be $5.85875 \text{ E-}5$ according to formula

$16^3 = 2^{12} = 4096$

but it actually is $5.9139 \text{ E-}5$ so, not bad, but not perfect

So $\frac{x!}{(x-\frac{1}{2})!} \approx \sqrt{x+\frac{1}{4}} \left(1 + \frac{.015381504}{x^2} - \frac{.006128496}{x^3} \right)$

Probably don't need so many figures at present

$\frac{.01538}{x^2} - \frac{.00613}{x^3} \text{ corr.} \Rightarrow 5.85815 \text{ E-}5$

Doesn't seem to work for $n = \frac{1}{2}$: $x_{\frac{1}{2}} = \frac{\sqrt{\frac{1}{2}}}{\sqrt{\frac{3}{2}}} = 1.023326$

formula gives 1.01248 about $\frac{1}{2}$ as much error as occurs.
Still, corrected formula turns gives $\frac{x!}{(x-\frac{1}{2})!}$ for x as low as $\frac{1}{2}$, w/ error $\sim 1.25\%$
 \approx error $\sim \frac{1}{n^2}$ perhaps, for larger n .

So maybe try $A \left(\frac{1}{x^2} - \frac{2}{x^3} + \dots \right)$
 $4 - \frac{16}{5} + \frac{8 \times 64}{25} + \frac{8}{16} - \frac{8}{10.16} = \frac{1}{10}$

So: Now I'm at the pt. at which I can evaluate simple (say daily bet) strats,
 \Rightarrow tell how bad ~~bad~~ much spurious yield I get from trying to apparently best of several strats.

Next Q is how to search over strats: A good large for stat Generalist.
In particular I think the problem I've resolved is the Q of when to \uparrow size of old strats, as opposed to trying new strats w. small sizes.

A Q. That I haven't really solved yet, is the just how expensive of a strat enters

into the entire system: Perhaps the idea is that if one tries a large no. of strats, even if the data is pure noise, one will get apparently good strats - so we want to subtract, or divide out, the amount of yield we get in trying $\frac{1}{p}$ strats, in considering strats w. $\frac{1}{p}$ yield. I did do some earlier discuss. of this Q: Not sure if I had any good results: Maybe look. At any rate, the entire system should be such that w. random noise ~~input~~ data, we should have very small likely hood of thinking we have a good strat. or that the expected value of a strategy found should be very low. The stuff

www.2m612b.com
66 used
4# 56
62
38
Solidox 25
pr2.org.
68
Solidox: Contains KClO₄

.15

.35

3.2897 TM

? 73.01 or 72.35 seem like start of this "Thread".

around 71.20 - 75.90 is very relevant. It does help solve the problem - maybe even good enough for the present problem area - but I still don't feel I have a general idea of just how to define a solution!

38.87

Would a narrow spread/round to origin, do it? I don't think so. I'd guess the spread would be concerned w. n, the data length, rather than m, the no. of shares being searched.

Perhaps I want the formulae of 71.20 - 75.40 for large "n".

A useful inequality is B.S. Page 7.1.13!

$$\frac{1}{x + \sqrt{x^2 + 2}} < e^{-x^2} \int_x^\infty e^{-t^2} dt \leq \frac{1}{x + \sqrt{x^2 + \frac{2}{\pi}}}$$

$$\frac{4}{\pi} = 1.273239545$$

Apparently the erf as defined in Maple is same as that in B.S. $\text{var} = \frac{1}{2}$; $\sigma = \frac{1}{\sqrt{2}}$

My modified erf may be wrong, however.

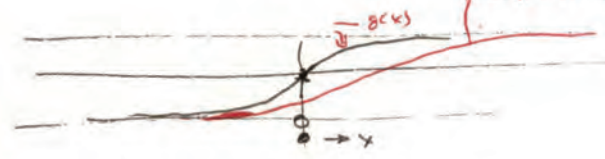
They define $\text{erf}(x) = \frac{2}{\sqrt{\pi}} \int_0^x e^{-t^2} dt$. $\text{erf}(0) = 0$; $\text{erf}(-\infty) = -1$

So it increases by 2 from $-\infty$ to $+\infty$: So it's not normalized as I meant! It starts at 0 w. $x=0$ - I want it to be 1 at $x=0$

So $\frac{1}{2} \text{erf} + \frac{1}{2}$ is correct (except for the var. - which I also corrected for)

So it looks like I did it right Peng.

I was using $F(x) = \frac{1}{\sqrt{\pi}} \int_{-\infty}^x e^{-t^2} dt$. Positive powered $F(x)$.



Actually, .08 is ²⁰ are close!

$$\int_0^\infty \text{vis.} \int_{-\infty}^x$$

$$\text{In fact } \int_x^\infty e^{-t^2} dt = \int_{-\infty}^{-x} e^{-t^2} dt$$

$$\frac{e^{-x^2}}{x + \sqrt{x^2 + 2}} < \int_{-\infty}^{-x} e^{-t^2} dt \leq \frac{e^{-x^2}}{x + \sqrt{x^2 + \frac{2}{\pi}}}$$

probably this approx was meant only for $x > 0$. But I don't know. But it's only $x < 0$.

So $\int_{-\infty}^{-x} e^{-t^2} dt$ is close to $\frac{e^{-x^2}}{2x}$

for $\ln(x)$.

for $|x| < 1$ say, it's not so close.

- but other approxs combined.

Work on 79.35: maybe 38.01 ff) 31.30 is on that problem.

10297 3 1/2 f

Another apparently related problem: Cover's Universal Portfolio: Using \vec{b} w. may yield is probably best, because yield is usually a very smooth function of \vec{b} .

For each \vec{b} corresponds to a different Strat. - However, Strats w. close \vec{b} values are highly correlated.

What is expected "over-optimism" involved in choosing \vec{b} w. ~~max~~ max past yield?

So: 3 apparently related problems:

1. T. problem of n uncorrelated Gaussians at $\sigma = 1$, mean all zero: What is an optimal value of max ? - also what is the var. of this max ? New problem. Easy to solve numerically: J-S put x^2 instead of x in 73.07.

I can get Maple to do this for a reasonable set of "n" values: Try to find way to get it to print it up in a more acceptable way! This would also solve problem of relative goodness of path vs. mean $|x|$ or $\text{max}(x)$

2. a way to get σ of a diff. (73.01 - 75.40 is recent work on this) 3. 2.01 - 22.40 is numerical work on offset σ of that offset.

01 (2) In Covari set of stats (80, 32, 35): what's expected (optimistic) of peak of output & continuum of highly correlated stats? — also for cov of P_{12} (optimism param).

(3) I guess General Problem of estimating expected future yield of a stat. or "action Alg" — and to var. of that expected yield: More generally, d.f. of expected future yield. This is a "harder problem" but it may be that way may make it easier to understand/solve it.

79.35 is some (perhaps) ~~relevant~~ relevant remarks.

02 (4) The work row day 57.01 - 58.40, 55.01 - 40: Given a ~~handful~~ set of empirical yields $\{S_i\}$, estimate $\frac{\mu_i}{\sigma_i}$ i. f. var. of μ_i . \rightarrow Also 67.20 is ~~78.15~~ 78.15 - 20

25% / Amo Rad Rho; $k=1.25 \rightarrow (1.25)^3 = \times 1.95/yr$; very smooth, $\times 1.95/125/yr$.

12 in 01 (2) Covari's problem, to get more: assume μ_i & σ_i underlying stocks are completely uncorrelated.

Each has its own μ_i, σ_i . It may be poss. to give them all same μ or all same σ , hvr.

In any case, I can see d.f. of any linear comb. of them.

\rightarrow Try it w. linear comb. of just 2 stocks.

Well, going back to k indep. stocks: It think linear combinations have a spurious path like "indiv. statistics". It ~~looks like~~ looks like Makm deals w. this properly by considering that σ_i coeffs have certain expected error, also that fitting a "Best set of coeffs" is an $n \times n$ & certain amt. of A. H. matrix is. \rightarrow N.B. those σ_i coeffs are "exactly" same as same σ spuriously good

\rightarrow So it does 2 things: Buy way to go. This is the general "Action Alg" problem?

See when optimizing.

Well, taking any stat: expresses set of yields as a μ, σ model. If ~~there~~ are no σ & P_{12} stats being considered: Compare it to μ, σ w. that of some data model. — Say same σ but $\mu=0$.

In the latter case, w. $\frac{\mu_0}{\sigma_0}$ of k stat will probly give its expected future yield! (also n is relevant to work of stats in table) \rightarrow $\mu_0 = \frac{n \cdot \mu_i}{n-k}$ & $\sigma = \frac{\sigma_i}{\sqrt{1 - \frac{n-k}{n}}}$ Makm relevant here.

If we have several stats — each has its own μ_i, σ_i $\frac{\mu_0}{\sigma_0} = \alpha_i$ i. its $n = n_i$.

well μ_i, σ_i will be of some interest, but α_i also $\frac{1}{T_i}$ (T_i is t. time in which yields occur) is of import. To first approx, we want stock w. best $\frac{\mu_i}{\sigma_i}$ $\frac{1}{T_i}$. To get P_{12} , we first

combine k stocks $\frac{\mu_i}{\sigma_i}$ observed i. it n_i w. $\frac{\mu_0}{\sigma_0}$ of all stocks, i. its n (= 552) from that into we get an μ & σ for that stock

See BS.2 for simple, exact, explain, treatment

27 If we have k stats, we can find k linear comb. of them that are statistically indep. — P_{12} (var σ^2) of these "linear comb." are k squares of the σ of k original cov. matrix.

(of course, obtaining cross correl. of non-synch. stats, is kn-trivial but I think what I want

30 is a set of $w_i \ni w_i$ ~~condition~~ balancing at every bet, μ mult yield is max.

SN En $\frac{27-30}{\dots}$ t. σ 's of k individual stats in a mix can be diff. — which makes it unclear as to what μ over all $\frac{\mu_i}{\sigma_i}$ really is!

1900 collect.

So just consider k indep. stats: Each w. own known $\frac{\mu_i}{\sigma_i}$ $\frac{1}{T_i}$ No not known, but of known σ & known no. of params

If we assume stationary ness, then μ_i & σ_i of each stat is clearly defined (for σ = cov). (T. successive dates may hvr, be correlated & But Rents auto cov is clearly defined).

If k data is correlated, use Makm to get d.f. of k indep. μ .

The original problem I may have been working on: That one has k $\frac{\mu_i}{\sigma_i}$ yield for a stat for a year, say $(\frac{1}{52})$ or number: One has k p.c. of k stats. From this info, it's hard to get much idea of reliability. If one is given k individual trades for a year, or monthly or weekly yields, ~~83.01~~ Spec

TM
3.29.97 Maple:

Attempt to get it to print output in nice format: See lines 83. If on Tables & arrays

One going to try to devise an array, and write a proc to fill it, then get maple to print it (P88 of Gauss)

$a := \text{array} (1..3, 1..5)$ hor. index. 3×5 array

$a[1, n] = n$ vertical index

$a[2, n] = \text{expected value of } x$

$a[3, n] = \text{expected value of } x^2$ $x := a[2, n]$ $a[3, n] - a[2, n]^2$
 $a[3, n] = \text{mean sqrt}(xx)$

```

first := 1; n := 5
while n <= 5 do
  nn = n+1; n = nn;
  a[1, n] = n;
  od;
Print(a);

```

end P122 PL

Print: expression PL
 | print: P35 PL
 L is not transfer
 "Printer"

5, 3

Tab creates extra blank line.

TM 7-81.ms.

first moment 0
 second moment 1
 zero moment

no 76 used in
 78
 952 hrs formula
 3 5 in every part.

n	2^n	first moment	σ	σ^2	σ^4	σ^6	σ^8
0	1	0	1.7071067	1.2329			
1	2	.3989422	.5838193701	.0680	.0363		
2	4	.727878	.49584	.0640	.024	13	1 2 4 8 16 32
3	8	1.00663	.4307569106	.0478	.0162	8 5	64
4	16	1.24874449	.38406367	.0364	.0134	5 3	7
5	32	1.4634768	.34757	.0287	.007	4	
6	64	1.657269	.318972	.023	.005	2	
7	128	1.834657	.296010091				

quite interesting! 6 + slower 2.8
 then $\ln n$!!

Note: .20 ff is ≈ 2 "upper bound"
 on SOY: If the stats have correlation 1, then
 this is a perfect no. of stats. If we have
 2 stats & corr = -1, the SOY is the same as for $N=1$
 If amounts to integrality on Abs value of one stat.

For several stats, w. diff. μ_i & σ_i^2 , it's not
 clear as to what's the best upper bound. If all σ_i are
 the same, then assuming all $\mu_i = \text{largest } \mu_i$ gives simple upper bound.

It would be interesting to know how good as
 $n \rightarrow \infty$. Does it \rightarrow an asymptote > 0 ?

If all σ_i are same ($\exists \sigma, \mu_i$) then
 there is a simple upper bound on SOY by assuming all
 μ_i 's are = max μ_i .

3.29.77 Maple

T.M. 7-82.ms

82.5
82.5

```
> a := array(0..7,0..3);
```

```
a := array(0..7,0..3, [ ])
```

```
> n := 0; ← Didn't seem to work w/ n := -1;
```

```
n := 0
```

```
> while n < 7 do
```

```
> nn := n + 1; n := nn; m := 2^n;
```

```
> a[n,1] := m; a[n,2] := evalf(int(m*x*(erf(x)/2+1/2)^(m-1)*exp(-x*x)/Pi^(1/2),x=-infinity..infinity));
```

```
> z := evalf(int(m*x*x*(erf(x)/2+1/2)^(m-1)*exp(-x*x)/Pi^(1/2),x=-infinity..infinity));
```

```
> a[n,3] := evalf(sqrt(z - a[n,2]^2));
```

```
> od;
```

```
>
```

```
> j := 0;
```

```
> while j < 6 do
```

```
> jj := j + 1; j := jj;
```

```
> print (a[jj,1], a[jj,2], a[jj,3]);
```

```
> od
```

```
> j := 2;
```

```
j := 2
```

```
> print (a[j,1], a[j,2], a[j,3]);
```

4, .7278783067, .4958403153

```
> for j from 1 by 1 to 7 do
```

```
> print (evalf(a[j,1]), a[j,2], a[j,3]);
```

```
> od;
```

m	first moment	σ
↓	↓	↓
1	ϕ	.707106781
2	first moment	

2., .3989422804, .5838193701
 4., .7278783067, .4958403153
 8., 1.006637430, .4317969106
 16., 1.248744490, .3840636718
 32., 1.463476863, .3475755328
 64., 1.657269826, .3189721555

$=\sqrt{2}$

Note: these values are for $\sigma^2 = \frac{1}{2}$; $\sigma = \frac{1}{\sqrt{2}}$.

For $\sigma = 1$, the values for first moment & second must be mult. by $\sqrt{2}$.

~~17, 1] 17, 2] 17, 3]~~

128 1.834657394 + 29601009)

```
>  
>
```

01: (31.40) Then its poss. to get the var. off. stat to some extent. I was thinking that if one knew the stat itself, one could know var. of yield (about 40) for random, MKK-like data - & this could be very useful for comparison - but if one had information about info to find var. of stat for "random data, one probably had not data to get ~~the~~ var. of ~~the~~ actual stat. itself!

About Maxm: Say a stat. has k continuous parms. + l bits of data - Then Sol 7 & 7B corr.

05 gives convergence rate as $\frac{1}{n} \leq \sum \text{err} < k + \frac{1}{2} \ln n \cdot \ln 2$
 If $\text{err} \propto \frac{1}{n}$: $\sum \text{err} \propto \ln n + k$
 $k = \frac{1}{2} l$ but k is wrong! it's k , (4.1) indep. of k !
 Say $\text{err} \propto \frac{1}{n^{2.2}}$ so $\sum \text{err} = \alpha \ln(n-2.2)$: not much better!
 I'm not sure of this! The "1/2" may have already been included in $\ln 2$.
 No!

So $\sum \text{err} = \frac{1}{n} + \text{a part that sums to } < k$: say $\frac{1}{n^{1.2}}$. This second part

is rapidly negligible for n sufficiently large.

So: Maxm ~~actually~~ actually has much to add to Sol 7 & 7B - by getting k .

expected error for linear regression, $\frac{1}{n-1}$ regression at n so the system is locally "near" Maxm, is related to stuff in Sol 7 & 7B! (See 20-21)

Well, in the case of MDL approx to ALP, one can find the no. of bits needed in each linear (or N.L.) coef, but a more exact value is obtained by integration

02 (in the param. space - has to do with Hessian at r. peak. So this integration obtains r. coef of l in (05)
 01 It also obtains $\frac{1}{n} \frac{n+1}{n-1}$ of Maxm.

02 I maybe we now have a usable ^{sufficiently} soln. to the (Action Algm) Evaln. Problem **AAE**
 for a stationary seq.:

We get the yields at successive bits $\{S_i\}$ - we then use a max. or several linear parms or N.L. parms, to get (w. Maxm) a Pd for yield of

next bit. I think this takes all imp. biases into account.
 03 (Not quite!) There is no mention of ϵ : pc of ϵ : AAE itself! 4.17: I'm not sure it's really relevant! See also some stuff up to 80.03 88.17
 T. fuzzy techniques will

Give a nice Pd. (Assignment) for the next bit of ϵ : AA. If we have several AA's, a pick the best, we have to modify results.

030 Here, say we have several objects & we have a d.f. for the force of each:
 (or, if the forces are not indep. & d.f. for the joint d.f. of their forces)

032 ^{where} So $P(\epsilon_1, \epsilon_2, \dots, \epsilon_k)$: What we want is the expected value of (the apparent peak minus the true peak). The true peak is, in the case of indep. forces, the max of the expected values of ~~the actual forces~~ the forces of each of the k objects.
 What it means, if the forces are not indep., is unclear!

036 Well, what's I want is expected value of errors in picking peak rather than max: So, for each poss. peak, we have an assoc. true expected value of the force of the object that took peak.
 Given P of 032 we can get the expected G_i of each of the k objects. At each pt. in the k dim space of parms: we have a choice of one of the k as peak, & we have to determine which peak is the ϵ_i of the assoc. object. We want the expected value

0) of Max difference. → (15)

C:\wind\System\PI\F16R.DLL

CTRL SHIFT E2

If we have a Continuum of AA's: (maybe not linear): \Rightarrow is we select a "best" one: This is like n.l. curve fitting w. Non-Gaussian Gorr. — I think its rooted like Maxm. → (15)

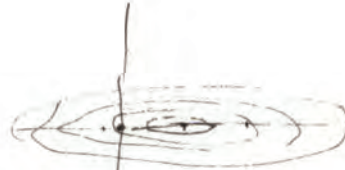
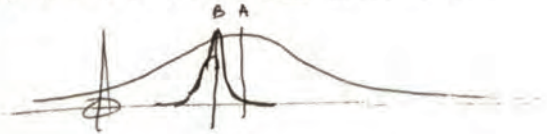
105 2 ① I think the AAE prob is what was formerly called SMA (Stock Mkt Advisor) problem. T. diffy was fixed if there were very many SMA's, there would be a spurious peak in the peak, even if mean were zero.

2 ② In the past, I may have been trying to solve SMA w. SSZ=1. While this is of interest, it may not occur so often, & it may give what look like unreasonable results, because usually one has SSZ>1.

3 ③ I think I did get some impt. progress on SMA: Maybe at SAAB (1990) or more recently.

4 ④ In some work of last few days, a prop ~~to~~ wasn't explicitly considered: This may be justified in that we may normally only consider stats w. mean values near that of the peak.

13 → 5 ⑤ Criticism of analysis of 83.36 - 84.01: It assumes that the peak was selected as "Best": This may not be so. A peak of a high var. stat may not be as good as a much lower low var. stat.



Overze.

At 89 Group howl.

20 Consider 2 d.f.s: a) One about zero w. zero d. b) One about .5 w. S.d. = .707. Picking a peak of b) gives an expected value of w. .5 - .707 = -.207 v.s. ← zero for d.f. a).

22 Here, the analysis of .20 - .22 is incorrect: But what would a correct analysis say?

→ It may be easy to consider ^{at least} 3 d.f.s: c), d) are zero mean var = 1; a) is mean = -.1, d = 0.

26 It's Important that I really understand what's going on! → 89.26

6 ⑥ Just exactly what was the AAE problem in TM? ~~How?~~ Seems Related to Jensen's Very Greedy algm. for evaluating "Pgm Modifn. stats".

1 ⑦ T. recent discn. of AAE has been for stationary systems using prediction from S: values alone, a very little about the structure of ϵ & σ . (perhaps xip value involved structure of strategy to some extent). For the more general non-stationary case we would use non-stationary predn. — with or w/o consideration of "structures" of strategies.

I want to first review what the problems are:

8036 - 81.08 lists 4 aspects of the problem: A-S aspect: we are given n strats (SMA advisors).
103 w. yearly yields of $\{Y_i\}$ $i=1/n$. How should we deal w. this? If n is large, choosing the best apparently
at best ones gives a spuriously high yield. If we only have a single S_i for each SMA, do we really
have much of a basis for decision?

Say we have n mutual funds or stocks, we have more detailed data: for some funds
monthly others daily others hourly (Fid. Sector funds might have hourly data available).

Or, we have SMAs that make buy/sell recommendations: They vary considerably in
108 f. no. of ~~the~~ bats/yr. Can we assign scores to these SMAs? One concern is that
f. set of SMA's "in business" is selected, so that no ones w. empirically poor
110 records ~~have~~ tend to drop out.

03-04 & 08-10 are major concerns. I think I was looking for a

"simpler" soln. involving ~~the~~ applied, because using a large no. of SMA's did ~~the~~ pc of each.

The recent work of ~~the~~ \$0.06 ff really didn't use ~~the~~ pc very critically (or simply):

[The applied was impl. in some of the work.] Note 83.28

(SN) Some impl. various ideas: 1) the "conjectures" of 72.08 -

- 2) the discuss of 81.12 ff.
- 3) the results of 73.01 - 75.10; 82.01 - 82.1.90. on Breslow
to choice of "best" of several strats.
- 4) Reviews of 57.01 - 68.01, 62.01 on expected
 $\frac{1}{6}$ of a strat is expected var. of that estimate

20 One very imp. applied re. SMA problem: A person ("SMA") make 10 by year choices
in a year - under 2^{10} diffrt names: The choice sequences in by yeards are published.
Under the assoc. names. one doesn't hear about low yield SMAs or
A tiny % to this occurs in "SM pickers" contests.
I felt that somehow the pc of the diffrt systems would somehow recognize what was
going on. To make money, the high yield SMA's simple use ~~the~~ by year, random
choices, using by management's by performance fees. Some of them will do very well!
So its a useful technique for making Money (by f. SMA's)

One Big Thing about TM's SM paramilitarism: the ~~the~~ specific diffrs of 20
do not occur: TM invents all of its own SMAs - it knows how many of them fail/succeeded.

On the other hand: A ~~the~~ reasonably good applied soln. to 20 would enable one to select
SMA's w/o pool them optimally - certainly a very valuable skill! - If there were
any really good SMA's one could find them. But there may not be any v.g. SMA's!

33 I did have one (initial) soln. for practical SMA problem: Take all yield data
of all off. SMA's: This can be modeled as a single μ_i or as a μ_i for each SMA -
or some mixture of the \geq models: See which has Most PG.

Almost all SMA's answer from step!

Once diffy is the dropping out of ~~the~~ SMA's w. low yield, (20) so one gets a spuriously
high yield for past. (SMA's w. poor past yields tend to drop out). One could, hrs. make that
part of the model - Not clear just how to do this - But I think I did figure out a
reasonable way.

Actually, one could study the "drop out" problem by using data (if available) of f.
list of all SMAs (say of a certain type): Plot yearly (or preferably monthly) yields.

3:31:57 TM SMA problem: AAE prob.

One could make a (to be sim. Anneal) model in which to dropout probly in each year was some simple function of yield for last k yrs.

By selecting only SMA's w. ≥ 5 yrs. "trch record", one gets rid of many that started & dropped out because of low yield. Hvr. this "byr rule" is partly a simple rule for skimming to for off. SMA's — So it gives a set of SMA's w. spuriously by yield.

If the best model was to some m.s. for all SMA's, w. a dropout probly based on near past yield, then this

Model could deal w. several SMA's w. large mean yields — (no base yields are sporadic.)

So, it looks like there are "Gradations" of the SMA problem: Depending on how much info one has about the creation of the SMA's.

One and extreme one has one or just a few SMA's & only a little data on their performances.

Then in the middle one has a larger no. of SMA's & more data on their strats — including some info on rate of exit of SMA's.

Out of the extreme one has, in TM, a complete knowledge of how all the SMA's were created & all of their know trades. **AAE problem**

ALP does not seem to offer any simple tricks to solve it all — (as far as I can tell... but I may some day discover such a trick); Use of strats to code corps may be part of such a trick. — Tho I'm not sure that that's so relevantly. What I'm doing now is just using the yield sequences of the strats, & using whatever models I can to fit the data & predict it. Presumably, knowledge of the structure of the strats would help devise better predictive models & this "yield as coding" seemed to be a step in that direction... but now I'm beginning to think, it can't have any more into that yield sequence itself!

A possible direction of using trying to use strats to get probly of yields.

Consider Carlos TK. Some basic ideas that behavior of MKT at different times of day is diffr. that behavior in one section could tell one about behavior in another section.

③ The "Occurance" model: that if a certain condition obtains we have an occurrence of "j". When j occurs, the MKT will run $\frac{dy}{dt}$ a distance $f(j)$ in $\frac{1}{2}$ time $T(j)$. From $f(j)$ & $T(j)$ we can construct an optimum betting strategy.

④ Time is in partoul: Just knowing the up down (at some level of granularity) history tend to be RW. — But knowing how long it took up or down betw. up v.s. dn steps could give very Non-Random walk.

⑤ Slope in the sense of less square error line can be useful for predn: maybe

"occurrence" criteria

(Actually, I'm not sure about how .27 it is supposed to work...)

0.27.18

0.09

0.19

0.27

0.28

Thresholded

3-31-97 TM SMA prob. AAE Prob. \rightarrow See (TM: LHL) 141.01 for complaints \odot
5.19.97

So 86.09-.19 characterizes diffrnt degrees of the SMA problem, depending on how much info one has on howt. info was generated.
T. extreme case where all info is known as in TM is the "AAE" case

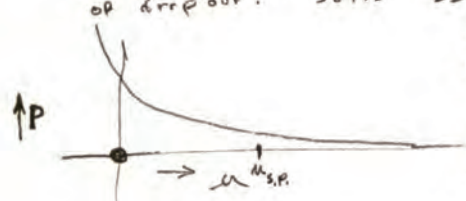
SN Analysis of clinical trials is a rather bad SMA problem - since the results under which trials were made are not very well known, a lots of cases get rejected that one does not know about.

One kind of soln of the SMA problems 85.33-.40 - to model the process of creation of the SMAs. This can be done in various ways & they do depend much on just what info - & how much info one has about the whole process. The ALP definitely is relevant to all of this.

As for the AAE problem, I may have it more or less under control, ~~see~~
83.22 was about to write a review of this, but we got sidetracked at 83.28 ff
Try to complete that review! (81.12 - 83.04) is relevant ^{very} (to ~~it~~).

On SMA: Anyway! 85.33 - 86.08 is a reasonable soln. to the SMA problem.
Dropout probly is a function of μ_i & σ_i . ("i" is SMA). Perhaps use $\mu_i - \frac{1}{2}\sigma_i^2$ as a variable.
Either \uparrow (i.e. life) ~~or~~ or $p = \text{dropout probly. per unit time}$. It has positive value $0 \leq p \leq 1$; $0 < p < 1$;
or use $\frac{1}{\tau}$ or τ . $\tau = \frac{1}{p}$ or $\tau = \frac{1}{p}$. ? probly of length τ is $e^{-\frac{t}{\tau}} = \lim_{t \rightarrow \infty} (1 - \frac{t}{\tau} p)^{\frac{t}{\tau}} = e^{-\frac{t}{\tau}}$

Expected value of τ so $\frac{1}{p} = \frac{1}{p}$ $p = \text{probly of dropout per unit time, for small times}$
of dropout. So its $(1 - p)^t$ as probly of no dropout after time t .
If σ is large, we can have a dropout in which a large (invisible to us) loss occurred during the first month & a low dropout - \rightarrow So this is



Anyway, the fitting of \uparrow as a function of μ_i & σ_i can be done separately (first).
Then, the second part, is to either assign a μ_i to each SMA; & a common σ ,
(or perhaps individual σ_i 's).

The fitting of data to $[\mu_i]$ or $[\mu_i, \sigma_i]$ is to σ, μ is a very direct, once \uparrow as a function of μ_i or $[\mu_i, \sigma_i]$ is obtained.

I think its a Maxm type fit, where so we need the location of each fit.
The "fit" removes the bias obtained by dropout of low μ_i SMA's.
The fitting "factors" into 2 separate parts. \uparrow is \uparrow $[\mu_i, \sigma_i]$ & μ, σ .

We want to val. wts. of the 2 models: $[\mu_i, \sigma_i]$ or $[\mu_i]$ v.s. μ, σ .
Another puzzle (seems combinatorial diff. to compute) is a given SMA can have either a common σ or common σ & a large σ up \uparrow or, its own σ_i (at higher cost).



4.1.97 TM

SMA solution \approx Final Soln.

MEXM:
(.23)

A simple explanation
for linear Regression

.01: 87.40: To do this study, I need several yrs. of data! SMA's predict preferably Monthly, maybe quarterly or yearly. There is a co. that sells such data. I might be able to get it free from them, if I promise to publish results! because ~~of course~~ such results would make their data more valuable, sellable.

First fit \uparrow to $\{u_i\}$ or $\{u_i, s_i\}$; Parameters part, ~~87.37~~

(\approx fitting $\{u_i, s_i\}$) is perhaps easy, but I would want to do a more complete ALP

~~analysis~~ analysis to make sure its all correct. This would involve studies like that ~~one~~

revised in 57.01-58.40; 65.01... also the "Bias due to Optimization": 83.36-84.01, 84.15-26

\rightarrow see 14.1.12 for Criticism

\uparrow General AAE problem is same as \neq SMA, but w'd drop out.

A preliminary review of AAE of General Soln. to AAE problem see 87.14;

83.22 was start of review, ~~see~~ also see 81.12-83.04

17: 80.03 There is also the Q of How relevant is the error of a strategy? presumably its then related to the no. of stats.

83.28R considers, so this could be relevant to the anti of spurious error and obtains by Optim.

Superficially, it would be not directly relevant. ALP ~~only~~ only gives a \approx upper bound to effective no of stats and is considerable. Here, in any real problem, one knows how many stats one is optimizing

over. Of more import., is the anti of correlation betw. the stats & their distribution of u_i 's & s_i 's.

\uparrow This may be an impr thing to understand, but I don't have time now! (look at stuff leading to 80.03 again)

23 \rightarrow SN Re: Maxm: I did this analysis of a 1 param linear fit - find the mean of n

measurements: If the mean were known (say μ_0) the best estimate of μ is ms, deviation from that mean. Here, one obtains in fact the empirical ms deviation from μ_0 , which \neq theta!

Moving from μ_0 to μ_{true} adds to var error by $(\mu_0 - \mu_{true})^2$: $E(\mu_0 - \mu_{true})^2 = \sigma_T^2 \cdot \frac{1}{n}$.

So $\sigma_T^2 \approx \sigma_0^2 + \frac{1}{n} \sigma_T^2$ or $\sigma_T^2 (1 - \frac{1}{n}) = \sigma_0^2$ | $\sigma_T^2 = \sigma_0^2 / (1 - \frac{1}{n})$ Error in next bot, has

σ_0^2 error plus error due to $E(\mu_0 - \mu_{true})^2 \approx \frac{1}{n} \sigma_T^2 \neq 0$, so error = $\sigma_T^2 (1 + \frac{1}{n}) = \sigma_0^2 \cdot \frac{n+1}{n}$

$\sigma_T^2 = \sigma_0^2 \cdot \frac{(1 + \frac{1}{n})}{(1 - \frac{1}{n})} = \frac{n+1}{n-1} \cdot \sigma_0^2$

In the case of linear regression: we have the same situation: error in coeffs causes

\approx possibly small σ_0^2 : $\sigma_T^2 = \sigma_0^2 (1 + \epsilon)$: Error in next bot is $\sigma_0^2 \cdot \frac{(1+\epsilon)}{1-\epsilon}$.

I guess $\frac{1}{1-\epsilon} = \frac{n}{n-k}$ where k is no. of coeffs! so $\epsilon = \frac{k}{n}$: $\sigma_T^2 = \sigma_0^2 \cdot \frac{n+1}{n-k}$

\rightarrow (actually if one variable has $k=2$, then have $E(\epsilon^2) = \frac{k}{n}$! $\frac{1+\frac{k}{n}}{1-\frac{k}{n}} = \frac{n+1}{n-k}$)

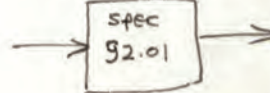
I should be able to get this result using the techniques of 57.01-58.40, 65.01-40.

i.e. the integrations leading to these results.

I think this result generalizes to curved fitting, & to large (for large) locally

linear n.1. curve fitting. Here, I'm not exactly sure of the results for linear curved fitting (i.e. 1.1.1. Curve fitting).

ALP address



ABCDEFGHI

Another apparently Impl Q: Maxm takes care of "overfitting errors" (as 88.23):
 It does this in a natural, general way. If I use Maxm for getting the d.f. for ~~the~~ "next step"
 of a stat. it would seem that this would not "interfer" w. the "next step" of correcting
 for the "Optzn effect" of selecting "Best" of several stats. It would seem that these
 are different, "separate" effects — so it would not be "double counting" by making
 the Second Correction.

So: In review of A.A.E. (Action/Plan. Evaln):

One has a bunch of A's: $\{A_i\}$: For each one one obtains the d.f. of yield
 G or whatever, for the next bet, trial, or next year or whatever it is one is concerned w.
 in the future. In most cases: These d.f.'s are maxm defined by $\{\mu_i, \sigma_i\}$

They are obtained by using the empirical yields of the A_i in the past: y_i^j says the yield
 of A_i in the j th trial. The d.f. $(\mu_i, \sigma_i, \text{say})$ can be obtained from y_i^j using any
 predn method one likes. ~~Then~~ One could just take μ_i mean of all y_i^j
 or use linear predn on past y_i^j 's or use data from other sources ~~at 2/0 use~~
 data from other A_k 's. At any rate, in the most genl case one has obtained the d.f. for the yield of the next bet (egmt)
 This is the most impl. desired output.

Anyway; In the simplest case one has this ~~set of~~ set of indep $\{\mu_i, \sigma_i\}$.
~~the~~ no of effect via MAPE.

From ~~the~~ this info using the ideas of 72.26 - 75.40; $82.01 - 82.40$: "one finds effect of fitting"
 For the correlated d.f.'s See discn. of 83.22 - 84.26 — Also raises impl. Q.

A review of where various ideas, derivations, are written: A "Bibliographic Review":

84.26: The Discn. of 83.30 - 84.26, on the General Case of "spurious & Assoc. w. Optzn" is
Very imp! I think I have all of the relevant ideas there, but I am suspect the details may not
 be correct. I really should write an exact formulation of what the problem is, as well as an
 outline of a Theoretical Soln.

Listing Impl. ideas, in the Throd in the notes:

- Some impl. ideas: 1) ~~the~~ MAPE: simple way to set it; b) relation to 50/78/73 cor. I.
- 2) SMA prob. vis. ~~the~~ AAE prob. ("Eliminats" \approx SMA): Stability vis. non-stationary
- 3) Spurious & of & assoc. w. "optzn": — size of effect & also is there
- 4) Analysis of simple Simple analysis of a betn. (or A.A.) in format $\frac{\mu}{\sigma}^2$ — applied to

under Genl. case:
 5) MAPE integrals: How to use MAPE. \odot . $\frac{x!}{(x-1)!}$ approxn. also, related to J's obtained 57, 58, 65.
 RAH

4.2.97 TM

Review of recent work: (Bibliog. review); Starts ~ 31.01

- .02 1) MEKM: 88.23-40: A simple way to understand $\frac{n+k}{n-k}$ factor! ~~push~~ for linear regression w. probab. extension to linear curve fitting; fin. regress. & curvature; 83.05-21: Relation of MEKM; to $C + A \ln n$ in Sol 78 Thm 3 (P427 cols) Also note: contin. of 88.23-40 suggested by 92.01ff
- .05 2) The SMA problem & the AAE problem (SMA Advisors; Actual Alarm Eval'n). Most recent work starts ~ 31.01. [up to RHP ph. was the "PATH" Thread] Relation of 2 probs: SMA was more recently recognized on 84-05; its relation to AAE was unclear in 86.09-19 (Dirat "isot SMA"; AAI is extreme case w. all info on how strats were created).
- .09 3) General Soln. of SMA problem in which one has some (but imperfect) knowl. of creation of strats ~~work~~; 85.33-86.08; 87.18-88.10: This is probly a quite practical soln. ^{work} 72.36 is 1st start of this thread. Some prelim. results on 75.27 This discusses only 1st part related to SMA, but not AAE. To solve SMA, we also need a soln to AAE problem.
- .15 4) The AAE problem: 89.08-29 various results! 89.23-24 is biblion work on Bias that is introduced when one selects the "Best" of many several strats \rightarrow we estimate the magnitude & varc. of this \rightarrow purp. \uparrow in "Yield". $\left\{ \begin{array}{l} \text{a general approach,} \\ 89.23-24 \text{ is biblion work on Bias that is introduced} \end{array} \right.$
- .18 5) The AAE Prob. Another aspect: The work summarized in 57.01-58.40; 65.01-40. ^{corrected version.} In this work, I was interested in the d.f. of $(\frac{M}{\sigma})^2$ of various strats. Say M_0 & σ_0 are observed empirical M & σ . ~~is not~~ for $\frac{M_0}{\sigma_0} \gg 0$ (this is hard to compute, but if $\frac{M_0}{\sigma_0} = 0$, (which is the usual case, since $(\frac{M}{\sigma})^2$ is for a single bet) then the d.f. is easy to obtain: its max is at ϕ ; its varc = $\frac{1}{n}$ (Discovered on 78.20) which simplifies things tremendously. It makes the main conjecture of 72-08-25 correct and so 72.08-25 does tell how to get an APSIP for each strat in a very simple way! - using some reasonable approx. about zero. (A. width varies in clear... try various values).
- .27 6) Some Mathematical results: Analysis of strategy w. yield data $\{F_2\}$ of empirical ~~work~~. M & σ of M_0 & σ_0 resp: n data pts. We find d.f. of $M, \sigma, (M, \sigma)$ also $\frac{M^2}{\sigma^2}$: eventually we find the d.f. & varc. of d.f. of $\frac{M^2}{\sigma^2}$ for ~~work~~ d.f. related about $(\frac{M}{\sigma})^2 = 0$ (which is most interesting case). Results on this are summarized on 57.01-58.40; 65.01-40. The part about d.f. of $(\frac{M}{\sigma})^2$ about $(\frac{M}{\sigma})^2$ peaked at zero is 78.20 One discovery: $\frac{X!}{(X-\frac{1}{2})!} \approx \sqrt{X+\frac{1}{4}}$: ~~work~~ 79.15 has makings of even better approx. w/o further correction error is $\frac{1}{8}$ for $x = 1$ } error $\propto \frac{1}{n}$ for large n

.28	for $x = .5$
.28	for $x = 1$
.28	" $x = 1.5$
.28	" $x = 2.5$
- .35 7) MAPLE: How to use: See 82.01-40 for one listing; See TM7-(), ms files in C:\Maple\BIN for examples of how to program MAPLE.

Going from 31.01 to present! Some random remarks!

32.17 puts in present SMA/AAE problems in the context of the Gault, M. problem.

One of the problems was to have a basis (in Larch) ^{for deciding} ^{between} increasing lengths of data sequences ~~are loaded~~ w. constant stats. V.S. creating new stats: evaluating them w. chart/sequences.

What looks like a soln. was obtained by getting the formula for expected $(\frac{M}{G})^2$ & its uncertainty "s".
as a function of the data for the stats considered, as well as their SS, and the assumed strip width of ~~$\frac{M}{G}$~~ $\frac{M}{G}$ about the origin.

T. main results re: Using Larch on a large data base of, say, SM data, is

90.02, .15, .18

① ④ ⑤

For later ~~return~~ return to this stuff: Read 90.01 - 40 is refs to get main ideas, then review

31.01 ~~to 92.46~~ to get some other impt ideas

.01: SPEC
88.40: MEXM: The discn of 88.23-40 seems relevant to Portfolio Management!

Also, it gives an example of a continuous ∞ of choices. In present case, say we have k SM's. We can find k linear combis. that are linearly indep. ~~linearly indep.~~ So we diagonalize f_i set. Each comb. has its own μ_i, σ_i & each μ_i, σ_i has its own uncertainty.

1.06 We then pick (within a constant factor) a set of wts $\rightarrow \left(\frac{\mu}{\sigma}\right)$ of sum is max. What is expected bias in the resultant $\left(\frac{\mu}{\sigma}\right)^2$? The idea of 88.23-40 certainly seems relevant!

As a "Study problem": We have k measurements of ~~for~~ f_i same quantity. ~~can~~ Try use of diff. precision & are indep! $[\mu_i, \sigma_i] \sum_{i=1}^k$. To make a linear comb. of these measurements we min σ^2 : ~~the~~ Hrr! what is σ^2 for next measurement, ~~is~~ ~~what~~ - well its unclear, depends on how one measures it! (~~is~~ which one of the "k" ways). I'm not sure I'm phrasing it.

rite Question!

Going back to .01-.06: ~~is~~ since C is optimizer $\frac{\mu}{\sigma}$; both μ & σ will tend to be biased.

Well: ~~is~~ This is an Emp. problem & ~~is~~ promising approach! But I don't have time just now!

See 141.31-40 (+ SMA problem) for more on this.
147.01 ff

4.16.97: TM: "50% solution"

A couple of related probs. probably having about same solns.:

.01 1) Given a system with ~~TM1~~ w. a TM1, & a TM2: How much ^{relative} time (a just when to switch) ~~betw.~~ betw. TM1 & TM2?

.03 2) ~~for~~ for $R_{TM1} = 100$ \$/yr: How much time to spend on predicting f. future, v.s. how much time spent on trying to solve its problems.

.06 In Barc (latest version) I dev'd to RLP problem by ~~finding~~ finding ^{largest poss.} probable/lower bound for $\sum z^{-ri}$: Given a certain available cc.

.06 was ~~an~~ an explication of a more el. version in which we wanted to find as

.10 many codes as poss. so as to maximize $\sum z^{-ri}$ in f. available cc: We wanted to generalize to continuous codes (equiv't to discrete codes).

Anyway, using .10 as a (prelim) model: The spirit of the thing is this:

Consider ① (.01): We list various TM1, TM2 models: T. order of listing we try to

approach is $\frac{T_{TM1} + T_{TM2}}{PC_{TM1} + PC_{TM2}}$ so its a single problem for optzn.

for ② (.02) again t. order of listing is $\frac{T_{predict future} + T_{solve future prob.}}{PC_{predict future} + PC_{solve future prob.}}$

1) dealing w. induction problems w.o. noise (\approx MTM) is an adequate form of f. general induction problem

Relevance of this stuff to AAE problem See Review 9.05 90.95 ff

Windows 10 tips

Actually ② is more complex: It involves expected future GCRs based on codes of f. past GCR values.

T. foug's (.01 ff) seems to me to ~~have~~ have great (likelihood of working).

It certainly needs a lot of work: The details have to be worked out: I have to draw up some examples.

T. reason its imp't. is that ~~it is~~ they are examples for fairly complex problem types that I want to be able to formalize so TM can usefully work on them.

Even .01 perhaps there is really only 1 goal: f. TM2 goal. What would soln. be if it had an enormous amt. of time (\approx cc)? Well, it would take a finite amt. of time to solve "present problem" of TM1: T. rest of t. time would be used by TM2.

Another possy: we work on TM2 only: in t. course of this work, TM2 ~~can~~ gets TM1 to work on its probs.

Looking at it another way: Say we have a TM that works on a bunch of INU's or 62 probs ("problem pool" approach). One of the problems in pool is "self improvement". It normally is given rewards (R) for solving or working on various probs, but no direct rewards for working on S.2. So how much time does it spend on S.2?

4.16.97 TM

TMC problem \approx IR problem (07)

PP 95 thru 100
do not exist

Well, the approach of TM: LHL \rightarrow 29.01 - 30.40! (29.2 + R in particular) seems relevant. jump to 101

We don't choose 2 times T_2 in the future, then try to optimally allocate time between these

tasks: S_1 , S_2
TM₁, TM₂

{ = X }

On the Q of how much time to spend making good models of the future: One way to look w. this is the "Problem Pool" approach: TM knows what future "Big Ticket" items are:

Def

TMC (\equiv Too many concepts \equiv IR problem): Say we have defined put into "many" every

Parvum we've proved, is every sub-function (in our function trees) that has proved "useful". We can associate unconditional PC's w. these concs, but for

use in building new concs is solving to problems, these uncond. PC's are

deriv

Usually too small.

The way out is conditional PC's: A particular conc. will have ^{been discovered} arrived in

a particular envt. This envt. will suggest where to use it next. Normally, various academic areas are divided up into special areas like Math, Physics, Chemistry... also sub-sub... classes.

A tree structure. Any conc. is normally closer to concs. & applications in its own branch of the tree. To get to other branches & sub branches requires a

certain amt. of traveling, & each step in this travel will have its peculiar kind of PC. (Wolfgang's "Directory" is one way of dealing w. this "compartmentalized

of knowledge") Soloved by the Academics ^{where} _{where} how long it took to be found, will give some constraints

The idea of .14-.15 is that how the conc. was invented, will give some constraints

(probistic) on how it is expected to be used in the future.

The TMC problem is closely related to the general IR (information) problem.

There are several implications here!

1) TMC \approx IR

2) Before TMC arises we have to generate/discover the relevant concs.

The process of discovery now, how long it takes, techniques used are related to the IR params used in the later retrieval of that conc. Furthermore, the techniques used in

IR will be somehow used to discover emerging concs. in the first place.

3) Human Heuristic: "State problem precisely, formally"

Machine heuristic: take precisely stated problem: generalize it related problems to other probs. that will help solve it. State problem in various languages; state problem visually, w. param. that machine takes. (Visual statement is more general)

ABCda